# Outmigrant Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site 1999 

September 2000

Submitted to
U.S. Fish and Wildlife Service
under subcontract to
CH2M Hill

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## EXECUTIVE SUMMARY

We operated two rotary screw traps side-by-side in the lower Stanislaus River near Caswell State Park (river mile (RM) 8.6) from January 18 to June 30, 1999 to estimate an index of outmigration abundance for juvenile fall-run chinook salmon. We estimated capture efficiency of the traps by releasing 4 groups of marked hatchery chinook and 5 groups of marked natural chinook, about $1 / 4$ mile upstream from the traps. Recovery rates of marked fish varied from $1.57 \%$ to $3.76 \%$. A multiple-regression was developed to predict daily trap efficiency based on flow, turbidity, and fish size, and this predictive function was recalculated for 1996, 1997, and 1998 outmigration data.

The estimated number of juvenile chinook that migrated past the traps between January 18 and June 30, 1999 was 1,321,042 with an approximate 95\% confidence interval of $1,007,443$ to $1,634,642$. The outmigration index by life stage was $1,155,424$ fry, 92,615 parr, and 73,003 smolts. Emigration of fry was already underway when sampling began in January 1999. Daily fry indices exceeded 20,000 on 27 days in 1999, but only once in 1998. In contrast, parr and smolt outmigrants were significantly less abundant ( $\mathrm{P}<0.05$ ) in 1999 than in 1998 as shown in the table that follows.

Caswell Outmigration Estimates by Life-Stage, 1996 through 1999

|  | $\frac{1996}{}$ <br> (Feb5-Jul 2) | (Mar $\frac{1997}{19-\text { Jun 27) }}$ | (Jan $\frac{1998}{29-\text { Jul 16) }}$ | (Jan $\frac{1999}{17-\text { Jun 30) }}$ |
| :--- | :---: | :---: | :---: | :---: |
| Fry | 31,767 | -- | 186,029 | $1,155,424$ |
| Parr | 1,596 | 7,011 | 209,911 | 92,615 |
| Smolts | 81,896 | 60,333 | 197,884 | 73,003 |

Note that fry estimates represent only a portion of fry outmigration.
The mean lengths of juvenile chinook gradually increased over the course of sampling, ranging from about 35 mm at the start of sampling (mid-January) to about 90 mm in late June. The length increase was slower than in 1996-1998 and the threshold size for classifying smolts ( 80 mm ) was not reached until May 10. This was 3 to 7 weeks later than other years. Mean lengths of fry captured at Caswell in January and February were similar from 1996 to 1999.

Mean lengths of natural chinook have been similar at Caswell and Oakdale in past years. In 1999 there was a noticeable difference in mean lengths between the sites beginning in March, once lengths began to increase above 40 mm . The difference was greatest in mid April, when fish reached approximately 75 mm at Caswell, but were still near 60 mm in length at Oakdale. The difference in lengths between the trap sites during 1999 may indicate that many parr paused to rear between Oakdale and Caswell. There was no sampling between the two sites to test this possibility, but data from our previous studies tend to support this hypothesis of rearing.

The smolt appearance index followed similar trends through time in 1998 and 1999. Fry (smolt index 1) were present through the end of April in both 1998 and 1999, with the exception of one fry captured in late May in 1999. Parr (smolt index 2) appeared later in 1999 (beginning of March) than in 1998 (late February), but in both years parr persisted until mid June.

During 1999, we captured 12 rainbow trout/steelhead at Caswell, ranging in size from 83 mm to 255 mm . Two distinct size classes were apparent (200-300 mm and <100mm), representing yearlings and young-of-the-year, respectively. More rainbow trout/steelhead were captured in 1999 than 1998, but 1999 counts were comparable to 1997.

## ACKNOWLEDGMENTS

This study was funded and directed by the USFWS. We are grateful to Marty Kjelson, Scott Spaulding, and Craig Fleming for their support and cooperation. We thank Bill Loudermilk, Tim Heyne, Jennifer Bull, and Steve Baumgartner with the CDFG for their help to plan, permit, and coordinate with our operations. We also thank the staff at the Merced River Hatchery for providing us with hatchery fish to conduct trap efficiency tests.

We thank the Tri-Dam Board of Directors, as well as Wayne Marcus, Manager of Oakdale Irrigation District, and Rick Martin, Manager of South San Joaquin Irrigation District, who provided equipment and personnel to facilitate this study.

We would also like to again express our appreciation to B\&B Farms who allowed us daily access to the river across their property, and the US Army Corps of Engineers for granting us special access through their parks, and for their protective surveillance of our equipment. We are especially grateful to Peggy Brooks, Lisa Vacarro, and Jason Anderson for their continued support of our activities throughout the year.

We are also grateful to the staff of the State of California Department of Parks and Recreation at Caswell Memorial State Park. Over the years they have provided us with continued support, including access to the river and our equipment, and provided surveillance of our equipment.

We also thank the Oakdale Waste Treatment Facility staff, Woody Woodruff, John Lane, and Lovanna Brown, for protecting and storing our equipment, and providing us access to the river.

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## BACKGROUND

Juvenile salmonid sampling at the Caswell site is part of a coordinated monitoring effort on the Stanislaus River to better understand how salmonid populations respond to both habitat restoration measures and flow management actions currently underway. Monitoring of juvenile salmon outmigrants and estimates of their abundance at the Caswell site are conducted as part of two CVPIA programs. The Central Valley Project Improvement Act (CVPIA), enacted in 1992, directs the U.S. Fish and Wildlife Service to develop and implement a series of restoration programs for fish and wildlife purposes, with the goal of doubling the natural production of anadromous fish in Central Valley streams by the year 2002. The first is the Comprehensive Assessment and Monitoring Program (CAMP). The goal of the CAMP juvenile monitoring program is to assess the relative effectiveness of categories of fishery restoration actions recommended by the Anadromous Fish Restoration Program (AFRP). The AFRP, established by the CVPIA, set anadromous fish production targets and recommended fishery restoration actions for Central Valley streams. The water management program on the Stanislaus River, which is one component of the AFRP and authorized under sections 3406(b1-3) of the CVPIA, has identified the need for juvenile salmonid monitoring at the Caswell site to help understand the effects of water management on salmonid population dynamics. Finally, outmigrant sampling 31.5 miles upstream of the Caswell site near the town of Oakdale presents unique opportunities to compare outmigration characteristics of juvenile salmon between the two sampling sites. Sampling at Oakdale is funded by the Oakdale and South San Joaquin irrigation districts. The monitoring described here for the Caswell site will also provide information that will inform the adaptive management process.

Sampling at the Caswell site in 1999 had four objectives:
$\varnothing$ Estimate the number of juvenile fall-run chinook salmon migrating out of the Stanislaus River in 1999,
Ù Determine the size and smolting characteristics of juvenile chinook salmon and rainbow trout/steelhead migrating out of the Stanislaus River,

Ú Identify factors that influence the timing, size and number of juvenile chinook salmon and rainbow trout/steelhead migrating out of the Stanislaus River.
Û Estimate the survival of coded wire tag releases from Knights Ferry and Oakdale Recreation Area to Caswell State Park in 1999.

## SUMMARY OF PREVIOUS MONITORING

Rotary screw traps have been used since 1993 to monitor timing and relative abundance of juvenile salmonids outmigrating from the Stanislaus River. Sampling has been conducted near Oakdale (RM 40.1) and near Caswell State Park (RM 8.6) by either California Department of Fish and Game (CDFG), US Fish and Wildlife Service (USFWS) or S.P. Cramer and Associates, Inc. (SPCA) (Table 1). Target species include fall-run chinook salmon and rainbow trout/steelhead. A summary of sampling in each past year follows.

1993 One trap was fished at the Oakdale site for a portion of the outmigration period in 1993, the first year of screw trap sampling in the Stanislaus River. The daily number of outmigrants was estimated from the results of two mark-recapture tests.

1994 One trap was operated at the Caswell site and no sampling occurred at the Oakdale site in 1994. Juvenile chinook catches were low in 1994, and no daily or seasonal

Table 1. Date, location and number of rotary-screw traps operated in the Stanislaus River, 1993-1999.

| Year | Trap <br> Location | Number of <br> Traps | Start <br> Date | End <br> Date | Flow-Year <br> Type |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1993 | Oakdale | 1 | Apr 21 | Jun 29 | Low |
| 1994 | Caswell | 1 | Apr 23 | May 26 | Low |
| 1995 | Oakdale | 1 | Mar 18 | Jul 1 | Low |
| 1995 | Caswell | 2 | Mar 27 | May 26 | Low |
| 1996 | Oakdale | 1 | Feb 1 | Jun 8 | High |
| 1996 | Caswell | 2 | Feb 5 | Jul 2 | High |
| 1997 | Caswell | 2 | Mar 19 | Jun 27 | High |
| 1998 | Oakdale | 1 | Jan 26 | Jul 15 | High |
| 1998 | Caswell | 2 | Jan 8 | Jul 16 | High |
| 1999 | Oakdale | 1 | Jan 18 | Jun 30 | High |
| 1999 | Caswell | 2 | Jan 18 | Jun 30 | High |

1995 Two traps were fished at the Caswell site in 1995. Catches of natural migrants were low, as were trap efficiencies estimated from recoveries of marked fish. However, since sampling was also conducted at Oakdale that year, it was possible to compare the size and timing of juvenile chinook between the upstream and downstream trapping locations. The abundance index of juvenile outmigrants was much higher at the Oakdale site. Screw trap efficiency was estimated at Oakdale with the release of 20 groups of marked natural or hatchery chinook.

1996 Two screw traps were fished at Caswell and one at Oakdale in 1996. Sampling began earlier in 1996 with the goal of estimating the total number of juvenile chinook outmigrants. However, we began sampling at Oakdale and Caswell in early
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February and found that fry were already migrating. We modified the trap set-up at Caswell to increase capture rates and released 15 groups of marked fish to estimate trap efficiency. Recapture rates varied from 0 to $12.08 \%$, and variation in capture efficiency was best accounted for by a logistic regression on turbidity.

1997 We fished two rotary screw traps at Caswell in 1997. No sampling occurred at Oakdale due to high flows. These high flows also delayed the initiation of sampling at Caswell until mid-March.

1998 Two rotary screw traps were fished at Caswell and one at Oakdale in 1998. Passage estimates at Oakdale were again higher than Caswell, however the ratio of Caswell to Oakdale was higher than seen in 1996.

The headwaters of the Stanislaus River originate on the western slope of the Sierra Nevada Mountains. The Stanislaus River and its tributaries flow southwest to the confluence with the San Joaquin River on the floor of the Central Valley (Figure 1). The San Joaquin River flows north and joins the Sacramento River in the Sacramento-San Joaquin Delta. There are several dams on the Stanislaus River that are used for flood control, power generation, and water supply. Water uses include irrigation and municipal needs, as well as recreational activities and water quality control.

Goodwin Dam, approximately 58.4 river miles (RM) upstream from the San Joaquin River confluence, blocks the upstream migration of anadromous fish. The lower river supports fall-run chinook salmon spawning between the town of Riverbank (RM 34) and Goodwin Dam (RM 58.4). Resident rainbow trout rear in 10-20 miles of the Stanislaus River below Goodwin Dam and adult steelhead are occasionally observed. However, it is not known whether a distinct anadromous population is present.

We reference river miles on the Stanislaus River throughout this report. River miles were determined with a map wheel and 7.5 minute series USGS quadrangle maps, (Knights Ferry, 1987 and Oakdale, 1987). Trapping locations and key area landmarks are listed below along with the river mile location for each site:

| Knights Ferry Bridge | RM 54.6 |
| :--- | :--- |
| Orange Blossom Bridge (OBB) | RM 46.9 |
| Highway 120/108 Bridge | RM 41.2 |
| Oakdale Trapping Location | RM 40.1 |
| Caswell Trapping Location | RM 8.6 |



Figure 1. Location map of study area on the Stanislaus River.

## JUVENILE OUTMIGRANT MONITORING

## Sampling Gear

We fished two rotary screw traps side-by-side in the mainstem of the lower Stanislaus River near Caswell State Park to sample juvenile salmonids as they migrated downstream. The screw traps, manufactured by E.G. Solutions in Eugene, Oregon, each consisted of a funnel shaped core suspended between two pontoons. Each trap was positioned in the current so that water entered the 8 ft wide funnel mouth and struck the internal screw core, causing the funnel to rotate. As the funnel rotated, fish were trapped in pockets of water and forced rearward into a livebox where captured fish could not escape.

Aluminum screen panels were placed in the rear of both liveboxes to provide fish with areas of refuge and to minimize stress and mortality. The screens caught wood and plant debris while allowing fish to pass through openings cut in the lower corners.

Each trap was held in place with $1 / 4$ inch cable fastened to large trees upstream on the north bank. The downstream force of the water on the traps kept the cables taught and near the water surface. Buoys marked the location of the cables for human safety. Although there is some recreational use of the river near the traps by small boats, canoes, and anglers in float tubes, the majority of river use in the vicinity of the State Park occurs downstream from the trap site.

## Trap Site Preparation

The Caswell trapping location was chosen by CDFG in 1994 since it was the farthest location downstream with adequate access to install and monitor the traps. In 1999, the traps fished in the same position as in 1996, 1997, and 1998, which was upstream approximately 100 yards from the site fished in 1994 and 1995. The trap nearest the left bank (looking upstream) was designated the north trap and the trap nearest the right bank was designated the south trap. These designations are the same as those used in the study since 1995. A sandbag wall extending approximately 5 ft out from the north bank was constructed in 1996 to divert flow into the traps and thereby increase trap efficiency. This wall remains at the site. The north trap fished about 10 ft downstream of this wall and approximately $8-12 \mathrm{ft}$ from the bank in an area where the velocity was highest.

## Safety Measures

Although recreational use of the river in this area was relatively light, we took precautions to warn park visitors and river users of the inherent dangers associated with the presence of rotary screw traps. One sign with large letters was placed upstream from the traps to warn river users traveling downstream towards the traps. The sign was approximately $4 \mathrm{ft} x 4 \mathrm{ft}$ with reflective red letters on a white background. A flashing light, similar to lights seen on roadside construction signs, was placed on the south trap to increase visibility at night. Reflective tape was applied to the A-frames of each trap to provide further warning.

To discourage people along the banks from swimming or floating towards the traps, numerous warning signs were placed at conspicuous places along the north bank and on the north trap facing the north bank. The signs warned of drowning danger near the traps and cautioned park visitors with messages such as "keep out" and "private property". The signs were in English and Spanish.

## Trap Monitoring

We installed the rotary screw traps on January 13, and began monitoring catches the morning of January 18. Monitoring continued until June 30 and the traps were removed July 6.

The traps were fished 24 hours per day, 7 days per week from January 17 to May 28. Beginning on May 28, and continuing through the end of sampling, the traps were raised after every Friday morning check and lowered again Sunday evening due to heavy weekend recreational traffic on the river. An exception was made the first weekend in June to accommodate sampling through the weekend following survival releases from Knights Ferry and Oakdale.

At times of high turbid flows and when we had recently released marked fish, we retrieved trap catches both in the morning and during the day to document daytime catches of juvenile chinook. Following the release of marked hatchery fish, we monitored the traps frequently until we were no longer recapturing marked fish.

During each trap check, we removed the contents of the liveboxes and identified and counted all fish captured. Random samples of 50 chinook and 20 of each other species were measured and their lengths recorded in millimeters every morning. We also measured all rainbow trout/steelhead and yearling chinook. When evening checks were conducted, random samples of 20 chinook and 10 of each other species were measured.

The traps were cleaned after all fish were recorded. When the river was carrying a high debris load, it was often necessary to clean the traps again in the afternoon to clear away debris accumulated against the funnel walls and in the liveboxes. Debris levels varied with changes in flow and weather conditions.

## Smolt Appearance Rating and Life Stage Classification

We recorded the external appearance of smolting characteristics for each chinook and rainbow trout/steelhead measured. Chinook smolting appearance was rated on a scale of 1 to 3 , with 1 an obvious fry (no scales, highly visible parr marks), and 3 an obvious smolt (silvery appearance, easily shed scales, blackened fin tips). Rainbow trout/steelhead smolting appearance was rated according to the CDFG scale of 1 to 5 (1= yolk-sac fry, 2= fry, 3= parr, 4= silvery parr, and 5= smolt).

In past years we estimated total outmigration for each juvenile chinook life-stage according to the following scale: fry < 45 mm ; parr 45 mm to 80 mm ; smolt > 80mm. Cutoff dates were chosen for each life stage when daily mean lengths were larger than the previous stage for 3 consecutive days. However, the daily lengths of sampled fish over consecutive days can bounce above and below values we use to separate the different stages. Therefore, in 1999 we used an algorithm to establish dates that separate fry from parr, and parr from smolts. When the number of consecutive days that fish fall into the larger life-stage permanently exceeds the previous number of consecutive days when the fish fall into the smaller life-stage, we used the date between the two runs of days to separate the smaller and larger size classes.

## Weight

A random sample of 50 chinook were weighed each week. If a sample of 50 could not be taken in a single day, weights were measured on consecutive days. In addition, weights were measured for all yearling chinook and rainbow trout/steelhead.

## Release Groups

Nine groups of marked chinook salmon (4 hatchery, 5 natural) were released to estimate trap efficiency (all releases made after dark). The CDFG supplied the hatchery fish from the Merced River Hatchery (MRH). All efficiency groups were released at the same site used in 1996, 1997, and 1998 (1/4 mile upstream of the traps). The number of fish in each group ranged from 671 to 2,550 .

## Holding Facility and Transport Method

Natural fish for mark-recapture experiments were marked and held at the Caswell site. Fish were held in net pens measuring $2 \mathrm{ft} \times 3 \mathrm{ft} \times 2 \mathrm{ft}$. The net pens consisted of 3/16 in. Delta mesh sewn onto frames constructed of $1 / 2 \mathrm{in}$. PVC pipe. The net pens were kept in an area of low water velocity off the south bank across from the traps. Fish were transported to the release site in coolers the morning of the release day. Towels were placed on top of the net pens when necessary to provide shade.

Hatchery fish were marked at MRH approximately one week prior to release. All hatchery groups were transported to the release site by CDFG the day of release. Fish were placed into net pens at the release site and allowed to recover prior to release.

## Marking Procedure

Juvenile chinook were marked by cold-brand or dye inoculation. Before marking, fish were anesthetized with MS-222 (Schoettger and Steucke 1970). Once anesthetized, the appropriate mark was applied. Fish were cold-branded by freezing a branding stick in a thermos of liquid nitrogen. Fish were placed into a PVC slide and the freeze brand was
applied by placing the tip of the branding tool against the body of the fish. Minimal pressure was applied for approximately 2 seconds. Each fish received only one mark. Fish were dye inoculated by placing the tip of a MadaJet or Pow'rject against the caudal (top or bottom lobe), dorsal, or anal fins (Hart and Pitcher 1969). Minimal pressure was applied as dye was injected into the fin rays. One mark was applied to each fish, and all fish in a group received the same mark. Location of the mark was varied between groups so that each group could be uniquely identified. The dyes used were photonic yellow and pink (NewWest Technologies, Santa Rosa, CA) and light green provided by CDFG. All were chosen because they were known to provide highly visible, long lasting marks.

## Prerelease Sampling

Marked fish were sampled for mean length and mark retention. Fifty fish were randomly selected from each distinctly marked group and anesthetized. Mark retention was rated as present or absent. If any of these 50 were found to have no mark, an additional 50 fish were sampled. The proportion of fish found to have clear marks in each group was used to estimate the actual number of fish released using the following expression:
number released = proportion mark retention * number in group.

All groups of natural chinook were also counted prior to release in order to obtain a more accurate number released. Release numbers for hatchery groups were calculated by subtracting the number of mortalities from the total number marked.

## Release Procedure

Fish were released directly from the net pens in which they were held. A dip net was used to remove and release about 50 fish per minute. The time required to release each

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 marked group ranged from 15 to 60 minutes. This release procedure was similar to the procedure used in 1996, 1997, and 1998. The gradual release of fish was intended to prevent the fish from behaving as a single school by dispersing them in time and space, to mimic the distribution of natural migrants. In 1996, 1997, and 1999, release of each mark group was separated by 5 to 15 minutes. In 1998, some mark groups (May 10, May 18) were released simultaneously.
## SURVIVAL TESTS

## Release Groups

Marked juvenile chinook salmon were released at five locations and recaptured at two locations for the purpose of determining the feasibility of estimating survival in different river reaches with CWT's. Five groups of marked hatchery reared chinook salmon were released in the Stanislaus River to estimate survival. The CDFG supplied the hatchery fish from the Merced River Hatchery (MRH). All groups were marked and held at MRH approximately one week prior to release and transported to the release sites by CDFG the day of release. One group of 25,536 coded wire tagged (CWT) fish were released from Knights Ferry (RM 54.6) the night of June1. Two groups of approximately 5,000 CWT/dye inoculated fish were released just below the Oakdale screw trap (RM 40) the nights of June 1 and June 2. Another two groups of approximately 5,000 CWT/dye inoculated fish were released just below the Oakdale Recreation Area at the Gambini Ranch (RM 38) the nights of June 1 and June 2. All survival test fish captured in the screw traps were given a secondary dye mark and released the morning following capture. The results of this evaluation are presented in Appendix D.

## Holding Facility and Transport Method

Fish for survival experiments were marked and held at MRH in outdoor troughs covered with 1" nylon mesh. Fish were transported to the release sites by CDFG the morning of the release day. All were held in $4 \mathrm{ft} x 4 \mathrm{ft} \times 4 \mathrm{ft}$ net pens until release, with the exception of the Knights Ferry group which were pumped directly from the 450 gallon hauling tank into the river.

## Marking Procedure

Juvenile chinook for survival tests were marked by coded wire tag (CWT) and dye inoculation. CWT insertion and adipose fin clips were done by Big Eagle and Associates prior to dye inoculation. One dye mark was applied to each fish, and all fish in a group received the same mark. Location of the mark was varied between groups so that each group could be uniquely identified. The dyes used were black waterproof engrossing, calligraphy, and India inks (Higgins) and photonic pink (Oakdale re-mark only) and yellow (Caswell re-mark only).

## MONITORING OF ENVIRONMENTAL FACTORS

## Flow Measurements

Daily flow in the Stanislaus River was obtained from the California Data Exchange Center (CDEC). All flows cited in this report were measured at the Orange Blossom Bridge gage by the US Geological Survey (USGS). The flow data are daily means; instantaneous flows during freshets were higher. Depth-velocity profiles were taken in front of the traps once per week.

Water velocity entering the traps was measured each day with a Global Flow Probe, manufactured by Global Water (Fair Oaks, CA). Daily average trap rotation speed for each trap was also recorded by measuring the time, in seconds, for three contiguous revolutions every morning. The average time per revolution for each trap was then calculated.

## Water Temperature and Turbidity

Daily water temperature was measured with a mercury thermometer at the trap site. An Onset StowAway recording thermograph was also installed to record water temperature once per hour throughout the sampling season. Recording thermographs maintained by SPCA are also stationed at Goodwin, Knights Ferry, Orange Blossom Bridge, Oakdale, and McHenry. Daily mean temperature was derived by averaging the hourly measurements. Temperature data is also available from stations maintained by USGS at Goodwin, Oakdale, and Ripon.

Turbidity was measured each day with a LaMotte turbidity meter, Model 2008. A water sample was collected each morning and later tested at the field station. Turbidity was recorded in Nephelometric Turbidity Units (NTU's).

## OAKDALE TRAPPING SITE

Rotary-screw trap sampling was conducted by S.P. Cramer and Associates under a separate contract at an upstream site near the Oakdale site (RM 40.1) between January 18 and June 30.

## STATISTICAL ANALYSIS

The outmigration estimates (Caswell and Oakdale) in this report were prepared by Dr. Doug Neeley (IntSTATS 712 12 ${ }^{\text {th }}$ Street, Oregon City, OR 97045). Dr. Neeley's complete methodology is presented as a complete report in Appendix A.

As part of this effort the AFRP had two independent statisticians review the statistical methods used for this and previous year's analyses. Comments by each reviewer and a response to comments from Dr. Doug Neeley are presented in the appendices B and C.

## OBJECTIVE 1: ESTIMATE THE NUMBER OF CHINOOK SALMON MIGRATING OUT OF THE STANISLAUS RIVER IN 1999

## Trap Catches of Chinook Salmon

From January 17 to June 30, 1999, we captured a total of 41,234 juvenile chinook salmon in the Caswell screw traps (Table 2). This compared to 19,903 in 1998, 2,357 in 1997, and 2,468 in 1996. Peak catch of 2,322 chinook fry during the outmigration occurred on February 10, 1999 (see Appendix A). With the exception of 3 days in March (trap repairs) and weekends from Memorial Day through June (high river traffic), the traps were fished everyday. It is uncertain whether or not a significant number of fish outmigrated prior to trap installation; however, degree day analysis suggests outmigration started only 10 days prior to the onset of sampling. This would indicate that only a small portion of the run was missed (discussed under Objective 3).

Table 2. Summary of 1996-1999 trapping seasons at Caswell.

|  | Period | Number of | Trap Catch |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Year | Sampled | Days Sampled | North | $\%$ | South | $\%$ | Catch |
| 1996 | Feb 6 - Jul 2 | 142 | 795 | $32.2 \%$ | 1,673 | $67.8 \%$ | 2,468 |
| 1997 | Mar 19 - Jun 27 | 98 | 408 | $17.3 \%$ | 1,949 | $82.7 \%$ | 2,357 |
| 1998 | Jan 8 - Jul 16 | 154 | 3,053 | $15.3 \%$ | 16,850 | $84.7 \%$ | 19,903 |
| 1999 | Jan 18 to Jun 30 | 152 | 31,949 | $77.5 \%$ | 9,285 | $22.5 \%$ | 41,234 |

## North and South Trap Catches

As in past years, during 1999 there were differences in the numbers of fish caught in the north and south traps. Also similar to past years, daily catch data were partitioned according to size class, where:

| fry: | length <= 45 mm |
| :--- | :--- |
| parr: | $45 \mathrm{~mm}<$ length <= 80 mm |
| smolt: | length> 80 mm. |

The daily mean number of fry sampled from the north trap was significantly greater than that from the south trap ( $\mathrm{P}<0.0001$ ), but the mean number of parr and smolt sampled from the north trap was significantly less than that from the south trap ( $P=0.0045$ for parr and $P=0.0003$ for smolt).

Mean lengths of each life-stage captured were compared in the north and south traps during 1999. North-trapped fry were significantly smaller (more than 1 mm on the average) than south-trapped fry ( $\mathrm{P}<0.0001$ ), but the differences between the lengths of north-trapped and south-trapped parr and smolt were less than 0.04 mm and not significantly different from 0 ( $\mathrm{P}>0.9$ ). In 1998 the north trap also caught smaller fry (and parr); however, the north trap caught larger smolts than the south trap. (No such comparison was made for 1996.)

Prior to 1999 the south trap consistently captured more fish than the north trap (see Table 2). In 1999, the north trap captured $77.5 \%$ of the total number of chinook captured at Caswell (Figure 2). In 1999 the traps were positioned using the same methods and hardware in near identical positions as in past years. In contrast to 1998 sampling, a large tree was lodged on the cables that position the traps during the 1999 fry outmigration period. The large size of the tree prevented removal, and it remained in the cables from
the second week of sampling to the end of fry outmigration.


Figure 2. Catch percentage for north and south traps 1996-1999.

The tree, present during February and March, pushed the trap approximately 6 feet closer to the north bank than it had been in previous years. The entrance velocities at the north trap ( $3.2 \mathrm{ft} / \mathrm{s}$ ) were higher than at the south trap ( $2.9 \mathrm{ft} / \mathrm{s}$ ) during the period the tree was caught in the cable (Figure 3). However, when the log was not present entrance velocities were consistently higher at the south trap. The presence of the tree during the fry outmigration might explain the changes in catch rates or size differences of fry; however, since it was gone during the majority of the parr and smolt outmigration, they were likely unaffected.


Figure 3. Schematic of trap positions in 1998 and in 1999 with tree lodged in cables.

The north trap may have captured more chinook this year because of the higher velocities through the trap when the majority of fish (fry) were migrating, and because the trap fished closer to the bank during fry migration, where fry probably prefer to migrate. The change in velocity pattern upstream from trap as a result of the tree may also have "guided" more fish into the trap, such that the higher catch is not solely a function of the increased entrance velocity or location.

We reviewed stream cross-sectional measurements taken directly in front of the traps to determine if changes in the stream channel may have influenced 1999 trap catch rates. The stream channel morphology did not appear to change significantly between any of the years measurements were taken.

Differences in the numbers of parr and smolts captured between traps during 1999 are likely a function of spatial abundance, not size bias. We found no significant difference in the mean lengths of parr and smolts captured in the north and south traps. Therefore, higher catches of parr and smolts in the south trap may indicate that more fish tend to migrate near the middle of the river channel, rather than near the bank.

## Trap Efficiency Estimates

Nine trap efficiency tests were made on 8 days during the months of February, March, and June (Table 3). A total of 4 marked hatchery groups and 5 marked natural groups were released about $1 / 4$ mile upstream of the traps at night. Capture efficiency ranged from $1.57 \%$ to $3.76 \%$.

Table 3. Release data for all chinook used for trap efficiency tests in 1999.

| Release <br> Code | Release <br> Date | Mark <br> Type | Fish <br> Stock | Release <br> Time | Adjusted <br> \# Released | Number <br> Recaptured | \% <br> Recaptured | Length at <br> Release <br> (mm) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| C1 | 20-Feb-99 | Brand | Natural | Night | 2550 | 96 | $3.76 \%$ | 33.2 |
| C2 | 27-Feb-99 | Brand | Natural | Night | 1672 | 43 | $2.57 \%$ | 35.6 |
| C3 | 02-Mar-99 | Brand | Natural | Night | 830 | 29 | $3.49 \%$ | 34.1 |
| C4 | 09-Mar-99 | Brand | Natural | Night | 962 | 20 | $2.08 \%$ | 36.1 |
| C5 | 17-Mar-99 | Photonic | Natural | Night | 671 | 15 | $2.24 \%$ | 42.8 |
| C6 | 02-Jun-99 | Panjet | Hatchery | Night | 2500 | 63 | $2.52 \%$ | 83.6 |
| C7 | 03-Jun-99 | Panjet | Hatchery | Night | 2487 | 39 | $1.57 \%$ | 84.2 |
| C8 | 04-Jun-99 | Photonic | Hatchery | Night | 2039 | 68 | $3.33 \%$ | 82.5 |
| C9 | 04-Jun-99 | Photonic | Hatchery | Night | 2002 | 35 | $1.75 \%$ | 83.3 |

The number of naturally migrating fry early in the season was such that we could have made additional releases with fry. However, at the direction of the CDFG we limited the releases to approximately once per week.

We were unable to make releases of naturally migrating parr and smolts during April through June due to low catches in the Caswell traps. In order to conduct a trap efficiency release, we needed approximately 500 marked fish. This would have given only 10 recaptured fish at an estimated $2 \%$ efficiency. Although we can save trap catches for several days to accumulate 500 fish, at the rate we were catching parr and smolts it would have taken well over one week. Saving natural fish for more than about 5 days results in stressed fish that are unhealthy at release, and therefore not representative of the natural population.

Due to the few fish available for trap efficiency releases, we did not release special groups to test differences between day and night releases or release location. Results in previous years indicated that the release location in the channel did not play a significant role in determining trap efficiency, and due to extremely low recapture rates day releases were determined to be ineffective in accurately determining trap efficiency.

## Effect of Mean Length on Trap Efficiency

To determine the effect of fish length on trapping efficiency, we compared the mean lengths of fish released to the mean lengths of fish recaptured for groups of marked fish used to test trapping efficiency. An indicator of possible sampling bias is whether or not the size of recaptured fish differed substantially from the mean of all released fish. Normally, there was less than twenty-four hours between release and recapture, so any detected difference in mean lengths would be associated with either a size-dependent sampling rate, or a size-dependent pre-recapture mortality rate. There would be insufficient time between release and recapture to result in any substantial change in the size of the fish.

There were not enough releases in a given size category in 1999, or within any of the years sampled at Caswell, to permit powerful comparisons of lengths; therefore, the 1996 through 1999 releases were pooled within size category. We found no substantial or significant difference between the lengths of released and recovered fry $(P=0.54)$. The recovered parr and smolts averaged more than 1 mm smaller than the released fish ( $\mathrm{P}=$ 0.0178 for parr and $P=0.0651$ for smolt). However, there were no parr releases in 1999, and the recovered smolts were actually larger than released smolts for all four smolt releases made in 1999. There may well be a size bias, but the nature of the size bias may differ over different years' river conditions. Such a size bias would bias the expansion of the catch if the size distribution of the released fish differed from that of the river-run passage. If there were such a difference, it would be difficult, if not impossible, to adjust for potential bias.

## 1999 Capture Efficiency Model

The daily outmigration index was calculated by dividing the number of chinook captured at Caswell each day by the predicted daily trap efficiency (proportion of released fish that were later recaptured):

Outmigration Index $\quad \frac{\text { Count }}{\text { Efficiency }}$

Daily counts from the two screw traps were available from February 6 through July 1, $199{ }^{1}$, from March 19 through June 27, 1997, from January 29 (although the trap was installed January 8) to July 16, 1998 (hereafter referred to as passage days), and from January 18 to June 30, 1999. On 25 dates during these monitoring periods for the 4 years combined, a total of 43 uniquely marked releases ${ }^{1}$ were made at a fixed distance upstream

1 The number of efficiency releases were: In 1996, 1 on Feb 14, 1 on Feb 19, 1 on March 22, 4 on April 6, 2 on May 2, 2 on May 10, 2 on May 26, and 2 on June 10; in 1997, 1 over a period from April 7 through 11 (denoted as April 9, mid-point day) and 4 releases on the night of May 28/29; and in 1998, 3 on March 14, 3 on March 25, 4 on April 18, 3 on May 10, 2 on May 18, 2 on June 4, and 2 on June 12. Of the 39 releases, 5 were day-time and 34 were night-time releases. from the Caswell screw traps for the purpose of estimating trap efficiency.

Trap efficiency releases were made in the same location, using the same release procedures, and within similar flow ranges in all years. Among-year differences in the variation in trap efficiency adjusted for fish size, flow, and turbidity did not differ substantially or significantly, allowing a combined analysis to estimate deviance (a measure of variability, see Appendix A). Combining data from all years enabled more precise estimation of efficiency rates for time periods when tests were not conducted.

## Developing the 1999 Model

In order to predict the efficiency for each passage day, the efficiency estimates were related as a response (dependent variable) to predictor variables (independent variables) that were measured on every day that the screw traps were operating. The predictor variables explored were flow (f) (in cubic feet per second, cfs) measured at OBB, fish size (s) (in millimeters, mm), and turbidity ( t ) (in nephelometric turbidity units, NTU). Efficiency (e), the proportion of released fish trapped per release, was related to the predictor variables using the logistic relation:
efficiency (e) $\quad$. $\frac{1}{1 \% \exp ^{[\delta b(0) \delta b(f)(f \delta b(s)(s \delta b(t)(t]}}$
or, using the "logit" linear transform,

$$
\text { logit }(e) \quad \ln \left[\frac{e}{1 \& e}\right] \quad b(0) \% b(f)(f \% b(s)(s \% b(t)(t \% b(s)(s
$$

In the above equations "exp" is the exponential function, "ln" is the natural log, "b(0)" is a

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September 2000 coefficient associated with the intercept ${ }^{2}$, and $b(f), b(s)$, and $b(t)$ are partial logistic regression coefficients relating the logit transform of efficiency to the indicated predictor variables. A major reason for choosing the logistic model is that the predicted efficiency can never be less than 0 or greater than 1 (100\%), but the logit transform is continuous from negative to positive infinity.

## Model Selection

A detailed description of the regression model fitting procedure is given in Appendix A, and a summary is presented here. Separate multiple regressions were calculated to predict trap efficiency in each year (1996-1997, 1998, 1999). A regression coefficient to account for the effect of turbidity could only be calculated with 1996 data, because that was the only year in which efficiency tests were completed on dates when turbidity exceeded 10 NTU (Appendix A). A strong effect of turbidity on trap efficiency was only apparent when turbidity exceeded 10 NTU. Although turbidity showed a significant influence on trapping efficiency in 1996, that influence was not measurable in other years due to relatively low turbidities (<10 NTU) for the dates on which efficiency was tested by releases of marked fish. However, turbidities did exceed 10 NTU for some days during the outmigration in all years. Therefore, the regression coefficient estimated for turbidity in 1996 was retained in the multiple regressions for other years, but it was only applied on days when turbidity exceeded 10 NTU.

The best fitting multiple regression on trap efficiency (logit transformed) was estimated separately for each year by a back-step procedure. Initially the predictor variables of turbidity, flow, and fish length were all included in the regressions for each year. Each regression was then reduced by dropping the predictor variables that did not contribute significantly ( $\mathrm{P}>0.2$ ) to the regression fit. The reason for choosing such a high

2 Intercept value $=1 /\left\{1+\exp ^{-b(0)}\right\}$ when $\mathrm{f}=\mathrm{s}=\mathrm{t}=0$. significance level was to reduce the chance of omitting a coefficient when it should be included (Type 2 error). During the back-step procedure, the turbidity regression coefficient estimated from 1996 data was retained in the regression for each year. The final regression model for 1999 differed from previous years in that it included only the intercept and turbidity (Table 4). The net result was that trap efficiency was estimated to be $2.6 \%$ on all days with turbidity < 10 NTU, and was estimated to increase at turbidities > 10 NTU. The final regression model for 1996-1997 included all three predictor variables, and the model for 1998 included turbidity and fish length (Table 4).

The 1999 analysis and model selection differed from previous years. Previously flow, fish size, and turbidity were all included in the regression for each year, whether or not the coefficients differed significantly from 0 in any given year. Previously, if a predictor variable was significant in at least one year, it was included in the regressions for all years. The reason for the inclusion of non-significant coefficients was that flow and turbidity coefficients had the same sign each year, but were only significant in 1996 and 1997. Conversely, fish length was a significant predictor variable in 1998, but not in 1996 or 1997 (although the coefficients had the same sign as the 1998 coefficient). Because of the sign consistency, the non-significant coefficients were retained to avoid the bias that could have resulted from statistically declaring the coefficient to be 0 , when in fact it was not (Type 2 error). In 1999, however, signs differed from the previous year's coefficients (Appendix 2.a). This finding prompted us to drop non-significant coefficients from all years. However, to reduce the chance of omitting a coefficient when it should be included, the significance level chosen was high ( $P=0.2$ ).

Table 4. Coefficient estimates and associated statistics for the regressions on trap efficiency.

1996-1997
Efficiency Predictor: er = 1/\{1+exp[-b(0)-b(f)*f-b(s)*s-b(t')*t']\}

| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio <br> (b/SE) | $\mathbf{P}$ <br> (Type I) |
| :---: | :---: | :---: | :---: | :---: |
| Intercept (0) | -1.423720 | 0.476060 | -2.99 | 0.0050 |
| Flow (f) | -0.000829 | 0.000167 | -4.97 | 0.0000 |
| Recovery Size (s) | -0.010380 | 0.005237 | -1.98 | 0.0551 |
| Turbidity $>10(\mathrm{t})$ | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | Deviance | D.F. | Deviance/D.F. |  |
|  | 163.47 | 36 | 4.54 |  |
|  |  | Variance-Covariance |  |  |
| Intercept | Flow | Recovery Size | Turbidity > 10 |  |
| Predictor | Intercept | $2.2663 \mathrm{E}-01$ |  |  |
| Flow | $-4.3819 \mathrm{E}-05$ | $2.7835 \mathrm{E}-08$ |  |  |
| Recovery Size | $-2.1342 \mathrm{E}-03$ | $6.5842 \mathrm{E}-08$ | $2.7427 \mathrm{E}-05$ |  |
| Turbidity $>10$ | $-3.4783 \mathrm{E}-03$ | $-3.7326 \mathrm{E}-07$ | $4.6771 \mathrm{E}-05$ | $2.0479 \mathrm{E}-04$ |

1998
Efficiency Predictor: er $=1 /\left\{1+\exp \left[-b(0)-b(s)^{*} s-b(t) * t\right]\right\}$

| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio <br> (b/SE) | P <br> (Type I) |
| :---: | :---: | :---: | :---: | :---: |
| Intercept $(0)$ | -2.251260 | 0.301509 | -7.47 | 0.0000 |
| Recovery Size (s) | -0.022160 | 0.004855 | -4.56 | 0.0001 |
| Turbidity $>10(\mathrm{t})$ | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | Deviance | D.F. | Deviance/D.F. |  |
|  | 163.47 | 36 | 4.54 |  |
|  |  | Variance-Covariance |  |  |
|  | Intercept | Recovery Size | Turbidity > 10 |  |
| Predictor | $9.0907 \mathrm{E}-02$ |  |  |  |
| Intercept | $-1.3759 \mathrm{E}-03$ | $2.3567 \mathrm{E}-05$ |  |  |
| Recovery Size | Turbidity $>10$ | $0.0000 \mathrm{E}+00$ | $0.0000 \mathrm{E}+00$ | $2.0479 \mathrm{E}-04$ |

1999
Efficiency Predictor: er $=1 /\left\{1+\exp \left[-b(0)-b\left(t^{\prime}\right)^{*} t^{\prime}\right]\right\}$

| Predictor | Estimate <br> $(\mathbf{b})$ | Standard <br> Error (SE) | t-ratio <br> (b/SE) | P <br> (Type I) |
| :---: | :---: | :---: | :---: | :---: |
| Intercept (0) | -3.624670 | 0.106971 | -33.88 | 0.0000 |
| Turbidity $>10(\mathrm{t})$ | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | Deviance | D.F. | Deviance/D.F. |  |
|  | 163.47 | 36 | 4.54 |  |
| Variance-Covariance |  |  |  |  |
| Predictor | Intercept | Turbidity > 10 |  |  |
| Intercept | $1.1443 \mathrm{E}-02$ |  |  |  |
| Turbidity $>10$ | $0.0000 \mathrm{E}+00$ | $2.0479 \mathrm{E}-04$ |  |  |

## Outmigration Indices

The predicted daily efficiency was used to expand the daily count to estimate a daily passage index. Daily passage indices were substantially greater during late January through February than for the remainder of the outmigration season (Figure 4). The fish migrating at that time were fry. Indices reached a daily peak of over 85,000 fry on February 12 (Figure 4). Outmigrant abundance during February was more than 10-times greater than during March through June. The daily indices were summed over the dates when fry, parr, and smolt were passing (Table 5), to obtain the season's outmigration index estimates for each of these life stages.


Figure 4. Daily outmigration index of juvenile chinook and flow of the Stanislaus River for 1999.
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Table 5. Estimates of outmigration indices for fry, parr, and smolts, 1996 through 1999, based on efficiency estimates from regressions in Table 4. Daily catches and abundance indices are presented in Appendix A.

1996 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits Lower Upper |  | 1998 Report Data Summary Date Domain Estimate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. |  |  |  |  |  |
| Fry | 02/06 | 03/20 | 31,767 | 10,320 | 11,540 | 51,994 | 02/07 | 03/20 | 28,653 |
| Parr | 03/21 | 03/31 | 1,596 | 470 | 675 | 2,516 | 03/22 | 03/31 | 1,465 |
| Smolt | 04/01 | 07/01 | 81,896 | 11,065 | 60,209 | 103,582 | 04/02 | 07/01 | 65,083 |
| All | 02/06 | 07/01 | 115,258 | 15,051 | 85,759 | 144,757 | 02/07 | 07/01 | 95,201 |

1997 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits Lower <br> Upper |  | 1998 Report Data Summary Date Domain Estimate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. |  |  |  |  |  |
| Fry Parr | 03/19 | 04/05 | 7,011 | 1,037 | 4,979 | 9,043 | 03/20 | 04/01 | 4,724 |
| Smolt | 04/06 | 06/27 | 60,333 | 7,478 | 45,676 | 74,990 | 04/03 | 06/27 | 48,861 |
| All | 03/19 | 06/27 | 67,344 | 8,000 | 51,663 | 83,024 | 03/20 | 06/27 | 53,585 |

1998 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits |  | 1998 Report Data Summary |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date | omain | Estimate | S.E. | Lower | Upper | Date D | main | Estimate |
| Fry | 01/29 | 03/07 | 186,029 | 44,908 | 98,009 | 274,049 | 01/30 | 03/07 | 287,801 |
| Parr | 03/08 | 04/23 | 209,911 | 31,238 | 148,685 | 271,137 | 03/09 | 04/21 | 179,448 |
| Smolt | 04/24 | 07/16 | 197,884 | 37,348 | 124,682 | 271,087 | 04/23 | 07/16 | 183,935 |
| All | 01/29 | 07/16 | 593,825 | 76,373 | 444,133 | 743,516 | 01/30 | 07/16 | 651,184 |

1999 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. | Lower | Upper |
| Fry | 01/18 | 03/15 | 1,155,424 | 145,284 | 870,668 | 1,440,181 |
| Parr | 03/16 | 05/09 | 92,615 | 11,169 | 70,723 | 114,506 |
| Smolt | 05/10 | 06/30 | 73,003 | 9,679 | 54,031 | 91,975 |
| All | 01/18 | 06/30 | 1,321,042 | 160,000 | 1,007,443 | 1,634,642 |

Note that the estimates for past years (Table 5) differ from those presented in the 1998 report. These differences resulted from the revised regressions on efficiency and some data corrections (see Appendix A). The previous estimates fall within the new confidence limits, which are generally narrower than in previous reports. This is because the variance estimate has been improved as described in Appendix A.

The estimated fry (<45mm) outmigration in 1999 was substantially greater than in 1998 (the previous greatest estimate) (Table 5 and Figure 5). The season totals for fry are not directly comparable between any years because sampling started on different dates and fry were already migrating when sampling began. Sampling began earlier in 1999 and fry passage continued until a later date in 1999 than in 1998 (1999 fry outmigration: January 18 - March 15; 1998 fry outmigration: Jan 29 - March 7), so the contrast in fry outmigration between years in Table 5 is exaggerated. However, daily fry indices exceeded 20,000 on 28 days in 1999, but only once in 1998. Clearly, substantially more chinook migrated as fry in 1999 than during the previous 2 years when fry were sampled (1996, 1998).

Unlike fry, the 1999 parr and smolt outmigrants were less abundant than in 1998. In fact, 1999 parr and smolt confidence intervals do not overlap those of 1998, indicating that the 1999 parr and smolt outmigration is truly less than that of 1998 (Table 5). Direct comparison of parr and smolt abundance are appropriate between years (excluding 1997 for parr) because the full outmigration period was sampled in each year from 1996 to 1999.

The parr outmigration of 1999 was substantially greater than in 1996 and 1997 and less than that of 1998. However, the differences were not as marked for smolts. In 1999, there was a 2 week period when the mean length was very near 80 mm , and the 80 mm demarcation between parr and smolt might be somewhat artificial, especially for the 1999 outmigration (see Appendix Table A.5.d). Because of the indistinct separation of parr and smolts in some years, there may be some benefit to combining the estimates of parr and smolts to compare abundance among years.


Figure 5. Estimates of total outmigration index by life stage classification from the Stanislaus River, 1996-1999.

## Oakdale and Caswell Passage Estimates

Outmigration indices from the Oakdale and Caswell sites were compared to evaluate differences in passage estimates and migration timing for different life stages of chinook (Figure 6). Sampling at the two locations in 1999 covered the same dates, so direct comparisons between the sites may reflect 1) additions or losses that occurred between the sites, 2) rearing between the sites, or 3) lack of precision in passage estimates. Overall, more fish were estimated to have passed the Oakdale trap $(1,669,000)$ than the Caswell trap $(1,321,042)$ (Table 6). Most of the difference in passage between the two sites was in parr (275,748 less parr at Caswell) (Figure 7).



Figure 6. Estimates of the daily passage index for juvenile chinook at Oakdale and Caswell in 1999. Data for Oakdale from Demko et al. (2000).

Table 6. Estimates for the outmigration index of fry, parr and smolts at Oakdale and Caswell in 1999. Oakdale data from Demko et al. (2000).

|  | Caswell 1999 |  |  |  | Oakdale 1999 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate |  | Date Domain | Estimate |  |
| Fry | $1 / 18 / 99$ | $3 / 15 / 99$ | $1,155,424$ |  | $01 / 18 / 99$ | $03 / 22 / 99$ | $1,198,144$ |
| Parr | $3 / 16 / 99$ | $5 / 9 / 99$ | 92,615 |  | $03 / 23 / 99$ | $05 / 26 / 99$ | 368,363 |
| Smolt | $5 / 10 / 99$ | $6 / 30 / 99$ | 73,003 |  | $05 / 27 / 99$ | $06 / 30 / 99$ | 102,493 |
| All | $\mathbf{1 / 1 8 / 9 9}$ | $\mathbf{6 / 3 0 / 9 9}$ | $\mathbf{1 , 3 2 1 , 0 4 2}$ |  | $\mathbf{0 1 / 1 8 / 9 9}$ | $\mathbf{0 6 / 3 0 / 9 9}$ | $\mathbf{1 , 6 6 9 , 0 0 0}$ |



Figure 7. Comparison between Caswell and Oakdale of fry, parr, and smolt passage indices during 1999. RIVER

## Length at Outmigration

The mean lengths of juvenile chinook gradually increased over the course of sampling, ranging from about 35 mm at the start of sampling (mid-January) to about 90 mm in late June (Figure 8). The gradual increase in mean lengths over time in 1999 was similar to the pattern seen in 1996, 1997, and 1998; however, the length increase was slower than other years and the threshold size for classifying smolts ( 80 mm ) was not reached until May 10 (Figure 9). This was 3 to 7 weeks later than other years. Environmental factors such as water temperature, turbidity, and habitat availability, as well as fish abundance, may have played roles in determining growth rates and the lengths of migrating juvenile chinook. Late spawners could have contributed to smaller fish seen later in the season. The above factors will be addressed in the following section.

## Comparison of Mean Lengths of Natural Migrants Between Years

Mean lengths of fry captured at Caswell in January and February were similar for all years sampled (1996-1999). However, at the onset of capturing parr in March, the mean lengths diverge between the years.


Figure 8. Maximum, minimum, and mean lengths of chinook captured daily at the Caswell trap in 1999.


Figure 9. Mean lengths of chinook captured daily at the Caswell trap 1996-1999

Comparison of rearing temperature data at Goodwin Dam for 1998 and 1999 revealed that slightly cooler temperatures prevailed during January, February, and March of 1999 (Figure 10). This may have contributed to the delayed emergence and slower growth of juvenile chinook. However, temperatures in 1996 were similar to those in 1998 while 1996 fish were larger than 1998 fish (Figure 10). Thus, temperature alone does not explain the differences in length between years.


Figure 10. Average monthly temperatures of the Stanislaus River at Goodwin Dam 1996 thru 1999.

The combination of temperature and juvenile chinook abundance may account for a large share of the variation in juvenile growth between years, but a longer time series of data will be necessary to confirm this. Chinook, like many salmonids, are highly territorial (Chapman 1966, Elliot 1990) and as they grow their territory expands (Elliot 1990). With their growth and subsequent expansion of territories, fish would inevitably continue to be displaced downstream in search of unoccupied habitats. In spite of the early exit of a majority of the population (fry migrants), the remaining fish would still compete for resources. The largest mean lengths at Caswell by the first week of May were collected in 1996. The year of lowest fish abundance and the smallest mean lengths the first week of May was 1999, also the year of lowest temperatures. Cramer et al. (1985) found after 10 years of study in the Rogue River that $85 \%$ of variation between years in mean lengths of juvenile chinook could be accounted for by regression on juvenile abundance and river temperature. In that study, juveniles grew faster when spring river temperatures were high and when juvenile chinook abundance was lower.

Flow may also influence growth indirectly through its influence on habitat availability or capacity. As flow increases, new areas could be inundated. At bankfull stages fish would have access to productive rearing habitats in the flood plain, more cover, and more total area for rearing (Yoshiyama et al. 1998). However, there is no indication in our years of data that lengths are larger in higher flow years. Further, there is generally a negative correlation between flow and river temperature during spring. Higher water years result in cooler river temperatures, which in turn can slow growth rates. This was the finding after 10 years of study on the Rogue River where growth rates of juvenile chinook tended to be highest in years of lowest flow (Cramer et al. 1985). However, Cramer et al. (1985) concluded from a variety of growth measures that warmer temperatures, rather than lower flows, were driving growth of juvenile chinook.

Spawner timing and abundance could affect the lengths of fish observed in the spring. Eggs spawned early would accumulate temperature units quickly while temperatures were still warm in the fall. Eggs from later spawners would be subject to a cooler temperature regime from the start. Fish exposed to cooler temperatures would incubate slower and emerge later, thus elongating the fry portion (and possibly parr and smolt stages) of the outmigration. This might result in the more gradual increase in lengths observed in 1998 and 1999. At this point spawner data has not been incorporated into this report, but should be considered in future studies.

Turbidity can play a role in growth of juvenile chinook (Gregory and Northcote 1993). It has been suggested that predation is reduced under turbid conditions (Gradall and Swenson 1982, Cezilly 1992, Gregory 1993 cited from Gregory and Levings 1998) and juvenile fish may engage in activities such as increased feeding activity that would otherwise be risky (Ginetz and Larkin 1976). Turbidity levels in 1996, 1998, and 1999 (turbidity was excluded for 1997 due to the relatively short sampling period) differed slightly between years, but generally only varied within a range of 3 to 10 NTU (Figure 11). Gregory and Northcote (1993) found highest foraging rates of juvenile chinook when they were in water with turbidities ranging form 35 to 150 NTU. We found no indication from available studies that the low range of turbidities we observed would have a detectable effect on chinook growth.


Figure 11. Daily turbidity of the Stanislaus River at Caswell for 1996, 1998, and 1999.

## Comparison of Mean Lengths at Oakdale and Caswell in 1999

Mean lengths of natural chinook captured at Caswell and Oakdale have been similar throughout the trapping season in past years. In 1999 there was a noticeable difference in mean lengths between the sites beginning in March, once fish reached the parr stage (Figure 12). The difference was greatest in mid April, when fish reached approximately 75 mm at Caswell, but were still near 60 mm in length at Oakdale.


Figure 12. Comparison of daily mean lengths for juvenile chinook at Oakdale and Caswell in 1999.

The difference in lengths between the trap sites during 1999 may indicate that many parr paused to rear between Oakdale and Caswell. There was no sampling between the two sites to test this possibility, but data from our previous studies tend to support the rearing hypothesis.

Although mean lengths were nearly the same between the two sites in previous years, the relative abundance of migrants that were either fry, parr, or smolts differed between the trap sites. Most notably, the abundance of both parr and smolts was greater at Caswell than at Oakdale in 1998 (Figure 13). Because only a small percentage of spawning occurs below Oakdale (roughly 20\%), the higher number of parr and smolts passing Caswell in 1998 would most likely have resulted from fry that migrated past Oakdale and then stopped to rear between Oakdale and Caswell. In order to make comparisons of relative fry abundance between stations and between years, we used the index of daily fry passage summed for the last 40 days of fry passage at each station (because we had at least 40 days of fry sampling at each station in each year). This 40 day index of fry abundance is the quantity shown in Figure 13. Data shown in Figure 13 indicates that the excess fry passing Oakdale compared to Caswell in 1998 must have then reared below Oakdale and later migrated to produce the excess parr and smolts at Caswell compared to Oakdale. Our finding that mean lengths of juvenile chinook on any given date in 1998 were similar between the two stations indicates that growth rates must have been similar between the two areas of the river.


# Standardized Passage Estimates by Lifestage at Caswell 



Figure 13. Comparison of passage indices between Oakdale and Caswell for fry, parr, and smolts during 1996, 1998, and 1999. Fry index is standardized for the last 40 days of fry passage in each year.

Chinook parr show habitat preferences for low velocities, small substrate particle size, and abundant cover (when available) (Chapman and Bjorrn 1969, Everest and Chapman 1972, Murphy et al. 1989 cited from Healey 1991). Such cover is frequently found in back eddies, behind fallen trees, and near undercut tree roots and other areas of bank cover (Healey 1991). These habitat characteristics are plentiful in the lower Stanislaus River during the spring. Relative to the river above Oakdale, the river downstream has temperatures during spring that may be more optimal for growth of juvenile chinook in some years, and higher turbidities that provide some cover from predation. However, warm water piscivores such as largemouth bass, smallmouth bass, and striped bass are more abundant below Oakdale. Overall, there may be survival advantages to juvenile chinook rearing between Oakdale and Caswell, but this is uncertain.

## Smolt Appearance Index

The external appearance of smolt characteristics among chinook captured in the traps started to increase at the beginning of March (Figure 14), when the daily mean smolt index gradually increased from 1 to 2 . Individual fish with a score of 2 appeared through mid-June and their lengths ranged up to 90 mm (Figure 14). Fish that were distinctly smolts (index $=3$ ) were at least 80 mm and began appearing in mid-April. (Figure 14).

The smolt appearance index followed similar trends through time in 1998 and 1999. Fry were present through the end of April in both 1998 and 1999, with the exception of 1 fry captured in late May in 1999. Parr appeared later in 1999 (beginning of March) than in 1998 (late February), but in both years parr persisted until mid June. The difference in the timing of parr could be attributed to a variety of factors affecting growth and development, as discussed previously. In both years smolts (smolt index 3) were observed around the $15^{\text {th }}$ of April through the end of sampling at the end of June.


Figure 14. Mean daily smolt appearance index and mean length of natural chinook captured in the Caswell screw traps during 1999.

## Rainbow Trout / Steelhead Lengths

During the sampling season, we captured 12 rainbow trout/steelhead at Caswell, ranging in size from 83 mm to 255 mm (Figure 15). Two distinct size classes were apparent (200-300 mm and <100mm), representing yearlings and young-of-the-year, respectively. More rainbow trout/steelhead were captured in 1999 than 1998, but 1999 counts were comparable to 1997.


Figure 15. Lengths of all rainbow trout/steelhead captured at Caswell 1995 through 1999.

## OBJECTIVE 3: IDENTIFY FACTORS THAT INFLUENCE THE TIMING, SIZE, AND NUMBER OF JUVENILE CHINOOK SALMON AND RAINBOW TROUT/STEELHEAD MIGRATING OUT OF THE STANISLAUS RIVER

## Effects of Streamflow on Chinook Salmon Outmigration

There was no clear indication that variation in flow stimulated fry movement, but the magnitude of flow may have encouraged fry movement. Similar to 1998, the peak of chinook fry passage in 1999 corresponded with elevated flows (Figure 16). Heavy rains resulted in elevated flows for most of January and February, and fry migration peaked on February 12, when over 80,000 fry were estimated to pass Caswell in a single night. Fry passage estimates were high throughout late January and all February, while flows generally ranged between 3,000-4,000 cfs.

Streamflow has previously been correlated with peak fry catches. In the Sacramento-San Joaquin delta, Kjelson et al. (1981) found peak fry catches were often associated with flow increases caused by storm run-off. They speculated flow pulses stimulated fry migration out of the upper river spawning grounds. The correlation between flow and fry movement was also observed in the Nanaimo River (Healey 1980).

Migration peaked in mid-February during 1996, 1998, and 1999 (fry outmigration was not sampled in 1997). Each peak was also associated with an increase in daily average flow of 300 to 700 cfs. However, smaller peaks were not associated with flow increases. Flow may influence fry migration, but its affects are not consistent, indicating that flow is not the only factor driving outmigration timing.


Figure 16. Daily index of fry passage at Caswell and flow for 1999.

It is possible that the magnitude of flows during the fry outmigration may play a role in determining the proportion of the total population that will migrate as fry. If flows during the onset of fry emergence are stable, more fry may establish feeding behavior and territories before they drift downstream with high flows. This may account for the low proportion of fry migrants in low flow years such as appeared to be the case in 1996. Other studies in the San Joaquin Delta tend to support the theory that a higher percentage of juvenile chinook migrate as fry in high flow years. For example, USFWS (1998) found that the abundance of chinook fry captured by seining in the northern Delta between 1985 and 1999 was positively correlated ( $r=0.91$ ) with the mean flow of the Sacramento River during February.

Between March and June, parr and smolt outmigrated without any dramatic peaks in migration, and flows were relatively stable between 1,200 cfs and 1,600 cfs (Figure 17). Unlike the fry outmigration, no relationship between outmigration and flow was observed for parr and smolts. Similar results were observed on the upper South Umpqua River basin, Oregon, where migration of 50-59 mm chinook (parr) was not cued by changes in discharge (Roper and Scarnecchia 1999). The relatively stable flows during the spring of 1999 provided little opportunity to observe the influence of flow variation on parr and smolt migration.


Figure 17. Daily index of parr and smolt passage at Caswell compared to flow for 1999.

## Effects of Turbidity on Chinook Salmon Outmigration

Fry outmigrated during January and February when turbidity levels were high, ranging from approximately 4 to 24 NTU's. Although there were two dramatic spikes in turbidity lasting 1 day in both January and February, one slightly over and one slightly under 24 NTU's, daily turbidity generally ranged between 4 and 10 NTU's. The highest fry passage day did occur two days after the second turbidity spike (just under 24 NTU's on February 10), but there was no apparent migration change to the first spike. Turbidity ranged only from 2 to 7 NTU's during the parr and smolt outmigration, so there was little opportunity to observe the effect on parr and smolt migration. The limited variation in turbidity did not correspond with any peaks in passage of parr or smolt (Figure 18).


Figure 18. Daily chinook passage at Caswell compared to turbidity in 1999.

Little work has been done on the relationship between turbidity and fry outmigration timing; however, many studies have related turbid conditions to reduced fry predation. Predators, such a birds and fish, use vision to detect and attack prey. High turbidity can impair visual abilities, thus reducing the detection range of predators. A study by Gregory and Levings (1998) concluded predation did occur at lower rates in the Fraser River (27108 NTU) compared to that of the clear water Harrison River. We might expect then that turbid water acts as cover for small fish, allowing more of them to outmigrate undetected. Previous studies have suggested juvenile fish engage in otherwise dangerous activities during turbid conditions, activities including feeding (Gregory and Northcote 1993, Gregory 1994), increased use of open water (Miner and Stein 1996), increased migration rate (Ginetz and Larkin 1976), and less cover seeking behavior (Gradall and Swenson 1982, Gregory 1993). These responses might encourage juvenile chinook to migrate during turbid conditions.

## Effects of Lunar Phase on Chinook Salmon Migration

Lunar phase has been correlated to peaks in fry and smolt movement in some streams. Reimers (1971) observed that downstream movement of chinook was inhibited by bright moonlight (cited from Healey 1991). Grau et al. (1981) found that thyroxine levels, which are associated with smoltification, peaked during the new moon phase in anadromous salmonids. They postulated migration during nighttime hours and during the darkest nights of the month made detection by predators more difficult, thus increasing smolts chances for survival.

Caswell passage estimates did not consistently peak during the new moon phases (Table 7), but there was indication of an increase in passage at that time (Figure 19). Proximity to the spawning grounds can effect the patterns observed. For example, fish choosing to migrate from spawning grounds might not reach Caswell for 3 days to a week depending on flow conditions. Therefore, we might expect lunar phase patterns to be more evident at the Oakdale trap, which is closer to spawning grounds than the Caswell trap. Additionally, weather conditions may also play a role. Overcast or cloudy skies could filter or completely block moonlight, thus producing darkness such as experienced during a new moon.

Table 7. Dates each month for peak passage and new moon from January thru May in 1996-1999. Months in which peak passage and new moon occurred within 5 days of one another are highlighted.

|  | 1996 |  | 1997 |  | 1998 |  | 1999 |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Month | Peak* $^{*}$ | New Moon | Peak* $^{*}$ | New Moon | Peak* $^{\text {New Moon }}$ | Peak* $^{\text {New Moon }}$ |  |  |
| January | 25 | 17 | - | - | 29 | 28 | 23 | 20 |
| February | 12 | 16 | - | - | 16 | 26 | 23 | 18 |
| March | 2 | 17 | 30 | 9 | 26 | 28 | 1 | 19 |
| April | 16 | 16 | 15 | 7 | 8 | 26 | 30 | 17 |
| May | 21 | 15 | 16 | 6 | 19 | 25 | 5 | 17 |
| *Peak $=$ peak passage of smolts from outmigration index estimates. |  |  |  |  |  |  |  |  |



Figure 19. New moon phase and daily chinook passage estimates for 1999.

## Effects of Temperature on Chinook Salmon Outmigration

Daily fluctuations in outmigration did not appear to correspond with changes in temperature. Temperature at Caswell gradually increased from around $10^{\circ} \mathrm{C}$ at the beginning of sampling to 19EC at the end of June (Figure 20). It is likely that temperature does affect seasonal migration timing, because it strongly affects growth and development. The timing of chinook smolt migration was significantly related to temperature on the South Umpqua River (Roper and Scarnecchia 1999). Migration was later in cool years and earlier in warm years on the South Umpqua River. Cooler temperatures during winter slowed development resulting in later emergence dates whereas warmer temperatures accelerated growth and promoted earlier emigration (Roper and Scarnecchia 1999). Water temperature was also seen to coincide with fish movement in a study conducted by Bjornn (1971) but he did not find a consistent relationship and concluded photoperiod and growth were more likely initiating movement. Bjornn's results (1971) are consistent with those from the Stanislaus in that we found a coincidence of increases in outmigration and temperature change (up or down), but not consistently. Other studies have found that fish will migrate under constant temperature regimes (Bjornn et al. 1968 cited from Bjornn 1971).


Figure 20. Daily chinook passage at Caswell, compared to river temperature during 1999.

## Effects of Incubation Temperature on Fry Migration Timing

Incubation temperature during fall and early winter plays a key role in the development of chinook and their subsequent time of emergence. The warmer the water, the faster the development of eggs. Chinook eggs on average take 780-814 degree days (EC) (Healey 1991) to incubate to hatching, and $\sim 890$ degree days (EC) to fry emergence from the gravel (USFWS 1998).

Water temperatures can be used to predict the start of chinook fry emergence. Temperatures at Goodwin Dam were recorded and used to perform a simple degree day analysis to estimate when fry first emerged. Degree days are the sum of the average degrees above freezing each day during incubation. Spawning was estimated to have begun at Knights Ferry around October 15, 1998 (personal communication, Duane Johnson, U.S. Army Corps of Engineers). Given this date, we summed average daily temperature until we achieved 888 degree days (CE) (literature value for emergence). The start of emergence was estimated to have begun on January $5^{\text {th }}$. This is consistent with trapping data. Traps were not fished until January $18^{\text {th }}$, at which time passage estimates already exceeded 10,000 fish/day. Therefore the start date of January $5^{\text {th }}$ is not an unlikely estimate; however, it does suggest that traps should be installed at an earlier day if the entire run is to be sampled. Also, more detailed information on the temporal distribution of spawning would enable more accurate estimates of emergence timing.

## Effects of Size on Timing of Migration

The variation in peak fry emergence among years did not appear to relate to fish size, as newly emerged fry were consistently $35-37 \mathrm{~mm}$ each year. This is at the low end of the ranges found for other populations (Mains and Smith 1964, Lister et al. 1971, Healey et al. 1977 cited from Healey 1991).

No relationship of parr/smolt lengths to migration timing could be discerned in the 1999 data (Figure 21).


Figure 21. Daily mean length and chinook passage estimate at Caswell for 1999.

## OBJECTIVE 4: ESTIMATE THE SURVIVAL OF CODED WIRE TAG RELEASES FROM KNIGHTS FERRY AND OAKDALE RECREATION AREA TO CASWELL STATE PARK IN 1999

## Coded Wire Tagged Chinook Released in 1999

In 1999, we cooperated with California Department of Fish and Game to mark and release approximately 45,000 CWT hatchery chinook near Oakdale and Knights Ferry in an attempt to evaluate their survival rates through different segments of the river. Oakdale releases were divided equally between one release point above and one below the Oakdale Recreation Area (RM 38 and 40) to determine if mortality was disproportionately
high as fish passed through the recreation area ponds. Previous radio tracking studies suggested that predation of chinook was high in the ponds (Demko et. al 1998). The Knights Ferry release was conducted to estimate survival for both the upper river (Knights Ferry to Oakdale) and the mid-section of the Stanislaus (Oakdale to Caswell). Results and survival estimates from this experiment are presented in Appendix D.

## Migration Rates

Average migration rates through the Stanislaus River were estimated for marked fish released at Knights Ferry and Oakdale that were subsequently recovered in the Caswell trap. Average migration rates for groups with multiple recoveries varied from 3.9 to 13.8 miles $/$ night among the different release groups (maximum $=23.0$ miles $/ \mathrm{night}$, minimum=1.3 miles/night) (Table 8). Most recoveries from the CWT groups released at Oakdale arrived at Caswell 2 nights later, while the greatest number of recoveries from the Knights Ferry release came 3 nights later. Thus, most fish migrated roughly 15 miles per night, while a few moved much slower. Of the 40 fish recaptured at Caswell from the upstream releases, $15 \%$ took more than two weeks to travel from Oakdale to Caswell. This supports the hypothesis that many juveniles reared between Oakdale and Caswell.

We found no obvious relationship of migration rate to either flow or fish size. One might expect fish to migrate faster during high flows or at larger sizes. In all probability, these relationships do exist but are obscured by data collected from fish opting to rear and extend their residence in the river.
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Table 8. Recapture data at the Caswell trap used to calculate migration rate of juvenile chinook from Oakdale Recreation, and from Oakdale or above down to Caswell, 1999. The groups titled "Oakx" were released at Oakdale to evaluate trapping efficiency and individual fish recaptured at Caswell.

| Night | KF-CDFG | RM38-1 | RM38-2 | RM40-1 | RM40-2 | Oak1 | Oak2 | Oak3 | Oak4 | Oak9 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 |  |  |  |  |  |  |  |  |  |  |
| 2 | 1 | 7 | 6 | 7 | 6 | 1 |  |  |  |  |
| 3 | 16 | 1 | 1 | 1 |  |  |  |  |  |  |
| 4 | 8 |  |  |  |  |  |  |  |  |  |
| 5 |  |  |  |  |  |  |  |  | 1 |  |
| 6 | 5 |  |  |  |  |  |  |  |  |  |
| 7 | 1 |  |  |  |  |  |  |  |  |  |
| 8 |  |  |  |  |  |  | 3 |  |  |  |
| 9 | , |  |  |  |  |  |  |  |  |  |
| 10 | 2 |  |  |  |  |  |  |  |  |  |
| 11 |  |  |  |  |  |  |  |  |  |  |
| 12 |  |  |  |  |  |  |  |  |  |  |
| 13 |  |  |  |  |  |  |  |  |  |  |
| 14 |  |  |  |  |  |  |  |  |  | 1 |
| 15 |  |  |  |  |  |  |  |  |  |  |
| 16 |  |  |  |  |  |  |  |  |  |  |
| 17 |  |  |  | 1 |  |  |  |  |  |  |
| 18 |  |  |  |  |  |  |  |  |  |  |
| 19 |  |  |  |  |  |  |  |  |  |  |
| 20 |  |  |  |  |  |  |  |  |  |  |
| 21 | 1 |  |  |  |  |  |  |  |  |  |
| 22 |  |  | 1 |  |  |  |  |  |  |  |
| 23 |  |  |  |  | 1 |  |  |  |  |  |
| 24 |  |  |  | 1 |  |  |  | 1 |  |  |
| Total Fish | 35 | 8 | 8 | 10 | 7 | 1 | 3 | 1 | 1 | 1 |
| avg. days | 4.8 | 2.1 | 4.6 | 5.8 | 5.0 | 2.0 | 8.0 | 24.0 | 5.0 | 14.0 |
| miles/day | 9.5 | 13.8 | 6.4 | 5.4 | 6.3 | 15.8 | 3.9 | 1.3 | 6.3 | 2.3 |
| flow | 1,229 | 1,229 | 1,365 | 1,229 | 1,365 | 4,129 | 4,158 | 3,535 | 2,641 | 1,146 |
| mean rec. LN | 89.2 | 85.4 | 86 | 87.3 | 85.7 | 34.2 | 35.8 | 35.2 | 35.8 | 49.6 |

## Mean Lengths at Release and Recapture

There was little change in mean lengths between all fish released and those that were recaptured (Figure 22). For both RM 40 releases, mean length was smaller at release than recapture; however, three fish from the two groups were recaptured 17-24 days after release and most likely grew during that period, contributing to the higher mean
S.P. Cramer \& Associates, Inc. lengths observed at recapture. Currently, we have too little evidence to determine whether vulnerability to predation is size dependent.


Figure 22. Mean lengths at release and recapture for 1999 test groups released to estimate survival.

## CONCLUSIONS

1. The estimated number of juvenile chinook that passed Caswell between January 18 and June 30, 1999 was $1,321,042$ with an approximate $95 \%$ confidence interval of $1,007,443$ to $1,634,642$.
2. The majority of chinook captured in 1999 were fry. Peak catch of 2,322 chinook fry occurred on February 10, 1999. Indices of daily fry passage exceeded 20,000 on 27 days in 1999. It is uncertain whether or not a significant number of fish outmigrated prior to trap installation; however, degree day analysis suggests outmigration started only 10 days prior to the onset of sampling.
3. In 1999 sampling began earlier than in 1998 (the only other year when sampling began in January), fry were already migrating when sampling began, and fry passage continued until a later date than in 1998. In contrast, parr and smolt outmigrants were less abundant in 1999 than in 1998.
4. During the sampling season, we captured 12 rainbow trout/steelhead at Caswell, ranging in size from 83 mm to 255 mm . Two distinct size classes were apparent (200-300 mm and <100mm), representing yearlings and young-of-the-year, respectively. More rainbow trout/steelhead were captured in 1999 than 1998, but 1999 counts were comparable to 1997.
5. There was no clear indication that variation in flow stimulated fry movement, but the magnitude of flow may have encouraged fry movement. Similar to 1998, the peak of chinook fry passage in 1999 corresponded with elevated flows. Heavy rains resulted in elevated flows for most of January and February, and fry migration peaked on February 12, when over 80,000 fry were estimated to pass Caswell in a single night. Fry passage estimates were high throughout late January and all
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February, while flows generally ranged between 3,000-4,000 cfs.
6. The mean lengths of juvenile chinook gradually increased over the course of sampling, ranging from about 35 mm at the start of sampling (mid-January) to about 90 mm in late June. The rate of length increase was slower than other years and the threshold size for classifying smolts ( 80 mm ) was not reached until May 10, which was 3 to 7 weeks later than other years.

## RECOMMENDATIONS

1. More releases of marked fish should be made over different environmental and biological conditions including flow, turbidity, and fish size. It is especially important that releases be conducted at turbidities exceeding 10 NTU, since there appears to be a strong relationship between trapping efficiency and turbidity above 10 NTU, but turbidity exceeding 10 NTU's was only tested in 1996.
2. Chinook spawner data should be obtained from CDFG and included in future analyses. Since multiple years of outmigration data are now available, spawner data can now be used to develop correlations between spawn timing, river temperature, emergence and migration timing, and size at outmigration. The percentage of the population that spawns below Oakdale and their relative timing may also help understand differences in chinook life stage abundance between years, and the extent to which different life stages may rear between Oakdale and Caswell.
3. To accurately estimate total outmigrant abundance the traps should be installed early enough to determine the start of migration each year. This will also allow emergence and migration timing to be correlated to spawn timing, and to environmental factors. Both the Caswell and Oakdale traps should begin sampling on the same dates, sometime in mid-December. In each year sampled at Caswell, catches of chinook were high immediately after the traps were installed.

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S.P. Cramer \& Associates, Inc. 1999 Caswell Report

September 2000
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# Appendix A. Estimated 1999 Trapping Efficiency and Fish Outmigration Index at Caswell (with updated 1996 through 1998 estimates) 

Prepared by<br>Doug Neeley<br>Statistical Consultant<br>International Statistical Training and Technical Services<br>Oregon City, Oregon

The daily screw-trap count on day i at Caswell was expanded by dividing it by the predicted daily trapping efficiency ( $\mathrm{e}_{\mathrm{i}}$, predicted proportion of fish trapped) to estimate the daily outmigration index $\left(\mathrm{o}_{\mathrm{i}}\right)$ :

$$
o_{i}=\frac{c_{i}}{e_{i}}=c_{i}\left\{\frac{1}{e_{i}}\right\}
$$

## Daily Counts ( $\mathrm{c}_{\mathrm{i}}$ )

Daily counts from two screw traps, referred to as the north and south traps, were available from February 6 through July 1, 1996; from March 19 through June 27, 1997; from January 29 to July 16, $1998^{3}$; and from January 18 through June 30, 1999 (hereafter referred to as passage days). The combined count over the two traps was the count that was expanded.

## Correction for Missing Counts

There were days when counts were not made. If no counts were made on a given day, the combined count over traps was estimated using combined counts from the previous five and subsequent five days. The estimation procedure involved the following steps:

1. Adding one to the combined counts from the five previous and five subsequent days,
2. Taking the natural logs of the resulting values,
3. Computing the weighted mean of those natural logs, and
4. Retransforming the resulting mean

The computation is summarized in the following equation:

$$
\bar{c}(i)=\exp \left\{\frac{\sum_{j=1}^{5} w(i+j)^{*} \ln [c(i+j)+1]+\sum_{j=1}^{5} w(i-j)^{*} \ln [c(i-j)+1]}{\sum_{j=1}^{5} w(i+j)+\sum_{j=1}^{5} w(i-j)}\right\}-1
$$

wherein, $\ln []$ represents natural $\log$ function, $\exp \}$ represents the exponential function, and $w()$ represents a weighting variable. The weights are greater for more proximal days, specifically,

$$
\begin{aligned}
& w(i+1)=5, w(i+2)=4, w(i+3)=3, w(i+4)=2, w(i+5)=1, \\
& w(i-1)=5, \quad w(i-2)=4, \quad w(i-3)=3, \quad w(i-4)=2, w(i-5)=1,
\end{aligned}
$$

unless the count on the day associated with the weight is also missing or is associated with a stopped trap in which case the associated weight is 0 .

## Adjusting Counts on Days when the Trap has Stopped

In previous reports, no adjustments were made for trap stoppages. Occasionally, the trap stopped prior to being checked. Under trap stoppage, an expanded unadjusted count would tend to under-estimate outmigration since only a portion of the outmigration would have passed while the trap was operating. Adjustments were made to the 1996 though 1999 counts made from stopped traps. If the adjusted count was greater than the actual count from the stopped trap, then the adjusted count was used; otherwise, the actual count was used. The adjustments depended on whether only one trap or both traps were stopped. If both traps were stopped, the combined trap count was treated as a missing value and was estimated in the same manner given in the previous section, Correction for Missing Counts. If only one of the two traps had stopped, then the count from the stopped trap was adjusted using the north-to-south trap count ratio, r[(north-trap count)/(south-trap count)].

Specifically, if the north trap but not the south trap were stopped on day i, the north-trap count was estimated by
count $(i$, north trap $)=$ count(i, south trap)* r[(north-trap count)/(south-trap count)],
and, if the south trap but not the north trap were stopped on day $i$, the south-trap count was estimated by
count (i, south trap) $=$ count $(i$, north trap) $/ r[($ north-trap count $) /($ south-trap count $)]$
The north-to-south trap count ratio was computed from the total counts from the two traps over days when neither trap experienced stoppages and neither trap experienced missing values.

The degree of difference between north trap and south trap counts varied over years and over life stages within year; therefore different ratios were estimated for each life stage within each year. The 1999 north-trap and south-trap count and mean fish size are given in Tables A.1.a, A.1.b, and A.1.c respectively for fry-, parr-, and smolt-cohort segments of the run. The counts given in the table are for only those days on which there were no missing values or trap stoppages. The life-stage cohorts are defined by fish length:

$$
\begin{array}{ll}
\text { fry: } & \text { length }<=45 \mathrm{~mm} \\
\text { fingerling: } & 45 \mathrm{~mm}<\text { length }<=80 \mathrm{~mm} \\
\text { smolt: } & \text { size }>80 \mathrm{~mm} .
\end{array}
$$

The partitioning of the outmigration period into three life-stage segments of contiguous days was complicated by size fluctuation of the sampled fish over the outmigration. Although fish size showed a strong tendency to increase with time, fluctuations did result in the lengths of sampled fish on one day sometimes being larger than those on a subsequent day. Therefore, the following algorithm was followed to identify a point to separate fry from parr and a point to separate parr from smolt: When the number of continuous days (run of days) that fish fell into the larger size category permanently exceeds the previous number of continuous days when the fish fell into the smaller size category, the point between these two runs of days was used to separate the smaller and larger size cohorts. This was the method used to establish the dates of cohort outmigration given in Tables A.1.a through A.1.c and elsewhere in this appendix. There was no such algorithm applied in the earlier reports; therefore the cohort outmigration periods presented in this report sometimes differ slightly from those presented in the 1998 report.

Referring to Tables A.1.a through A.1.c, the daily fry number sampled by the north trap was significantly greater than that by the south trap, but the number of parr and smolt sampled by the north trap was significantly less ( $\mathrm{P}<0.0001$, Tables A.1.a. through A.1.c). A larger number of fry in the north trap relative to the south trap was not observed in previous years. In 1997 and 1998, the north-trap counts were less than the south-trap counts throughout the run ${ }^{4}$. There were also fish-size differences between the north- and south-trap that were inconsistent over cohorts and years. The length of fry sampled from the north trap in 1999 was significantly smaller than from the south trap ( $\mathrm{P}<0.0001$, Table A.1.a), but the lengths of sampled north- and south-trapped parr and smolt did not significantly differ ( $\mathrm{P}>0.5$ from Tables A.1.b and A.1.c). In 1998, early-run north-trapped fish were significantly smaller than south-trapped fish but late-run north-trapped fish were significantly larger; whereas, 1997 north-trapped fish tended be larger throughout the run. The variation between north and south count differences are not clearly associated with the changing size differences. The extent to which changing count differences are associated with changing fish morphology/physiology or changing river conditions is unknown.

Refer to previous reports: Outmigration Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site 1997 (June, 1998) and Outmigration Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site (June, 1999).

Because of the changing differences between the north-trapped and south-trapped fish, the north-to-south-trap count ratio was computed separately for each age-cohort run segment for each year. These ratios are presented in Table A.2. The application of the 1999 fry ratio to the stopped trap counts may result in biases for part of the fry outmigration. The north trap counts exceed the south trap counts for all recoveries up to March 2 but are less than the south counts beginning on March 9 (Table A.1.a). The intervening days involve trap stoppages and the application of the ratio to the counts could either overestimate or underestimate the counts, depending on whether the catch was more similar to the majority of the fry outmigration (up to March 2) or to the later part of the outmigration (March 9 through March 15). The adjusted values given in Appendix A.4.d.

Table A.1.a. Comparisons between Caswell north- and south-trap spring Chinook fry counts and sizes in 1999

| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | $\begin{gathered} \text { Difference } \\ \text { in means } \\ \{\mathrm{m}(\mathrm{~N})-\mathrm{m}(\mathrm{~S})\} \end{gathered}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North Trap c(N) | South Trap c(S) | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $\mathbf{c}(\mathbf{N})-\mathrm{c}(\mathrm{~S})$ | $\begin{aligned} & \{\ln [c(N)+1]- \\ & \ln \{c(S)+1]\} \end{aligned}$ | $\begin{gathered} \hline \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \\ \hline \end{gathered}$ | $\begin{gathered} \text { (sample) } \\ \text { size) } \end{gathered}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { (sample) } \\ \text { size) } \\ \hline \end{gathered}$ |  |  |
| 01/18/99 | 13 | 0 | 13 | 2.639 | 34.62 | 13 |  |  |  |  |
| 01/19/99 | 10 | 6 | 4 | 0.452 | 35.00 | 10 | 36.00 | 6 | -1.00 | 7.5 |
| 01/20/99 | 111 | 1 | 110 | 4.025 | 34.08 | 50 | 37.00 | 1 | -2.92 | 2.0 |
| 01/21/99 |  |  |  |  |  |  |  |  |  |  |
| 01/22/99 | 1,738 | 111 | 1,627 | 2.743 | 34.38 | 50 | 35.38 | 50 | -1.00 | 50.0 |
| 01/23/99 | 793 | 19 | 774 | 3.681 | 35.06 | 50 | 35.21 | 19 | -0.15 | 27.5 |
| 01/24/99 | 185 | 0 | 185 | 5.226 | 34.64 | 50 |  |  |  |  |
| 01/25/99 | 938 | 0 | 938 | 6.845 | 34.48 | 50 |  |  |  |  |
| 01/26/99 | 688 | 78 | 610 | 2.166 | 35.14 | 50 | 35.52 | 50 | -0.38 | 50.0 |
| 01/27/99 | 590 | 156 | 434 | 1.326 | 33.96 | 50 | 35.82 | 50 | -1.86 | 50.0 |
| 01/28/99 | 909 | 0 | 909 | 6.813 | 34.24 | 50 |  |  |  |  |
| 01/29/99 | 578 | 45 | 533 | 2.533 | 34.44 | 50 | 35.53 | 30 | -1.09 | 37.5 |
| 01/30/99 | 486 | 105 | 381 | 1.525 | 34.48 | 50 | 34.68 | 50 | -0.20 | 50.0 |
| 01/31/99 | 395 | 226 | 169 | 0.556 | 34.32 | 50 | 34.65 | 51 | -0.33 | 50.5 |
| 02/01/99 | 187 | 123 | 64 | 0.416 | 35.04 | 50 | 36.26 | 50 | -1.22 | 50.0 |
| 02/02/99 | 569 | 356 | 213 | 0.468 | 35.26 | 50 | 34.68 | 50 | 0.58 | 50.0 |
| 02/03/99 | 330 | 203 | 127 | 0.484 | 35.12 | 69 | 35.39 | 64 | -0.27 | 66.4 |
| 02/04/99 | 392 | 190 | 202 | 0.722 | 34.80 | 70 | 35.18 | 60 | -0.38 | 64.6 |
| 02/05/99 | 358 | 228 | 130 | 0.450 | 35.00 | 59 | 35.84 | 63 | -0.84 | 60.9 |
| 02/06/99 | 798 | 312 | 486 | 0.937 | 34.24 | 50 | 33.80 | 50 | 0.44 | 50.0 |
| 02/07/99 | 535 | 188 | 347 | 1.042 | 35.46 | 50 | 35.14 | 50 | 0.32 | 50.0 |
| 02/08/99 | 379 | 32 | 347 | 2.444 | 34.48 | 50 | 35.59 | 32 | -1.11 | 39.0 |
| 02/09/99 | 1,382 | 8 | 1,374 | 5.035 | 34.72 | 50 | 38.00 | 8 | -3.28 | 13.8 |
| 02/10/99 | 1,921 | 401 | 1,520 | 1.565 | 34.53 | 70 | 35.99 | 70 | -1.46 | 70.0 |
| 02/11/99 | 1,903 | 0 | 1,903 | 7.552 | 34.36 | 50 |  |  |  |  |
| 02/12/99 | 1,326 | 906 | 420 | 0.381 | 35.10 | 50 | 35.44 | 50 | -0.34 | 50.0 |
| 02/13/99 | 1,261 | 175 | 1,086 | 1.970 | 34.76 | 50 | 36.38 | 50 | -1.62 | 50.0 |
| 02/14/99 | 855 | 288 | 567 | 1.086 | 35.55 | 51 | 35.66 | 50 | -0.11 | 50.5 |
| 02/15/99 | 1,520 | 2 | 1,518 | 6.229 | 34.90 | 50 | 37.50 | 2 | -2.60 | 3.8 |
| 02/16/99 | 42 | 114 | (72) | -0.984 | 36.17 | 42 | 35.53 | 51 | 0.64 | 46.1 |
| 02/17/99 | 613 | 130 | 483 | 1.545 | 35.34 | 50 | 35.60 | 50 | -0.26 | 50.0 |
| 02/18/99 | 527 | 40 | 487 | 2.556 | 35.06 | 50 | 36.68 | 40 | -1.61 | 44.4 |
| 02/19/99 | 978 | 0 | 978 | 6.887 | 34.30 | 50 |  |  |  |  |
| 02/20/99 | 761 | 136 | 625 | 1.716 | 35.98 | 51 | 36.92 | 50 | -0.94 | 50.5 |
| 02/21/99 | 740 | 287 | 453 | 0.945 | 34.90 | 50 | 35.71 | 51 | -0.81 | 50.5 |
| 02/22/99 | 799 | 335 | 464 | 0.868 | 34.92 | 50 | 35.69 | 52 | -0.77 | 51.0 |
| 02/23/99 | 767 | 35 | 732 | 3.060 | 35.48 | 50 | 36.06 | 35 | -0.58 | 41.2 |

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| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | Difference in means$\{\mathrm{m}(\mathrm{~N})-\mathrm{m}(\mathrm{~S})\}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North Trap c(N) | South Trap c(S) | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $\mathrm{c}(\mathrm{~N})-\mathrm{c}(\mathrm{~S})$ | $\begin{aligned} & \{\ln [c(N)+1]- \\ & \ln \{c(S)+1]\} \\ & \hline \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \text { (sample) } \\ \text { size) } \end{gathered}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \text { (sample) } \\ \text { size) } \end{gathered}$ |  |  |
| 02/24/99 | 503 | 277 | 226 | 0.595 | 35.26 | 50 | 35.30 | 50 | -0.04 | 50.0 |
| 02/25/99 | 353 | 138 | 215 | 0.935 | 33.56 | 50 | 34.66 | 50 | -1.10 | 50.0 |
| 02/26/99 | 242 | 154 | 88 | 0.450 | 35.48 | 50 | 36.16 | 51 | -0.68 | 50.5 |
| 02/27/99 | 268 | 86 | 182 | 1.129 | 34.34 | 50 | 35.42 | 50 | -1.08 | 50.0 |
| 02/28/99 | 213 | 116 | 97 | 0.604 | 35.20 | 76 | 35.35 | 68 | -0.16 | 71.8 |
| 03/01/99 | 162 | 75 | 87 | 0.763 | 35.12 | 50 | 35.54 | 50 | -0.42 | 50.0 |
| 03/02/99 | 208 | 106 | 102 | 0.670 | 37.06 | 51 | 37.10 | 51 | -0.04 | 51.0 |
| 03/03/99 |  |  |  |  |  |  |  |  |  |  |
| 03/04/99 |  |  |  |  |  |  |  |  |  |  |
| 03/05/99 |  |  |  |  |  |  |  |  |  |  |
| 03/06/99 |  |  |  |  |  |  |  |  |  |  |
| 03/07/99 |  |  |  |  |  |  |  |  |  |  |
| 03/08/99 |  |  |  |  |  |  |  |  |  |  |
| 03/09/99 | 43 | 50 | (7) | -0.148 | 40.56 | 43 | 42.32 | 50 | -1.76 | 46.2 |
| 03/10/99 | 31 | 78 | (47) | -0.904 | 39.71 | 31 | 41.67 | 63 | -1.96 | 41.6 |
| 03/11/99 | 19 | 20 | (1) | -0.049 | 37.11 | 19 | 37.65 | 20 | -0.54 | 19.5 |
| 03/12/99 |  |  |  |  |  |  |  |  |  |  |
| 03/13/99 |  |  |  |  |  |  |  |  |  |  |
| 03/14/99 |  |  |  |  |  |  |  |  |  |  |
| 03/15/99 | 14 | 24 | (10) | -0.511 | 44.79 | 14 | 40.83 | 24 | 3.95 | 17.7 |
| Mean | 605 | 135 | 470 | 1.945 |  |  |  |  | -0.618 |  |
| Standard Error (SE) |  |  |  | 0.317 |  |  |  |  | 0.132 |  |
| Degrees of Freedom |  |  |  | 46 |  |  |  |  | 40 |  |
| t-ratio [(Mean-0)/SE] |  |  |  | 6.14 |  |  |  |  | -4.67 |  |
| P (Type I Error) |  |  |  | 0.0000 |  |  |  |  | 0.0000 |  |

Table A.1.b. Comparisons between Caswell north- and south-trap spring Chinook parr counts and sizes in 1999

| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | $\begin{gathered} \text { Difference } \\ \text { in means } \\ \{\mathrm{m}(\mathrm{~N})-\mathrm{m}(\mathrm{~S})\} \\ \hline \end{gathered}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North <br> Trap <br> c(N) | South <br> Trap $c(S)$ | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $\mathbf{c}(\mathrm{N})-\mathrm{c}(\mathrm{~S})$ | $\begin{aligned} & \{\ln [\mathrm{c}(\mathrm{~N})+1]- \\ & \ln \{\mathrm{c}(\mathrm{~S})+1]\} \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathbf{m}(\mathrm{N})\} \end{gathered}$ | $\begin{aligned} & \text { (sample) } \\ & \text { size) } \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{aligned} & \text { (sample) } \\ & \text { size) } \end{aligned}$ |  |  |
| 03/20/99 | 15 | 24 | -9 | -0.446 | 48.31 | 13 | 45.04 | 24 | 3.27 | 16.9 |
| 03/21/99 |  |  |  |  |  |  |  |  |  |  |
| 03/22/99 | 9 | 12 | -3 | -0.262 | 43.78 | 9 | 47.42 | 12 | -3.64 | 10.3 |
| 03/23/99 | 8 | 10 | -2 | -0.201 | 49.75 | 8 | 53.20 | 10 | -3.45 | 8.9 |
| 03/24/99 |  |  |  |  |  |  |  |  |  |  |
| 03/25/99 | 2 | 12 | -10 | -1.466 | 50.50 | 2 | 52.75 | 12 | -2.25 | 3.4 |
| 03/26/99 | 19 | 20 | -1 | -0.049 | 52.16 | 19 | 51.50 | 20 | 0.66 | 19.5 |
| 03/27/99 | 20 | 41 | -21 | -0.693 | 52.50 | 20 | 54.10 | 41 | -1.60 | 26.9 |
| 03/28/99 | 22 | 35 | -13 | -0.448 | 55.73 | 22 | 55.71 | 35 | 0.02 | 27.0 |
| 03/29/99 | 23 | 30 | -7 | -0.256 | 50.78 | 23 | 54.69 | 30 | -3.91 | 26.0 |
| 03/30/99 | 5 | 15 | -10 | -0.981 | 56.20 | 5 | 58.33 | 15 | -2.13 | 7.5 |
| 03/31/99 | 4 | 10 | -6 | -0.788 | 58.75 | 4 | 57.20 | 10 | 1.55 | 5.7 |
| 04/01/99 | 14 | 49 | -35 | -1.204 | 60.82 | 11 | 64.16 | 45 | -3.34 | 17.7 |
| 04/02/99 | 13 | 64 | -51 | -1.535 | 59.46 | 13 | 62.28 | 50 | -2.82 | 20.6 |
| 04/03/99 | 9 | 38 | -29 | -1.361 | 69.33 | 9 | 62.92 | 38 | 6.41 | 14.6 |
| 04/04/99 |  |  |  |  |  |  |  |  |  |  |
| 04/05/99 | 10 | 16 | -6 | -0.435 | 63.75 | 8 | 61.94 | 16 | 1.81 | 10.7 |
| 04/06/99 | 10 | 25 | -15 | -0.860 | 62.50 | 10 | 65.96 | 25 | -3.46 | 14.3 |

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| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | Difference in means$\{\mathrm{m}(\mathrm{~N})-\mathrm{m}(\mathrm{~S})\}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North <br> Trap <br> c(N) | South <br> Trap $\mathrm{c}(\mathrm{~S})$ | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $\mathbf{c}(\mathrm{N})-\mathrm{c}(\mathrm{~S})$ | $\begin{aligned} & \{\ln [c(\mathrm{~N})+1]- \\ & \ln \{\mathrm{c}(\mathrm{~S})+1]\} \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathbf{m}(\mathrm{N})\} \end{gathered}$ | $\begin{aligned} & \text { (sample) } \\ & \text { size) } \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{aligned} & \text { (sample) } \\ & \text { size) } \end{aligned}$ |  |  |
| 04/07/99 | 8 | 15 | -7 | -0.575 | 58.50 | 8 | 61.13 | 15 | -2.63 | 10.4 |
| 04/08/99 | 6 | 5 | 1 | 0.154 | 67.50 | 6 | 65.20 | 5 | 2.30 | 5.5 |
| 04/09/99 2.30 ${ }^{\text {2 }}$ |  |  |  |  |  |  |  |  |  |  |
| 04/10/99 | 10 | 11 | -1 | -0.087 | 59.90 | 10 | 69.91 | 11 | -10.01 | 10.5 |
| 04/11/99 | 3 | 4 | -1 | -0.223 | 62.67 | 3 | 69.25 | 4 | -6.58 | 3.4 |
| 04/12/99 | 10 | 23 | -13 | -0.780 | 70.70 | 10 | 74.83 | 23 | -4.13 | 13.9 |
| 04/13/99 | 8 | 10 | -2 | -0.201 | 74.50 | 8 | 74.80 | 10 | -0.30 | 8.9 |
| 04/14/99 | 20 | 18 | 2 | 0.100 | 72.40 | 20 | 75.44 | 18 | -3.04 | 18.9 |
| 04/15/99 | 23 | 23 | 0 | 0.000 | 71.57 | 23 | 67.32 | 23 | 4.25 | 23.0 |
| 04/16/99 |  |  |  |  |  |  |  |  |  |  |
| 04/17/99 | 12 | 45 | -33 | -1.264 | 75.75 | 12 | 74.02 | 45 | 1.73 | 18.9 |
| 04/18/99 | 21 | 31 | -10 | -0.375 | 77.14 | 21 | 74.52 | 31 | 2.63 | 25.0 |
| 04/19/99 | 16 | 43 | -27 | -0.951 | 76.19 | 16 | 73.76 | 43 | 2.43 | 23.3 |
| 04/20/99 | 27 | 11 | 16 | 0.847 | 76.22 | 27 | 69.91 | 11 | 6.31 | 15.6 |
| 04/21/99 | 19 | 8 | 11 | 0.799 | 78.58 | 19 | 73.75 | 8 | 4.83 | 11.3 |
| 04/22/99 | 15 | 13 | 2 | 0.134 | 74.53 | 15 | 71.54 | 13 | 2.99 | 13.9 |
| 04/23/99 |  |  |  |  |  |  |  |  |  |  |
| 04/24/99 | 27 | 22 | 5 | 0.197 | 81.30 | 27 | 78.45 | 22 | 2.84 | 24.2 |
| 04/25/99 | 15 | 25 | -10 | -0.486 | 82.33 | 15 | 75.76 | 25 | 6.57 | 18.8 |
| 04/26/99 | 10 | 35 | -25 | -1.186 | 77.40 | 10 | 77.80 | 35 | -0.40 | 15.6 |
| 04/27/99 |  |  |  |  |  |  |  |  |  |  |
| 04/28/99 | 20 | 34 | -14 | -0.511 | 76.85 | 20 | 76.65 | 34 | 0.20 | 25.2 |
|  |  |  |  |  |  |  |  |  |  |  |
| 04/30/99 |  |  |  |  |  |  |  |  |  |  |
| 05/01/99 |  |  |  |  |  |  |  |  |  |  |
| 05/02/99 | 24 | 68 | -44 | -1.015 | 82.63 | 24 | 81.40 | 50 | 1.22 | 32.4 |
| 05/03/99 | 9 | 25 | -16 | -0.956 | 80.44 | 9 | 76.92 | 25 | 3.52 | 13.2 |
| 05/04/99 |  |  |  |  |  |  |  |  |  |  |
| 05/05/99 |  |  |  |  |  |  |  |  |  |  |
| 05/06/99 | 45 | 74 | -29 | -0.489 | 77.09 | 45 | 79.38 | 50 | -2.29 | 47.4 |
| 05/07/99 | 16 | 39 | -23 | -0.856 | 78.44 | 16 | 78.56 | 39 | -0.13 | 22.7 |
| 05/08/99 | 24 | 40 | -16 | -0.495 | 80.13 | 24 | 78.83 | 40 | 1.30 | 30.0 |
| 05/09/99 | 23 | 45 | -22 | -0.651 | 80.43 | 23 | 79.44 | 45 | 0.99 | 30.4 |
| Mean | 15.2 | 27.4 | -12.2 | -0.509 |  |  |  |  | 0.302 |  |
| Standard Error (SE) |  |  |  | 0.0895 |  |  |  |  | 0.520 |  |
| Degrees of Freedom |  |  |  | 39 |  |  |  |  | 38 |  |
| t-ratio [(Mean-0)/SE] |  |  |  | -5.686 |  |  |  |  | 0.58 |  |
| P (Type I Error) |  |  |  | 0.0000 |  |  |  |  | 0.5645 |  |

Table A.1.c. Comparisons between Caswell north- and south-trap spring Chinook smolt counts and sizes in 1998

| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | Difference in means \{m(N)-m(S) $\}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North <br> Trap <br> c(N) | South <br> Trap $\mathbf{c}(\mathrm{S})$ | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $\mathbf{c}(\mathrm{N})-\mathbf{c}(\mathrm{S})$ | $\begin{aligned} & \{\ln [\mathrm{c}(\mathrm{~N})+1]- \\ & \ln \{\mathrm{c}(\mathrm{~S})+1]\} \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \text { (sample) } \\ \text { size) } \end{gathered}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \text { (sample) } \\ \text { size) } \end{gathered}$ |  |  |
| 05/10/99 | 16 | 39 | -23 | -0.856 | 78.00 | 16 | 80.85 | 39 | -2.85 | 22.7 |
| 05/11/99 |  |  |  |  |  |  |  |  |  |  |
| 05/12/99 | 15 | 47 | -32 | -1.099 | 81.20 | 15 | 81.57 | 47 | -0.37 | 22.7 |
| 05/13/99 | 10 | 37 | -27 | -1.240 | 77.50 | 10 | 79.19 | 37 | -1.69 | 15.7 |
| 05/14/99 | 24 | 48 | -24 | -0.673 | 82.92 | 24 | 80.83 | 48 | 2.08 | 32.0 |

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| Date | Number Caught |  |  |  | Mean Lengths of Fish |  |  |  | $\begin{gathered} \text { Difference } \\ \text { in means } \\ \{\mathrm{m}(\mathrm{~N})-\mathrm{m}(\mathrm{~S})\} \end{gathered}$ | Weight for mean Difference1 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | North <br> Trap <br> c(N) | South <br> Trap $\mathbf{c}(\mathrm{S})$ | Difference |  | North Trap |  | South Trap |  |  |  |
|  |  |  | in Counts $c(N)-c(S)$ | $\begin{aligned} & \{\ln [c(N)+1]- \\ & \ln \{c(S)+1]\} \end{aligned}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \hline \text { (sample) } \\ \text { size) } \\ \hline \end{gathered}$ | $\begin{gathered} \text { Mean } \\ \{\mathrm{m}(\mathrm{~N})\} \end{gathered}$ | $\begin{gathered} \hline \text { (sample) } \\ \text { size) } \\ \hline \end{gathered}$ |  |  |
| 05/15/99 | 25 | 15 | 10 | 0.486 | 81.48 | 25 | 80.00 | 15 | 1.48 | 18.8 |
| 05/16/99 | 19 | 31 | -12 | -0.470 | 81.11 | 19 | 80.52 | 31 | 0.59 | 23.6 |
| 05/17/99 | 15 | 34 | -19 | -0.783 | 82.87 | 15 | 81.29 | 34 | 1.57 | 20.8 |
| 05/18/99 | 34 | 61 | -27 | -0.572 | 81.85 | 34 | 83.76 | 50 | -1.91 | 40.5 |
| 05/19/99 | 18 | 20 | -2 | -0.100 | 82.56 | 18 | 83.70 | 20 | -1.14 | 18.9 |
| 05/20/99 | 11 | 29 | -18 | -0.916 | 85.27 | 11 | 81.90 | 29 | 3.38 | 16.0 |
| 05/21/99 | 42 | 84 | -42 | -0.681 | 81.62 | 39 | 83.62 | 50 | -2.00 | 43.8 |
| 05/22/99 |  |  |  |  |  |  |  |  |  |  |
| 05/23/99 | 15 | 60 | -45 | -1.338 | 81.80 | 15 | 85.36 | 50 | -3.56 | 23.1 |
| 05/24/99 | 11 | 54 | -43 | -1.522 | 83.55 | 11 | 82.70 | 50 | 0.85 | 18.0 |
| 05/25/99 |  |  |  |  |  |  |  |  |  |  |
| 05/26/99 | 18 | 100 | -82 | -1.671 | 84.67 | 18 | 85.14 | 50 | -0.47 | 26.5 |
| 05/27/99 | 19 | 54 | -35 | -1.012 | 83.58 | 19 | 85.44 | 50 | -1.86 | 27.5 |
| 05/28/99 | 5 | 22 | -17 | -1.344 | 87.40 | 5 | 86.05 | 22 | 1.35 | 8.1 |
| 05/29/99 |  |  |  |  |  |  |  |  |  |  |
| 05/30/99 |  |  |  |  |  |  |  |  |  |  |
| 05/31/99 |  |  |  |  |  |  |  |  |  |  |
| 06/01/99 | 32 | 49 | -17 | -0.416 | 85.81 | 32 | 83.29 | 49 | 2.53 | 38.7 |
| 06/02/99 | 14 | 13 | 1 | 0.069 | 88.21 | 14 | 85.85 | 13 | 2.37 | 13.5 |
| 06/03/99 | 22 | 37 | -15 | -0.502 | 87.32 | 22 | 86.24 | 37 | 1.07 | 27.6 |
| 06/04/99 | 18 | 37 | -19 | -0.693 | 83.83 | 18 | 84.57 | 37 | -0.73 | 24.2 |
| 06/05/99 | 5 | 18 | -13 | -1.153 | 87.20 | 5 | 84.33 | 18 | 2.87 | 7.8 |
| 06/06/99 |  |  |  |  |  |  |  |  |  |  |
| 06/07/99 | 5 | 26 | -21 | -1.504 | 83.40 | 5 | 88.04 | 26 | -4.64 | 8.4 |
| 06/08/99 | 7 | 23 | -16 | -1.099 | 86.29 | 7 | 85.17 | 23 | 1.11 | 10.7 |
| 06/09/99 | 17 | 17 | 0 | 0.000 | 89.06 | 17 | 88.82 | 17 | 0.24 | 17.0 |
| 06/10/99 | 6 | 12 | -6 | -0.619 | 91.00 | 6 | 88.25 | 12 | 2.75 | 8.0 |
| 06/11/99 | 4 | 25 | -21 | -1.649 | 89.75 | 4 | 87.88 | 25 | 1.87 | 6.9 |
| 06/12/99 |  |  |  |  |  |  |  |  |  |  |
| 06/13/99 |  |  |  |  |  |  |  |  |  |  |
| 06/14/99 |  |  |  |  |  |  |  |  |  |  |
| 06/15/99 |  |  |  |  |  |  |  |  |  |  |
| 06/16/99 | 3 | 9 | -6 | -0.916 | 97.33 | 3 | 90.33 | 9 | 7.00 | 4.5 |
| 06/17/99 | 8 | 7 | 1 | 0.118 | 93.57 | 8 | 90.14 | 7 | 3.43 | 7.5 |
| 06/18/99 | 3 | 4 | -1 | -0.223 | 84.00 | 3 | 89.00 | 4 | -5.00 | 3.4 |
| 06/19/99 |  |  |  |  |  |  |  |  |  |  |
| 06/20/99 |  |  |  |  |  |  |  |  |  |  |
| 06/21/99 | 0 | 2 | -2 | -1.099 | 0.00 | 0 | 94.00 | 2 | -94.00 | 0.0 |
| 06/22/99 |  |  |  |  |  |  |  |  |  |  |
| 06/23/99 | 2 | 4 | -2 | -0.511 | 93.50 | 2 | 92.75 | 4 | 0.75 | 2.7 |
| 06/24/99 | 2 | 4 | -2 | -0.511 | 89.50 | 2 | 89.67 | 4 | -0.17 | 2.7 |
| 06/25/99 | 2 | 3 | -1 | -0.288 | 98.00 | 2 | 90.33 | 3 | 7.67 | 2.4 |
| 06/26/99 |  |  |  |  |  |  |  |  |  |  |
| 06/27/99 |  |  |  |  |  |  |  |  |  |  |
| 06/28/99 |  |  |  |  |  |  |  |  |  |  |
| 06/29/99 | 0 | 1 | -1 | -0.693 | 0.00 | 0 | 103.00 | 1 | -103.00 | 0.0 |
| 06/30/99 | 2 | 2 | 0 | 0.000 | 88.50 | 2 | 97.00 | 2 | -8.50 | 2.0 |
| Mean | 12.8 | 29.4 | -16.5 | -0.728 |  |  |  |  | 0.0136 |  |
| Standard Error (SE) |  |  |  | 0.0909 |  |  |  |  | 0.394 |  |
| Degrees of Freedom |  |  |  | 33 |  |  |  |  | 32 |  |
| t-ratio [(Mean-0)/SE] |  |  |  | -8.01 |  |  |  |  | 0.03 |  |
| P (Type I Error) |  |  |  | 0.0000 |  |  |  |  | 0.9727 |  |

Table A.2. Caswell north-trap to south-trap count ratios for 1996-1999 fry, parr, and smolt.

| Year of <br> Outmigration | Life <br> Stage | Used in this Report <br> Beginning |  | Counts |  | North/South |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1996 | Ery | $02 / 06 / 96$ | $03 / 20 / 96$ | 48 | 385 | 0.1247 |
|  | Parr | $03 / 21 / 96$ | $03 / 31 / 96$ | 7 | 19 | 0.3684 |
|  | Smolt | $04 / 01 / 96$ | $07 / 01 / 96$ | 478 | 767 | 0.6232 |
| 1997 | Fry |  |  |  |  |  |
|  | Parr | $03 / 19 / 97$ | $04 / 05 / 97$ | 48 | 456 | 0.1053 |
|  | Smolt | $04 / 06 / 97$ | $06 / 27 / 97$ | 222 | 999 | 0.2222 |
| 1998 | Fry | $01 / 29 / 98$ | $03 / 07 / 98$ | 1,201 | 6,922 | 0.1735 |
|  | Parr | $03 / 08 / 98$ | $04 / 23 / 98$ | 594 | 3,748 | 0.1585 |
|  | Smolt | $04 / 24 / 98$ | $07 / 16 / 98$ | 406 | 1,664 | 0.2440 |
| 1999 | Fry | $01 / 18 / 99$ | $03 / 15 / 99$ | 28,433 | 6,360 | 4.4706 |
|  | Parr | $03 / 16 / 99$ | $05 / 09 / 99$ | 594 | 1,068 | 0.5562 |
|  | Smolt | $05 / 10 / 99$ | $06 / 30 / 99$ | 449 | 1,028 | 0.4368 |

## Daily Efficiency ( $\mathbf{e}_{\mathrm{i}}$ )

On 25 days during the 1996 though 1999 outmigration periods, a total of 43 uniquely marked night-time releases ${ }^{5}$ were made at a fixed distance upriver from Caswell screw traps for the purpose of estimating trapping efficiency. Estimated efficiencies were simply the proportions of the released fish that were later trapped. In order to predict the efficiency for each passage day, the efficiency estimates had to be related as a response or "dependent" variable to predictor or "independent" variables that were measured every day that the screw traps were operating. Substituting a given day's values of the predictor variables into the predictive relation would then provide an estimate of that day's efficiency.

The predictor variables explored were flow (f in cubic feet per second, cfs) measured at Orange Blossom Bridge (OBB), size of recovered fish (s as length in millimeters, mm), and turbidity ( t in nephelometric turbidity units, ntu) when turbidity reached 10 . Efficiency (e), the proportion

5 The number of standard efficiency night-time releases:
In 1996, 1 on Feb 14, 1 on Feb 19, 1 on Mar 22, 4 on Apr 6, 2 on May 2, 2 on May 10, 2 on May 26, and 2 on Jun 10;
In 1997, 1 over a period from Apr 7 through 11 (denoted as Apr 9, mid-point day) and 4 releases on the night of May 28/29 (designated as a May $28^{\text {th }}$ release) ;
In 1998, 3 on Mar 14, 3 on Mar 25, 2 on Apr 18, 2 on May 10, 2 on May 18, 2 on Jun 4, and 2 on June 2;
In 1999, 1 each on Feb 20 and 27, Mar 2, 9 and 17, Jun 2 and 3; and 2 on Jun 4
Day-time releases were omitted for reasons given in the 1998 report (June 1999)
of released fish trapped per release, was related to the predictor variables using the following logistic function:

$$
e_{i}=\frac{1}{1+\exp \left[-b(0)-b(f) * f-b(s) * s-b\left(t^{\prime}\right) * t^{\prime}\right]}
$$

or, in the form of the "logit" linear transform,

$$
\log i t\left(e_{i}\right)=\ln \left[\frac{e_{i}}{1-e_{i}}\right]=b(0)+b(f)^{*} f+b(s)^{*} s+b\left(t^{\prime}\right)^{*} t^{\prime}
$$

In the above equations, " $\mathrm{b}(0)$ " is a coefficient associated with the intercept ${ }^{6}$, and $\mathrm{b}(\mathrm{f}), \mathrm{b}(\mathrm{s})$, and $\mathrm{b}\left(\mathrm{t}^{\prime}\right)$ are partial logistic regression coefficients relating the logit transform of efficiency predictor respectively to flow, size, and turbidity when turbidity is at least 10 . A major reason for choosing the logistic model is that the predicted efficiency can never be less than 0 and can never exceed 1 ( $100 \%$ ). The logistic regression used assumes that the underlying distribution of the number of captured fish is binomial when the model is accurate. Adjustments to the standard errors, variances, and covariances of the estimated coefficients for failure of the residuals to be binomially distributed had to be made, the adjustment procedures being discussed in Appendix A.1.

## Predictor Variables

The predictor variables evaluated in this 1999 analysis were the same as in previous years; however, many of the 1996 through 1998 values differed from those used in the 1998 report:

1. Flow: In the current analysis I used the mean of release-day and recovery-day flows. In last year's analysis I used release-day flows. The reason for using the mean of release-day and recovery-day flows is that releases were made in the evening of the release day and almost all were recovered by the following morning (recovery day). Therefore the mean of the two days’ flows was considered to be a better indicator of the flow during the recovery period then was the release-day flow. In general, the predictor variable is the mean of the flows from the day of recovery and from the day prior to recovery.
2. Fish size: The fish length used in the 1998 report was not consistent over years. The
lengths of a sample of released fish was used for the 1996-1997 ${ }^{\text {data set; }}$ whereas, the lengths of a sample of recovered fish was used for the 1998 data set. In the current analysis, recovered fish length was used for all three data sets--1996-97, 1998, and 1999. The use of release size in previous reports could have contributed to a bias. Tables A.3.a though A.3.c present comparison between the size of fish sampled prior to release and the size of a sample of those same fish after recovery for all efficiency releases. While the mean release-recovery size difference is small and not significant for fry ( $\mathrm{P}=0.5385$ in Table A.3.a), the mean sizes of sampled released parr and smolt exceeded those of sampled recovered parr and smolt by more than 1 $\mathrm{mm}(\mathrm{P}=0.0178$ for parr in Table A.3.b and $\mathrm{P}=0.0637$ for smolt in Table A.3.c). Whether the sizes of released parr and smolt actually exceeded that of recovered parr and smolt in 1999 is unclear since no efficiency releases were made during the parr outmigration period in 1999 (Table A.3.b) and since the mean sizes of sampled recovered smolt were somewhat larger than those of sampled released smolt for the four 1999 efficiency releases (Table A.3.c).
3. Turbidity: Turbidity was never used in the prediction unless it was at least 10 because a threshold of 10 resulted in the greatest precision of the estimated coefficients in the model. For 1996 releases, I used release-day turbidities; whereas for 1997 and 1998 releases, I used recovery-day turbidities. Since turbidities were checked and recorded in the morning when the recovered fish are enumerated, I used recovery-day turbidities in the current analysis for all three data sets--1996-97, 1998, and 1999. The turbidity can change dramatically from one day to another. For the 1996 fit, the recovery-day turbidity had a dramatic affect on the fit when the turbidity reaches 10 , but the effect was substantially reduced if the release-day turbidity was used. It should be noted that the only efficiency releases experiencing turbidities of at least 10 were those made in 1996. However, turbidities did exceed 10 for some days (release or non-release days) during the outmigration in all years.

Linear interpolation was used to estimate missing predictor variable values from the nearest straddling days’ values. For example, if there was a flow of 1000 cfs on Day 4 (Day j) and there was a flow of 1200 cfs on Day 9 (Day j') and if there were no intervening flow measures, then the missing values for Day 5 through Day 8 (Day i) would be computed as follows:

7 The 1997 and 1996 efficiency-release data sets were combined to obtained 1996-1997 coefficient estimates because there were only two release days in 1997 and because, for the flows on those two days, the efficiency estimates seemed more comparable to those in 1996 than they did to those in 1998 and 1999 (refer to Table A.4).

Day 4: 1000 (actual value)
Missing Value for day $i=$
[Day j' - Day i)*(Day j value)+(Day i - Day j)*(Day j' value)]/(Day j' - Day i)
Day 5: [(9-5)*1000 + (5-4)*1200]/(9-4) $=[4 * 1000+1 * 1200] /(9-4)=1040$
Day 6: $[(9-6) * 1000+(6-4) * 1200] /(9-4)=[3 * 1000+2 * 1200] /(9-4)=1080$
Day 7: $[(9-7) * 1000+(7-4) * 1200] /(9-4)=[2 * 1000+3 * 1200] /(9-4)=1120$
Day 8: $[(9-8) * 1000+(8-4) * 1200] /(9-4)=[1 * 1000+4 * 1200] /(9-4)=1160$

## Day 9: 1200 (actual value)

Table A.3.a. Comparisons between released and recovered spring Chinook fry sizes from 1996-1999 efficiency tests.

| Release <br> Date | Mark Type | Fish Stock | Release Time | Released Fish |  | Recovered Fish |  | Released <br> - Recovered Lengths |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Length | Sample | Length | Sample |  |  |
|  |  |  |  | (mm) | Size | (mm) | Size | Difference | Weight1 |
| 14-Feb-96 | Brand | Natural | Night | 34.3 | 30 | 35.2 | 62 | -0.9 | 40.4 |
| 19-Feb-96 | Brand | Natural | Night | 33.8 | 30 | 35.5 | 56 | -1.7 | 39.1 |
| 22-Mar-96 | Panjet | Hatchery | Night | 42.7 | 30 | 41.8 | 15 | 0.9 | 20.0 |
| 14-Mar-98 | Brand | Natural | Night | 36.2 | 50 | 37.3 | 101 | -1.1 | 66.9 |
| 25-Mar-98 | Panjet | Hatchery | Night | 41.2 | 50 | 42.1 | 34 | -0.9 | 40.5 |
| 25-Mar-98 | Panjet | Hatchery | Night | 41.1 | 50 | 41.8 | 32 | -0.7 | 39.0 |
| 20-Feb-99 | Brand | Natural | Night | 33.2 | 50 | 34.1 | 86 | -0.9 | 63.2 |
| 27-Feb-99 | Brand | Natural | Night | 35.6 | 100 | 34.7 | 43 | 0.9 | 60.1 |
| 02-Mar-99 | Brand | Natural | Night | 34.1 | 50 | 34.7 | 29 | -0.6 | 36.7 |
| 09-Mar-99 | Brand | Natural | Night | 36.1 | 50 | 34.7 | 20 | 1.4 | 28.6 |
| 17-Mar-99 | Photonic | Natural | Night | 42.8 | 50 | 39.0 | 15 | 3.8 | 23.1 |
| Weighted ${ }^{1}$ Mean |  |  |  |  |  | -0.26 |  |  |  |
| Standard Error (SE) |  |  |  |  |  | 0.409 |  |  |  |
| Degrees of Freedom |  |  |  |  |  | 10 |  |  |  |
|  |  |  |  |  |  | -0.64 |  |  |  |
| P(Type I Error) |  |  |  |  |  | 0.5385 |  |  |  |

${ }^{1}$ Weight is the harmonic mean of the release and recovery sample sizes to account for differences in sample sizes within and among pairs

Table A.3.b. Comparisons between released and recovered spring Chinook parr sizes from 1996-1999 efficiency tests.

| Release <br> Date | Mark Type | Fish <br> Stock | Release Time | Released Fish |  | Recovered Fish |  | Released <br> - Recovered Lengths |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Length | Sample | Length | Sample |  |  |
|  |  |  |  | (mm) | Size | (mm) | Size | Difference | Weight1 |
| 06-Apr-96 | Brand | Hatchery | Night | 67.4 | 30 | 71.6 | 22 | -4.2 | 25.4 |
| 06-Apr-96 | Brand | Hatchery | Night | 70.2 | 30 | 72.9 | 8 | -2.7 | 12.6 |
| 02-May-96 | Panjet | Hatchery | Night | 75.5 | 30 | 75.9 | 30 | -0.4 | 30.0 |

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| 02-May-96 | Panjet | Hatchery | Night | 76.1 | 30 | 76.7 | 30 | -0.6 | 30.0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 10-May-96 | Panjet | Hatchery | Night | 74.2 | 30 | 73.4 | 50 | 0.8 | 37.5 |
| 10-May-96 | Panjet | Hatchery | Night | 76.1 | 30 | 72.9 | 55 | 3.2 | 38.8 |
| 26-May-96 | Panjet | Hatchery | Night | 72.7 | 30 | 68.2 | 65 | 4.5 | 41.1 |
| 26-May-96 | Panjet | Hatchery | Night | 71.7 | 30 | 69.9 | 60 | 1.8 | 40.0 |
| 28-May-97 | Panjet | Hatchery | Night | 71.9 | 30 | 71.5 | 35 | 0.4 | 32.3 |
| 28-May-97 | Panjet | Hatchery | Night | 72.5 | 30 | 71.9 | 30 | 0.6 | 30.0 |
| 28-May-97 | Panjet | Hatchery | Night | 71.3 | 30 | 71.9 | 52 | -0.6 | 38.0 |
| 29-May-97 | Panjet | Hatchery | Night | 73.3 | 30 | 72 | 66 | 1.3 | 41.3 |
| 14-Mar-98 | Panjet | Hatchery | Night | 55.2 | 50 | 54.1 | 35 | 1.1 | 41.2 |
| 14-Mar-98 | Panjet | Hatchery | Night | 55.1 | 50 | 53.6 | 45 | 1.5 | 47.4 |
| 25-Mar-98 | Panjet | Natural | Night | 52.4 | 50 | 48.1 | 43 | 4.3 | 46.2 |
| 18-Apr-98 | Panjet | Hatchery | Day | 75.3 | 50 | 70.7 | 4 | 4.6 | 7.4 |
| 18-Apr-98 | Panjet | Hatchery | Night | 75.1 | 50 | 73.7 | 26 | 1.4 | 34.2 |
| 18-Apr-98 | Panjet | Hatchery | Night | 74.6 | 50 | 70.3 | 15 | 4.3 | 23.1 |
| 18-Apr-98 | Panjet | Natural | Day | 65.6 | 50 | 66 | 12 | -0.4 | 19.4 |
| Weighted1 Mean |  |  |  |  |  |  |  | 1.28 |  |
| Standard Error (SE) |  |  |  |  |  |  |  | 0.492 |  |
| Degrees of Freedom |  |  |  |  |  |  |  | 18 |  |
| t-ratio [(Weighted Mean)/SE] |  |  |  |  |  |  |  | 2.61 |  |
| P(Type I Error) |  |  |  |  |  |  |  | 0.0178 |  |

Table A.3.c. Comparisons between released and recovered spring Chinook smolt sizes from 1996-1999 efficiency tests.

| Release <br> Date | Mark <br> Type | Fish Stock | Release Time | Released Fish |  | Recovered Fish |  | Released <br> - Recovered Lengths |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  | Length | Sample | Length | Sample |  |  |
|  |  |  |  | (mm) | Size | (mm) | Size | Difference | Weight1 |
| 10-Jun-96 | Panjet | Hatchery | Night | 91.6 | 30 | 85.5 | 43 | 6.1 | 35.3 |
| 10-Jun-96 | Panjet | Hatchery | Night | 90.5 | 30 | 86.8 | 56 | 3.7 | 39.1 |
| 7-11-Apr-97 | Panjet | Natural |  | 82.5 | 30 | 81.7 | 3 | 0.8 | 5.5 |
| 10-May-98 | Panjet | Hatchery | Day | 87.7 | 50 | 83 | 1 | 4.7 | 2.0 |
| 10-May-98 | Panjet | Hatchery | Night | 86.4 | 50 | 86.3 | 8 | 0.1 | 13.8 |
| 10-May-98 | Panjet | Hatchery | Night | 87.4 | 50 | 84.5 | 4 | 2.9 | 7.4 |
| 18-May-98 | Panjet | Natural | Night | 88.8 | 50 | 83.6 | 16 | 5.2 | 24.2 |
| 18-May-98 | Panjet | Hatchery | Night | 88.2 | 50 | 86.9 | 31 | 1.3 | 38.3 |
| 04-Jun-98 | Panjet | Hatchery | Night | 100.5 | 50 | 98.4 | 16 | 2.1 | 24.2 |
| 04-Jun-98 | Panjet | Hatchery | Night | 98.6 | 50 | 97.7 | 15 | 0.9 | 23.1 |
| 12-Jun-98 | Panjet | Hatchery | Night | 102.8 | 50 | 104.8 | 6 | -2.0 | 10.7 |
| 12-Jun-98 | Panjet | Hatchery | Night | 102.8 | 50 | 95.3 | 4 | 7.5 | 7.4 |
| 02-Jun-99 | Panjet | Hatchery | Night | 83.6 | 50 | 84.8 | 41 | -1.2 | 45.1 |
| 03-Jun-99 | Panjet | Hatchery | Night | 84.2 | 50 | 84.4 | 39 | -0.2 | 43.8 |
| 04-Jun-99 | Photonic | Hatchery | Night | 83.3 | 50 | 83.7 | 33 | -0.4 | 39.8 |
| 04-Jun-99 | Photonic | Hatchery | Night | 82.5 | 50 | 83.6 | 53 | -1.1 | 51.5 |
| Weighted1 Mean |  |  |  |  |  |  |  | 1.32 |  |

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|  |  |  |  | Relea | sed Fish | Recov | ered Fish | Rele | ad |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Release | Mark | Fish | Release | Length | Sample | Length | Sample | - Recovere | Lengths |
| Date | Type | Stock | Time | (mm) | Size | (mm) | Size | Difference | Weight1 |
| Standard Error (SE) |  |  |  |  |  |  |  | 0.660 |  |
| Degrees of Freedom |  |  |  |  |  |  |  | 15 |  |
| t-ratio [(mean)/SE] |  |  |  |  |  |  |  | 2.00 |  |
| P(Type I Error) |  |  |  |  |  |  |  | 0.0637 |  |
| 1 Weight is the harmonic mean of the release and recovery sample sizes to account for differences in sample sizes within and among pairs |  |  |  |  |  |  |  |  |  |

## Model Selection

The data in Table A.4. were used to develop efficiency predictors. An analysis of variation procedure was undertaken to initially evaluate the effectiveness of each predictor variable. The residual deviances produced from logistic regression are analogous the residual sums of squares from least squares regression. Therefore, the logistic regression deviances were subjected to the same analytic partitioning that sums of squares are subjected to in an analysis of variance.

Analyses of variation were initially performed separately within each set of years (19961997, 1998, 1999). As indicated in Table A.5, insufficient numbers of releases were made to permit estimation of coefficients within each size cohort within each year. For example, within each year $x$ size-cohort category having only 3 releases, up to 3 coefficients (e.g., intercept, flow, size) could be estimated, but there would be no true variation measure within these categories. Therefore, individual coefficients were not estimated within each year x size-cohort category. Instead, the following organized step-wise procedures were followed within each year and are presented in Appendix A.2.

1. Size: Size was assessed by
2. using recovery length as a one predictor variable [for estimating $b(s)$ ],
3. using separate size-cohort intercepts within each year (fry, fingerling, and smolt cohort indicator variables), and
4. using both the size and size-cohort-indicator variables.

Within each year, using just size (a. above) gave the best fit, and the further inclusion of size-cohort indicator variables did not substantially or significantly improve the fit. Sizecohort was dropped from the model. Size was dropped from the model if its contribution was not significant.
2. Flow was then added as a variable to previously retained variables and its effect assessed.
3. Recovery-day turbidity was then added to other retained variables when it exceeded 10 ntu
(which happened only in 1996). As mentioned earlier, the turbidity threshold that resulted in the best fit (the lowest deviance) was 10 ntu.

Table A.4. Predictor variables and efficiency response variable used to develop logistic efficiency predictor

| Release <br> Date | Flows |  |  |  | Recovery Day Turbidity |  | Efficiency (Proportion Recovered) | Adjusted <br> Release <br> Number |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Release Day | Recovery Day | Average | Recovery <br> Length |  |  |  |  |
|  |  |  |  |  | Value | Value >= 0 |  |  |
| 02/14/96 | 1179 | 1595 | 1387 | 35.2 | 14.7 | 14.7 | 0.1208 | 1324 |
| 02/19/96 | 2014 | 2841 | 2427.5 | 35.5 | 10.5 | 10.5 | 0.0566 | 1078 |
| 03/22/96 | 3413 | 3010 | 3211.5 | 41.8 | 7.3 | 0 | 0.0137 | 1097 |
| 04/06/96 | 1791 | 1780 | 1785.5 | 71.6 | 5.9 | 0 | 0.0295 | 746 |
| 04/06/96 | 1791 | 1780 | 1785.5 | 72.9 | 5.9 | 0 | 0.0107 | 748 |
| 05/02/96 | 1680 | 1659 | 1669.5 | 76.7 | 10.2 | 10.2 | 0.0763 | 1979 |
| 05/02/96 | 1680 | 1659 | 1669.5 | 75.9 | 10.2 | 10.2 | 0.0442 | 1990 |
| 05/10/96 | 1667 | 1653 | 1660 | 73.4 | 8.7 | 0 | 0.0223 | 2242 |
| 05/10/96 | 1667 | 1653 | 1660 | 72.9 | 8.7 | 0 | 0.0252 | 2341 |
| 05/26/96 | 921 | 955 | 938 | 69.9 | 6.8 | 0 | 0.067 | 2374 |
| 05/26/96 | 921 | 955 | 938 | 68.2 | 6.8 | 0 | 0.0544 | 2298 |
| 06/10/96 | 1279 | 1300 | 1289.5 | 85.5 | 9.0 | 0 | 0.0276 | 1559 |
| 06/10/96 | 1279 | 1300 | 1289.5 | 86.8 | 9.0 | 0 | 0.0298 | 1981 |
| 04/09/97 | 599 | 598 | 598.5 | 81.7 | 8.8 | 0 | 0.0165 | 182 |
| 05/28/97 | 1608 | 1615 | 1611.5 | 71.9 | 9.8 | 0 | 0.0273 | 1905 |
| 05/28/97 | 1608 | 1615 | 1611.5 | 71.5 | 9.8 | 0 | 0.0242 | 1444 |
| 05/28/97 | 1608 | 1615 | 1611.5 | 71.9 | 9.8 | 0 | 0.0209 | 1433 |
| 05/28/97 | 1608 | 1615 | 1611.5 | 72 | 9.8 | 0 | 0.0363 | 1817 |
| 03/14/98 | 1577 | 1574 | 1575.5 | 54.1 | 7.4 | 0 | 0.0339 | 1033 |
| 03/14/98 | 1577 | 1574 | 1575.5 | 37.3 | 7.4 | 0 | 0.047 | 2149 |
| 03/14/98 | 1577 | 1574 | 1575.5 | 53.6 | 7.4 | 0 | 0.0429 | 1049 |
| 03/25/98 | 2657 | 2351 | 2504 | 41.8 | 8 | 0 | 0.0284 | 1128 |
| 03/25/98 | 2657 | 2351 | 2504 | 48.1 | 8 | 0 | 0.049 | 877 |
| 03/25/98 | 2657 | 2351 | 2504 | 42.1 | 8 | 0 | 0.0271 | 1254 |
| 04/18/98 | 1996 | 1996 | 1996 | 70.3 | 6.1 | 0 | 0.0152 | 988 |
| 04/18/98 | 1996 | 1996 | 1996 | 73.7 | 6.1 | 0 | 0.0261 | 995 |
| 05/10/98 | 2005 | 2004 | 2004.5 | 84.5 | 6 | 0 | 0.0062 | 649 |
| 05/10/98 | 2005 | 2004 | 2004.5 | 86.3 | 6 | 0 | 0.0079 | 1009 |
| 05/18/98 | 2023 | 2016 | 2019.5 | 86.9 | 4.9 | 0 | 0.0304 | 1020 |
| 05/18/98 | 2023 | 2016 | 2019.5 | 83.6 | 4.9 | 0 | 0.0145 | 1102 |
| 06/04/98 | 1527 | 1537 | 1532 | 98.4 | 7.1 | 0 | 0.0148 | 1079 |
| 06/04/98 | 1527 | 1537 | 1532 | 97.7 | 7.1 | 0 | 0.0144 | 1044 |
| 06/12/98 | 1593 | 1564 | 1578.5 | 104.8 | 9.75 | 0 | 0.0076 | 791 |
| 06/12/98 | 1593 | 1564 | 1578.5 | 95.3 | 9.75 | 0 | 0.004 | 1000 |
| 02/20/99 | 4316 | 4291 | 4303.5 | 34.1 | 7.9 | 0 | 0.0376 | 2550 |
| 02/27/99 | 4207 | 3842 | 4024.5 | 34.7 | 3 | 0 | 0.0257 | 1672 |
| 03/02/99 | 2800 | 2861 | 2830.5 | 34.7 | 7.7 | 0 | 0.0349 | 830 |
| 03/09/99 | 1736 | 1734 | 1735 | 34.7 | 5.1 | 0 | 0.0208 | 962 |

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| Release <br> Date | Flows |  |  |  | Recovery Day Turbidity |  | Efficiency <br> (Proportion <br> Recovered) | Adjusted <br> Release <br> Number |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Release Day | Recovery Day | Average | Recovery Length |  |  |  |  |
|  |  |  |  |  | Value | Value >= 0 |  |  |
| 03/17/99 | 1577 | 1602 | 1589.5 | 39 | 4.6 | 0 | 0.0224 | 671 |
| 06/02/99 | 1365 | 1369 | 1367 | 84.8 | 6.3 | 0 | 0.0252 | 2500 |
| 06/03/99 | 1369 | 1360 | 1364.5 | 84.4 | 4.5 | 0 | 0.0157 | 2487 |
| 06/04/99 | 1360 | 1356 | 1358 | 83.6 | 1.4 | 0 | 0.0333 | 2039 |
| 06/04/99 | 1360 | 1356 | 1358 | 83.7 | 1.4 | 0 | 0.0175 | 2002 |

Table A.5. Number of releases within each year $\mathbf{x}$ cohort grouping

|  | Year Category |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Size | $\mathbf{1 9 9 6 - 1 9 9 7}$ | $\mathbf{1 9 9 8}$ | $\mathbf{1 9 9 9}$ | TOTAL |
| Fry | 3 | $3^{*}$ | 5 | 11 |
| Fingerling | 12 | 5 | 0 | 17 |
| Smolt | $3^{*}$ | 8 | 4 | 15 |
| TOTAL | 18 | 16 | 9 | 43 |
| * 2 of the 3 releases made on same day |  |  |  |  |

The within-year mean deviances ${ }^{8}$ from the models were found not to differ substantially or significantly based on paired f-tests (full-model mean deviances compared: 1996-1997 versus 1998, 1996-1997 versus 1999, and 1998 versus 1999). Therefore, a combined analysis was performed including all coefficients except the size-cohort indicators. The coefficients for this full model are presented in Appendix A.3. The model was then reduced by dropping the predictor variable associated with any given year's coefficients that did not differ significantly from 0 at the $20 \%$ significance level ( $\mathrm{P}>0.2$ ). The reason for choosing such a high significance level was to reduce the chance of omitting a coefficient when it should be included (Type 2 error). The coefficients from this refit reduced model are presented in Table A.6.a.

The above procedure is different than that used in last year's analysis wherein the full model was retained; i.e., in last year's analysis, all years $b(f), b(s)$, and $b\left(t t^{\prime}\right)$ coefficient estimates were included whether or not a coefficient differed significantly from 0 in any given year, as long as it significantly differed from 0 in at least one year. The reason for the inclusion of non-significant coefficients in last year's report was because of sign consistency: In the 1998 report, the 1996-1997 flow coefficient differed significantly from 0, and, although the 1998 flow coefficient did not differ significantly from 0, it had the same sign as the 1996-1997 coefficient. Conversely, the 1998 fishsize coefficient differed significantly from 0; whereas the 1996-1997 coefficient did not but had the same sign as the 1998 coefficient. This year, however, there were sign differences between the 1999 and the previous year's coefficients. Since non-significant 1999 coefficients had sign differences

[^0]from previous years, non-significant coefficients were dropped from the model irrespective of the sign.

It turned out that flow did not make a significant contribution to the 1998 predictor, and neither size nor flow made a significant contribution to the 1999 predictor; therefore they were dropped from the model, and the data were refit to estimate the coefficients for the reduced model. The turbidity coefficient was used for all years because the turbidity reached or exceeded 10 ntu in each year for some days during the outmigration.

Table A.6.a. Coefficient estimates and associated statistics for the efficiency model with the data sets having 1996 turbidity > 10 included

| 1996-1997 <br> Efficiency Predictor: er = 1/\{1+exp[-b(0)-b(f)*f-b(s)*s-b(t')*t']\} |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio <br> (b/SE) | $\begin{gathered} \mathbf{P} \\ \text { (Type I) } \end{gathered}$ |
| Intercept (0) | -1.423720 | 0.476060 | -2.99 | 0.0050 |
| Flow (f) | -0.000829 | 0.000167 | -4.97 | 0.0000 |
| Recovery Size (s) | -0.010380 | 0.005237 | -1.98 | 0.0551 |
| Turbidity > 10 (t) | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | $\begin{gathered} \text { Deviance } \\ 163.47 \end{gathered}$ | $\begin{gathered} \text { D.F. } \\ 36 \end{gathered}$ | $\begin{gathered} \text { Deviance/D.F. } \\ 4.54 \end{gathered}$ |  |
| Predictor | Intercept | ance-Covari <br> Flow | Recovery Size | Turbidity > 10 |
| Intercept | $2.2663 \mathrm{E}-01$ |  |  |  |
| Flow | -4.3819E-05 | 2.7835E-08 |  |  |
| Recovery Size | -2.1342E-03 | 6.5842E-08 | $2.7427 \mathrm{E}-05$ |  |
| Turbidity > 10 | -3.4783E-03 | -3.7326E-07 | $4.6771 \mathrm{E}-05$ | 2.0479E-04 |

1998
Efficiency Predictor: er $=1 /\left\{1+\exp \left[-b(0)-b(s)^{*} s-b(t) * t\right]\right\}$

| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio (b/SE) | $\begin{gathered} \mathbf{P} \\ \text { (Type I) } \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
| Intercept (0) | -2.251260 | 0.301509 | -7.47 | 0.0000 |
| Recovery Size (s) | -0.022160 | 0.004855 | -4.56 | 0.0001 |
| Turbidity > 10 (t) | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | $\begin{gathered} \text { Deviance } \\ 163.47 \\ \hline \end{gathered}$ | $\begin{gathered} \text { D.F. } \\ 36 \\ \hline \end{gathered}$ | $\begin{gathered} \text { Deviance/D.F. } \\ 4.54 \\ \hline \end{gathered}$ |  |
| Predictor | Intercept | Variance-Covariance Recovery Size | Turbidity > 10 |  |
| Intercept | $9.0907 \mathrm{E}-02$ |  |  |  |
| Recovery Size | -1.3759E-03 | $2.3567 \mathrm{E}-05$ |  |  |
| Turbidity > 10 | $0.0000 \mathrm{E}+00$ | $0.0000 \mathrm{E}+00$ | 2.0479E-04 |  |

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| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | $\begin{aligned} & \hline \text { t-ratio } \\ & \text { (b/SE) } \end{aligned}$ | $\begin{gathered} \mathbf{P} \\ \text { (Type I) } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{gathered} \text { Intercept (0) } \\ \text { Turbidity > } 10 \text { (t) } \end{gathered}$ | -3.624670 | 0.106971 | -33.88 | 0.0000 |
|  | 0.074650 | 0.014311 | 5.22 | 0.0000 |
|  | $\begin{gathered} \text { Deviance } \\ 163.47 \end{gathered}$ | $\begin{gathered} \text { D.F. } \\ 36 \end{gathered}$ | $\begin{gathered} \text { Deviance/D.F. } \\ 4.54 \end{gathered}$ |  |
| Predictor | $\begin{gathered} \text { Vari } \\ \text { Intercept } \end{gathered}$ | ovariance <br> Turbidity > 10 |  |  |
| $\begin{gathered} \text { Intercept } \\ \text { Turbidity > } 10 \end{gathered}$ | $\begin{gathered} 1.1443 \mathrm{E}-02 \\ 0.0000 \mathrm{E}+00 \end{gathered}$ | 2.0479E-04 |  |  |

The data set used did not distinguish between releases made on the same day and those made on different days; i.e., the data from each release was treated as an independent set. This creates "pseudo-replication" as far as flow and turbidity are concerned. Releases made on the same day always experience the same measured flow and turbidity; they only differ in their fish size and recovery data. Since variation is expected to be less between releases made on the same day than between releases made on different days, the concern would be whether treating multiple releases on the same day as independent releases would result in the measured precision being greater than it should be. To test this, fish sizes, release numbers, and efficiency estimates were pooled over releases within day, and these pooled data were fit using the full model. The resulting residual mean deviance was compared to the simple among-release within-day mean deviance (a measure of within-day random variation). The results of this analysis of variation is summarized at the bottom of Appendix A.2. The residual mean deviance from the fit of the pooled data was 4.52 based on 15 degrees of freedom, and the among-release within-day mean deviance was 3.85 based on 18 degrees of freedom. The first mean deviance did not significantly exceed the second mean deviance based on an F-test $(\mathrm{P}=0.3684)$. This indicates that the model was effective in explaining most, if not all, of the among day variation in efficiencies. It also suggests that the model is a reasonably good predictor. Since the two mean deviances were nearly equal, the decision was made to use fit based on treating data from each release as an independent set (i.e., use the coefficients from Table A.6.) to boost the degrees of freedom and the power of the test used.

Since the turbidity never reached 10 for the efficiency releases in 1997, 1998, and 1999, it is possible that the turbidity did not have an effect in 1997, 1998, and 1999. An additional fit was made using all years' night-time, fixed-point efficiency releases except for the 1996 efficiency releases for which the recovery day's turbidity equaled or exceed 10 ntu . The coefficients and associated statistics from this fit are present in Table A.6.b. It should be noted that t-tests were used for testing the coefficients in Tables A.6.a and A.6.b. The t-test is not truly appropriate because the estimated efficiencies are not expected to be normally distributed, it was only used because the asymptotic z-test would have been too liberal.

Table A.6.b. Coefficient estimates and associated statistics for the efficiency model with the data sets having 1996 turbidity > 10 excluded

1996-1997
Efficiency Predictor: er $=1 /\{1+\exp [-b(0)-b(f) * f-b(s) * s]\}$


1998
Efficiency Predictor: er = 1/\{1+exp[-b(0)-b(s)*s]\}

| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio <br> (b/SE) | P <br> (Type I) |
| :---: | :---: | :---: | :---: | :---: |
| Intercept (0) | -2.251260 | 0.275819 | -8.16 | 0.0000 |
| Recovery Size (s) | -0.022160 | 0.004441 | -0.20 | 0.8424 |
|  |  |  | Deviance/D.F. |  |
|  | Deviance | D.F. | 3.80 |  |

1999
Efficiency Predictor: er $=1 /\{1+\exp [-b(0)]\}$

| Predictor | Estimate <br> (b) | Standard <br> Error (SE) | t-ratio <br> (b/SE) | P <br> (Type I) |
| :---: | :---: | :---: | :---: | :---: |
| Intercept (0) | -3.624670 | 0.097857 | -37.04 | 0.0000 |
|  |  |  |  |  |
|  | Deviance | D.F. | Deviance/D.F. |  |
|  | 125.4 | 33 | 3.80 |  |


|  | Variance-Covariance |
| :---: | :---: |
| Predictor | Intercept |
| Intercept | $9.5760 \mathrm{E}-03$ |

## Estimated Outmigration ( $\mathbf{o}_{\mathbf{i}}=\mathbf{c}_{\mathbf{i}} / \mathbf{e}_{\mathbf{i}}$ )

The daily counts were expanded by dividing them by the predicted daily efficiencies (multiplying by the inverse of the efficiencies) to estimate daily outmigration:


These expansions were then accumulated over days to estimate the cumulative outmigration. The daily predictor variables used and the counts are given in Appendices A.4.a. through A.4.d. respectively for 1996 through 1999. The associated estimated daily and cumulative outmigration for those years are respectively given in Appendices A.5.a. through A.5.d. for the Table A.6.a. coefficients using the 1996 turbidity coefficient. Table A.7.a presents fry-, fingerling-, and smolt-cohort-period cumulative outmigration summaries from these appendices. Note that the 1996, 1997, and 1998 estimates are different than those presented in the 1998 report because of the data modifications. The confidence intervals presented are generally narrower than in previous reports. This is because the previous confidence interval estimate was based, in part, on the approximate variance of a ratio [ $\mathrm{S}^{2}\left(\mathrm{c}_{\mathrm{i}} / \mathrm{e}_{\mathrm{i}}\right)$ ] which turned out to be conservative (bigger than it should be). The variance estimate has improved by using instead an unbiased estimate of the variance of a product $\left[\mathrm{S}^{2}\left(\mathrm{c}_{\mathrm{i}}{ }^{*} 1 / \mathrm{e}_{\mathrm{i}}\right)\right]$. The methodology is spelled out in Appendix A.1.a.

The estimated fry outmigration in 1999 was greater than in 1998 (the previous greatest estimate). It should be borne in mind that the fry enumeration began earlier in 1999 and fry passage continued until a later date in 1999 than in 1998 (1999 fry outmigration: January 18 - March 15; 1998 fry outmigration: January 29 - March 7). Unlike fry, the 1999 parr and smolt outmigration index estimates are smaller than in 1998. Parr and smolt estimates may be more reliable than the fry estimates because parr and smolt may well be more actively outmigrating than fry. The 1999 parr and smolt confidence intervals do not overlap those of 1998, indicating that the 1999 parr-smolt outmigration was truly less than that of 1998.

The 1999 parr outmigration index is substantially greater than the 1996 and 1997 indices. While the same is not true of the 1999 smolt index, it is true of the 1999 combined parr and smolt
indices which exceeds those of 1996 and 1997. In 1999 there was an extended period when the mean length was very near 80 mm , and the 80 mm demarcation between parr and smolt might be somewhat artificial. The 1999 fry-outmigration index did exceed that of 1996; however, the frymonitoring period started earlier in 1999. There was no monitoring of the fry segment of the outmigration in 1997.

Table A.7.a. Estimates of outmigration indices for fry, parr, and smolt based on the efficiency model with the data sets having 1996 turbidity > 10 included

1996 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits Lower <br> Upper |  | 1998 Report Data Summary Date Domain Estimate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. |  |  |  |  |  |
| Fry | 02/06 | 03/20 | 31,767 | 10,320 | 11,540 | 51,994 | 35,101 | 03/20 | 28,653 |
| Parr | 03/21 | 03/31 | 1,596 | 470 | 675 | 2,516 | 35,145 | 03/31 | 1,465 |
| Smolt | 04/01 | 07/01 | 81,896 | 11,065 | 60,209 | 103,582 | 35,156 | 07/01 | 65,083 |
| All | 02/06 | 07/01 | 115,258 | 15,051 | 85,759 | 144,757 | 35,101 | 07/01 | 95,201 |

1997 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits Lower Upper |  | 1998 Report Data Summary Date Domain Estimate |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. |  |  |  |  |  |
| Fry |  |  |  |  |  |  |  |  |  |
| Parr | 03/19 | 04/05 | 7,011 | 1,037 | 4,979 | 9,043 | 35,508 | 04/01 | 4,724 |
| Smolt | 04/06 | 06/27 | 60,333 | 7,478 | 45,676 | 74,990 | 35,522 | 06/27 | 48,861 |
| All | 03/19 | 06/27 | 67,344 | 8,000 | 51,663 | 83,024 | 35,508 | 06/27 | 53,585 |

1998 Cumulative Outmigration

|  | Date Domain |  | Current |  | Approximate 95\% Confidence Limits |  | 1998 Report Data Summary |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Estimate | S.E. | Lower | Upper | Date D | ain | Estimate |
| Fry | 01/29 | 03/07 | 186,029 | 44,908 | 98,009 | 274,049 | 35,824 | 03/07 | 287,801 |
| Parr | 03/08 | 04/23 | 209,911 | 31,238 | 148,685 | 271,137 | 35,862 | 04/21 | 179,448 |
| Smolt | 04/24 | 07/16 | 197,884 | 37,348 | 124,682 | 271,087 | 35,907 | 07/16 | 183,935 |
| All | 01/29 | 07/16 | 593,825 | 76,373 | 444,133 | 743,516 | 35,824 | 07/16 | 651,184 |

1999 Cumulative Outmigration

|  |  |  |  |  | Approximate 95\% <br> Confidence Limits |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain | Estimate | S.E. | Lower | Upper |  |
| Fry | $01 / 18$ | $03 / 15$ | $1,155,424$ | 145,284 | 870,668 | $1,440,181$ |
| Parr | $03 / 16$ | $05 / 09$ | 92,615 | 11,169 | 70,723 | 114,506 |
| Smolt | $05 / 10$ | $06 / 30$ | 73,003 | 9,679 | 54,031 | 91,975 |
| All | $01 / 18$ | $06 / 30$ | $1,321,042$ | 160,000 | $1,007,443$ | $1,634,642$ |

As mentioned earlier, the 1996 turbidity-greater-than-10 coefficient is used for all years. It may well be that turbidity should not be included in years other than 1996, but there is no statistical basis for either excluding or including it in years other than 1996. Exclusion of the positive turbidity coefficient would lead to a greater estimated outmigration index in 1997, 1998, and 1999. The counts were reexpanded using the efficiencies based on the reestimated coefficients and standard errors in Table A.6.b. which resulted from dropping the 1996 turbidity-greater-than-10 data sets. These expansions are given in Appendices 6.a through 6.d for 1996 through 1999 respectively. The cohort-based cumulative outmigration summary from these appendices is given in Table A.7.b.

The exclusion of the turbidity coefficient increased the 1997 parr cumulative estimate by 83\% (from 7 K to 13 K ) and the 1997 smolt estimate by $71 \%$ (from 60 K to 103 K ). The 1998 fry outmigration index was increased by $146 \%$ (from 186 K to 458 K ), the 1998 par outmigration index by $6 \%$ (from 210 K to 223 K ), and the 1998 smolt outmigration index by $1 \%$ (from 197 K to 200 K ). The 1999 fry outmigration index was increased by $25 \%$ (from $1,155 \mathrm{~K}$ to $1,445 \mathrm{~K}$ ), there being no change in the 1999 parr and smolt outmigration indices because of the failure of the turbidity to reach 10 ntu during their portion of the outmigration.

Table A.7.b. Estimates of outmigration indices for fry, Smolt, and smolt within years based on efficiency model EXCLUDING TURBIDITY >=10 (1996) records

1996 Cumulative Outmigration

|  | Current Date Domain |  | $146,271$ <br> Estimate | S.E. | Approxim Confidenc Lower | te 95\% <br> Limits <br> Upper | Turbidity- <br> Based <br> Estimate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Fry | 02/06 | 03/20 | 29,658 | 13,347 | 3,499 | 55,818 | 31,767 |
| Parr | 03/21 | 03/31 | 2,090 | 654 | 809 | 3,372 | 1,596 |
| Smolt | 04/01 | 07/01 | 124,279 | 29,943 | 65,590 | 182,968 | 81,896 |
| All | 02/06 | 07/01 | 156,028 | 26,758 | 103,582 | 208,474 | 115,258 |

1997 Cumulative Outmigration

|  |  |  |  | Approximate 95\% <br> Confidence Limits |  | Turbidity- |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Based |  |  |  |  |  |  |  |
| Lurrent | Estimate | S.E. | Lower | Upper | Estimate |  |  |
| Fry |  |  |  |  | 0 | 0 |  |
| Parr | $03 / 19$ | $04 / 05$ | 12,818 | 1,368 | 10,136 | 15,500 | 7,011 |
| Smolt | $04 / 06$ | $06 / 27$ | 103,436 | 19,540 | 65,137 | 141,735 | 60,333 |
| All | $03 / 19$ | $06 / 27$ | 116,254 | 20,006 | 77,042 | 155,466 | 67,344 |

1998 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits |  | TurbidityBased |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. | Lower | Upper | Estimate |
| Fry | 01/29 | 03/07 | 458,377 | 77,283 | 306,902 | 609,851 | 186,029 |
| Parr | 03/08 | 04/23 | 223,489 | 31,521 | 161,708 | 285,270 | 209,911 |
| Smolt | 04/24 | 07/16 | 200,320 | 35,291 | 131,150 | 269,491 | 197,884 |
| All | 01/29 | 07/16 | 882,186 | 99,514 | 687,137 | 1,077,234 | 593,825 |

1999 Cumulative Outmigration

|  | Current |  |  |  | Approximate 95\% Confidence Limits |  | TurbidityBased Estimate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Date Domain |  | Estimate | S.E. | Lower | Upper |  |
| Fry | 01/18 | 03/15 | 1,444,960 | 170,901 | 1,109,995 | 1,779,925 | 1,155,424 |
| Parr | 03/16 | 05/09 | 92,615 | 10,470 | 72,094 | 113,136 | 92,615 |
| Smolt | 05/10 | 06/30 | 73,003 | 9,182 | 55,006 | 91,001 | 73,003 |
| All | 01/18 | 06/30 | 1,610,578 | 184,044 | 1,249,852 | 1,971,305 | 1,321,042 |

## Survival

I have taken the turbidity-based efficiencies from Table A.6.a and applied them to Caswell recoveries of survival releases in 1998 and 1999. These are summarized in Table A.8.

Table A. 8. Estimated survival rates of survival to Caswell of releases made some distance upstream of the Caswell traps based on expanded recoveries.

| Year | 1998 |  |  | 1999 |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Knight's | Knight's | Pooled | Knight's | River Mil | River Mi | iver Mil | iver Mil | Pooled |
| Release Site | Ferry (1) | Ferry (2) | Oakdale | Ferry | 38 (1) | 38 (2) | 40 (1) | 40 (2) | Oakdale |
| Number Released (N) | - | 2763 | 26693 | 25536 | 4981 | 5007 | 4975 | 4403 | 6165 |
| Caswell Recoveries (n) | 1 | 6 | 4 | 35 | 8 | 8 | 10 | 7 | 0 |
| Proprotion Recovered (n/N) | - | 0.0022 | 0.0002 | 0.0014 | 0.0016 | 0.0016 | 0.0020 | 0.0016 | 0.0011 |
| Expanded Recoveries (er) | 60.9 | 365.1 | 86.4 | 1347.9 | 308.1 | 308.1 | 385.1 | 269.6 | 269.6 |
| Survival (er/N) | - | 0.1322 | 0.0032 | 0.0528 | 0.0619 | 0.0615 | 0.0774 | 0.0612 | 0.0437 |

What was surprising to me is the relative uniformity of the survival estimates in 1999. Even so, with the exception of the 1999 Knight's Ferry release, the estimates in both years are based on very few recoveries, and far larger release sizes would be needed to obtain estimates that even approach a reasonably precise estimate of survival.

None of the 1999 survival recoveries involved days when the turbidity was 10 or greater, so the issue of a possible turbidity bias does not apply. This, however, is not true of the 1999 recoveries, and survival estimates would be higher if turbidity were not included in the prediction.
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The 1998 and 1999 pooled Oakdale survival estimates given in the Table A. 8 have no inherent meaning since the Oakdale releases were released over an extended period of time over which survival conditions likely varied greatly. Individual Oakdale-release estimates would be meaningless because the number of recoveries from the individual releases is too small, usually 0 .

## Appendix A.1. Variance of the Estimated Outmigration Index

The method of approximating the standard error of the confidence interval of the outmigration passage index is different than that used in the past. The outmigration index for day i is estimated by Equation 1.

Equation 1. $\quad o_{i}=\frac{c_{i}}{e_{i}}=c_{i}\left\{\frac{1}{e_{i}}\right\}$
wherein $c_{i}$ is the screw-trap count on day $i$ and $e_{i}$ is the predicted efficiency (Equation 2) based on the logistic model

Equation 2. $e_{i}=\frac{1}{1+\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right]}$
wherein $\mathrm{x}_{\mathrm{i}}(\mathrm{j})$ is the jth predictor variable (e.g., flow) measured on day i . In the past, the standard error was approximated by using the estimated variance $\left(s^{2}\right)$ of the ratio: $s^{2}\left(o_{i}\right)=s^{2}\left(c_{i} / o_{i}\right)$; however, the estimate of the variance of a ratio was an approximation. In this report, the variance estimator used was the estimated variance of a product, which can be more accurately estimated that the variance of a ratio. Substituting Equation 2 into Equation 1, puts the equation into product form (Equation 3):

Equation 3.

$$
\begin{aligned}
& o_{i}=\frac{c_{i}}{\frac{1}{1+\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right]}} \\
& =c_{i}\left\{1+\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right]\right\}
\end{aligned}
$$

The cumulative outmigration index can be written
Equation 4. $\quad \sum_{i} o_{i}=\sum_{i} c_{i}\left\{\frac{1}{e_{i}}\right\}$
The variance of the cumulative outmigration index is of the form given in Equation 5:
Equation 5. $\quad s^{2}\left[\sum_{i} o_{i}\right]=s^{2}\left[\sum_{i} c_{i}\left\{\frac{1}{e_{i}}\right\}\right]=\sum_{i} s^{2}\left[c_{i}\left\{\frac{1}{e_{i}}\right\}\right]+\sum_{i} \sum_{i \neq i} s\left[c_{i}\left\{\frac{1}{e_{i}}\right\}, c_{i^{\prime}}\left\{\frac{1}{e_{i^{\prime}}}\right\}\right]$

The form $s^{2}(y)$ representing the estimated variance of $y$ and $s(x, y)$ representing the estimated covariance between $x$ and $y$.

In the above equation, $s^{2}\left[c_{i}\left\{1 / e_{i}\right\}\right]$, a variance of a product, can written in the form:
Equation 5.a. $\quad s^{2}\left[c_{i}\left\{\frac{1}{e_{i}}\right\}\right]=c_{i}^{2} * s^{2}\left[\frac{1}{e_{i}}\right]+s^{2}\left[c_{i}\right] *\left[\frac{1}{e_{i}}\right]^{2}-s^{2}\left[c_{i}\right] * s^{2}\left[\frac{1}{e_{i}}\right]$
Which is unbiased if the count, $\mathrm{c}_{\mathrm{i}}$, and the expansion factor, $1 / \mathrm{e}_{\mathrm{i}}$, are stochastically independent and as long as their variance and covariance estimates are unbiased.

Within Equation 5.a, the variance of the count on day $\mathrm{i}, \mathrm{s}^{2}\left[\mathrm{c}_{\mathrm{i}}\right]$, is estimated by Equation 5.a.1, the variance among the counts involving the count on day $\mathrm{i}, \mathrm{c}_{\mathrm{i}}$, and involving the counts on the immediate preceding and following days ${ }^{9}, \mathrm{c}_{\mathrm{i}-1}$ and $\mathrm{c}_{\mathrm{i}+1}$, respectively

Equation 5.a.1. $\quad s^{2}\left[c_{i}\right]=\frac{\left[c_{i-1}-\overline{c(i)}\right]^{2}+\left[c_{i}-\overline{c(i)}\right]^{2}+\left[c_{i+1}-\overline{c(i)}\right]^{2}}{3-1}$

$$
\text { wherein } \overline{c(i)}=\frac{c_{i-1}+c_{i}+c_{i+1}}{3}
$$

and, within Equation 5.a, the variance of the inverse of the efficiency, $\mathrm{s}^{2}\left[1 / \mathrm{e}_{\mathrm{i}}\right]$, is estimated by Equation 5.a. 2 using the delta method

Equation 5.a.2.

$$
\begin{aligned}
s^{2}\left[\frac{1}{e_{i}}\right] & =s^{2}\left[1+\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right]\right] \\
& =\exp ^{2}\left[-b(0)-b(1) x_{i}(1)-b(q) x_{i}(q)\right] * s^{2}\left[b(0)+b(1) x_{i}(1)+\ldots+b(q) x_{i}(q)\right]
\end{aligned}
$$

The delta method is also used to estimate the covariance terms in equation 5.

## Equation 5.b.

9
If the day is the first day of monitoring, then the variance is the variance between the counts on that day and the following day; if the day is the last day of monitoring, then the variance is the variance between the counts on that day and the preceding day.

$$
\begin{gathered}
s\left[\frac{1}{e_{i}}, \frac{1}{e_{i}}\right]= \\
s\left\{c_{i}\left[1+\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right]\right], c_{i^{\prime}}\left[1+\exp \left[-b(0)-b(1) x_{i i}(1)-\ldots-b(q) x_{i^{\prime}}(q)\right]\right]\right\}= \\
c c_{i^{\prime}}\left\{\exp \left[-b(0)-b(1) x_{i}(1)-\ldots-b(q) x_{i}(q)\right] * \exp \left[-b(0)-b(1) x_{i j}(1)-\ldots-b(q) x_{i_{i}}(q)\right]\right\} * \\
s\left\{\left[b(0)+b(1) x_{i}(1)+\ldots+b(q) x_{i}(q)\right],\left[b(0)+b(1) x_{i^{\prime}}(1)+\ldots+b(q) x_{i i}(q)\right]\right\}
\end{gathered}
$$

Within equation 5.a. 2

$$
\begin{gathered}
s^{2}\left[b(0)+b(1) x_{i}(1)+\ldots+b(q) x_{i}(q)\right]= \\
\left\{s^{2}[b(0)]+x_{i}^{2}(1) s^{2}[b(1)]+\ldots+x_{i}(q)_{i}^{2} s^{2}[b(q)]\right\}+ \\
\left\{x_{i}(1) * s[b(0), b(1)]+\ldots+x_{i}(q) * s[b(0), b(q)]\right\}+ \\
\left\{\left[x_{i}(1) * x_{i}(2)\right] * s[b(1), b(2)]+\ldots\left[x_{i}(1) * x_{i}(q)\right] * s[b(1), b(q)]+\right. \\
\left.\ldots\left[x_{i}(2) * x_{i}(q)\right] * s[b(2), b(q)]+\ldots\right\}
\end{gathered}
$$

and within Equation 5.b.

$$
\begin{gathered}
s\left\{\left[b(0)+b(1) x_{i}(1)+\ldots+b(q) x_{i}(q)\right],\left[b(0)+b(1) x_{i^{\prime}}(1)+\ldots+b(q) x_{i^{\prime}}(q)\right]\right\}= \\
\left\{s^{2}[b(0)]+x_{i}(1) x_{i^{\prime}}(1) s^{2}[b(1)]+\ldots+x_{i}(q) x_{i^{\prime}}(q) s^{2}[b(q)]\right\}+ \\
\left\{\left[x_{i}(1)+x_{i^{\prime}}(1)\right] * s[b(0), b(1)]+\ldots+\left[x_{i}(q)+x_{i^{\prime}}(q)\right] * s[b(0), b(q)]\right\}+ \\
\left\{\left[x_{i}(1) * x_{i^{\prime}}(2)+x_{i}(2) * x_{i^{\prime}}(1)\right] * s[b(1), b(2)]+\ldots\left[x_{i}(1) * x_{i^{\prime}}(q)+x_{i}(q) * x_{i^{\prime}}(1)\right] * s[b(1), b(q)]+\right. \\
\left.\ldots\left[x_{i}(2) * x_{i^{\prime}}(q)+x_{i}(q) * x_{i^{\prime}}(2)\right] * s[b(2), b(q)]+\ldots\right\}
\end{gathered}
$$

The variances of and covariances among the logistic regression coefficient estimates $\left\{\mathrm{s}^{2}[\mathrm{~b}(\mathrm{j})]\right.$ and $\left.\mathrm{s}\left[\mathrm{b}(\mathrm{j}), \mathrm{b}\left(\mathrm{j}{ }^{\prime}\right)\right]\right\}$ were obtained from the variance-covariance-matrix output from logistic regression software; however, the output matrix assumes that the distribution of the efficiencies around the true predictor is binomial. The residual deviances suggests this is not likely to be the case; therefore, the variance-covariance matrix was multiplied by the mean deviance and the standard errors were multiplied by the square root of the mean deviance to correct a greater-than-assumed variation due to contagious movement and possible lack of fit of the model.

## Appendix A.2. Logistic Analysis of Variation used for Preliminary Selection of Model

| MODEL | Residual Deviance (Dev) | Degrees of Freedom (DF) | Dev/DF | Compared to Model | Change <br> Dev | DF | Dev/DF | $\begin{gathered} \text { F- } \\ \text { ratio } \end{gathered}$ | $\begin{gathered} \text { Type I } \\ \text { P } \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| YEAR: 1996-1997 |  |  |  |  |  |  |  |  |  |
| 1. Constant | 400.02 | 17 | 23.53 |  |  |  |  |  |  |
| 2.a. Age Cohort+Constant | 329.46 | 15 | 21.96 | 1 | 70.56 | 2 | 35.28 | 1.61 | 0.2333 |
| 2.b. Recovery Size+Constant | 320.52 | 16 | 20.03 | 1 | 79.50 | 1 | 79.50 | 3.97 | 0.0637 |
| 2.c. Age Cohort+Recovery Size+Constant | 310.62 | 14 | 22.19 | 2.a. | 18.84 | 1 | 18.84 | 0.85 | 0.3724 |
|  |  |  |  | 2.b. | 9.90 | 2 | 4.95 | 0.22 | 0.8028 |
| 3. Average Flow+Constant | 350.76 | 16 | 21.92 | 1 | 49.26 | , | 49.26 | 2.25 | 0.1533 |
| 4. Average Flow+Recovery Size+Constant | 185.39 | 15 | 12.36 | 2.b. | 135.13 | 1 | 135.13 | 10.93 | 0.0048 |
|  |  |  |  | 3 | 165.37 | 1 | 165.37 | 13.38 | 0.0023 |
| 5. Average Flow+Recovery Size+Turbidity > 10+Constant | 68.92 | 14 | 4.92 | 4 | 116.47 | 1 | 116.47 | 23.66 | 0.0003 |
| 6. Drop Recovery Size from 5. | 86.93 | 15 | 5.80 | 5 | 18.01 | 1 | 18.01 | 3.66 | 0.0765 |
| 7. Drop Average Flow from 5. | 191.29 | 15 | 12.75 | 5 | 122.37 | 1 | 122.37 | 24.86 | 0.0001 |
| YEAR: 1998 |  |  |  |  |  |  |  |  |  |
| 1. Constant | 159.21 | 15 | 10.61 |  |  |  |  |  |  |
| 2.a. Age Cohort+Constant | 69.02 | 13 | 5.31 | 1 | 90.19 | 2 | 45.10 | 8.49 | 0.0044 |
| 2.b. Recovery Size+Constant | 55.98 | 14 | 4.00 | 1 | 103.23 | 1 | 103.23 | 25.82 | 0.0002 |
| 2.c. Age Cohort+Recovery Size+Constant | 44.07 | 12 | 3.67 | 2.a. | 24.95 | 1 | 24.95 | 6.79 | 0.0230 |
|  |  |  |  | 2.b. | 11.91 | 2 | 5.96 | 1.62 | 0.2380 |
| 3. Average Flow+Constant | 158.32 | 14 | 11.31 | 1 | 0.89 | 1 | 0.89 | 0.08 | 0.7832 |
| 4. Average Flow+Recovery Size+Constant | 51.99 | 13 | 4.00 | 2.b. | 3.99 | 1 | 3.99 | 1.00 | 0.3361 |
|  |  |  |  | 3 | 106.33 | 1 | 106.33 | 26.59 | 0.0002 |
| YEAR: 1999 |  |  |  |  |  |  |  |  |  |
| 1. Constant | 38.57 | 8 | 4.82 |  |  |  |  |  |  |
| 2.a. Age Cohort+Constant | 29.75 | 7 | 4.25 | 1 | 8.82 | , | 8.82 | 2.08 | 0.1929 |
| 2.b. Recovery Size+Constant | 29.25 | 7 | 4.18 | 1 | 9.32 | 1 | 9.32 | 2.23 | 0.1790 |
| 2.c. Age Cohort+Recovery Size+Constant | 25.28 | 6 | 4.21 | 2.a. | 4.47 | 1 | 4.47 | 1.06 | 0.3427 |
|  |  |  |  | 2.b. | 3.97 | 1 | 3.97 | 0.94 | 0.3692 |
| 3. Average Flow+Constant | 24.35 | 7 | 3.48 | 1 | 14.22 | 1 | 14.22 | 4.09 | 0.0829 |
| 4. Average Flow+Recovery Size+Constant | 24.32 | 6 | 4.05 | 2.b. | 4.93 | 1 | 4.93 | 1.22 | 0.3124 |


| RESIDUAL FROM FULL MODEL AMONG DAYS VERSUS WITHIN DAY VARIATION |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Residual |  |  | F-Ratio |  |
|  |  |  |  | Pooled Full | Type I |
|  | Dev | DF | Dev/Df | to within Day | P |
| Full Model Fit among days with recovery data and fish size pooled over releases within days |  |  |  |  |  |
| 1996-97: Cons, flow, size (pooled), turbidity >10 | 31 | 6 | 5.17 |  |  |
| 1998: Cons, flow, size (pooled) | 22.88 | 4 | 5.72 |  |  |
| 1999: Cons, flow, size (pooled) | 13.87 | 5 | 2.77 |  |  |
| Pooled Full Model mong Days | 67.75 | 15 | 4.52 | 1.17 | 0.3684 |
| Variation among releases within days |  |  |  |  |  |
| among releases within day 1996-97 | 36.94 | 8 | 4.62 |  |  |
| among releases within day 1998 | 21.85 | 9 | 2.43 |  |  |
| among releases within day 1999 | 10.43 | 1 | 10.43 |  |  |
| Pooled within Day | 69.22 | 18 | 3.85 |  |  |

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## Appendix A.3. Coefficients from Full Model

| Predictor | Estimate | Standard <br> Error | t-ratio | Type I <br> P |
| :---: | :---: | :---: | :---: | :---: |
| 1997-97 Intercept | -1.423720 | 0.468668 | -3.04 | 0.0046 |
| 1998 Intercept | -1.709590 | 0.633254 | -2.70 | 0.0109 |
| 1999 Intercept | -4.041050 | 0.839527 | -4.81 | 0.0000 |
| 1996 1997 Flow | -0.000829 | 0.000164 | -5.05 | 0.0000 |
| 1998 Flow | -0.000254 | 0.000268 | -0.95 | 0.3506 |
| 1999 Flow | 0.000158 | 0.000154 | 1.03 | 0.3099 |
| 1996-97 size | -0.010380 | 0.005156 | -2.01 | 0.0523 |
| 1998 size | -0.023080 | 0.004820 | -4.79 | 0.0000 |
| 1999 size | 0.000733 | 0.008152 | 0.09 | 0.9289 |
| Turbidity $>10$ | 0.074650 | 0.014088 | 5.30 | 0.0000 |


|  |  |
| :---: | :---: |
| Deviance | 145.23 |
| P(Type 1) | 0.0000 |
| Degrees of Freedom | 33 |
| Dev/DF | 4.40 |


| Coefficient | Type 1 Probabilities for Comparing Coefficients among Years |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Year | $\begin{gathered} \text { Intercept } \\ \text { 1996-1997 } \end{gathered}$ | 1998 | $\begin{gathered} \text { Flow } \\ \text { 1996-1997 } \end{gathered}$ | 1998 | $\begin{gathered} \text { Size } \\ \text { 1996-1997 } \end{gathered}$ | 1998 |
| Intercept | 1998 | 0.7190 |  |  |  |  |  |
|  | 1999 | 0.0103 | 0.0336 |  |  |  |  |
| Flow | 1998 |  |  | 0.0761 |  |  |  |
|  | 1999 |  |  | 0.0001 | 0.1912 |  |  |
| Size | 1998 |  |  |  |  | 0.0024 |  |
|  | 1999 |  |  |  |  | 0.2575 | 0.0170 |

Appendix A.4.a. Variables used to estimate 1996 Outmigration

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| Cohort | DATE | COUNT |  | FLOW (cfs) |  |  | Previous Day | Average | SIZE(Length, mm) | TURBIDITY (NTU) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | North Trap | South Trap | Total |  | Current Day |  |  |  | Current Day | Used for fit |
|  | 04/05/96 | 2 | 7 | 9 | 1.4 | 1809.0 | 1788.0 | 1798.5 | 82.78 | 6.00 |  |
|  | 04/06/96 | 5 | 9 | 14 |  | 1791.0 | 1809.0 | 1800.0 | 87.50 | 6.20 |  |
|  | 04/07/96 | 1 | 12 | 13 |  | 1780.0 | 1791.0 | 1785.5 | 76.92 | 5.90 |  |
|  | 04/08/96 | 1 | 0 | 1 |  | 1779.0 | 1780.0 | 1779.5 | 81.00 | 5.10 |  |
|  | 04/09/96 | 3 | 5 | 8 |  | 1775.0 | 1779.0 | 1777.0 | 86.17 | 4.20 |  |
|  | 04/10/96 | 0 | 4 | 4 |  | 1776.0 | 1775.0 | 1775.5 | 80.75 | 6.50 |  |
|  | 04/11/96 | 1 | 1 | 2 |  | 1791.0 | 1776.0 | 1783.5 | 85.00 | 4.20 |  |
|  | 04/12/96 | 2 | 7 | 9 | 1.4 | 1731.0 | 1791.0 | 1761.0 | 82.56 | 4.60 |  |
|  | 04/13/96 | 2 | 0 | 2 |  | 1598.0 | 1731.0 | 1664.5 | 80.50 | 9.90 |  |
|  | 04/14/96 | $0.000 \quad 1.1$ | 0 | 0.000 |  | 1595.0 | 1598.0 | 1596.5 | 83.00 3 | 5.20 |  |
|  | 04/15/96 | 2.0 | 8 | 10 |  | 1599.0 | 1595.0 | 1597.0 | 85.50 | 9.70 |  |
|  | 04/16/96 | 0.0 | 2 | 2 |  | 1656.0 | 1599.0 | 1627.5 | 97.50 | 5.70 |  |
|  | 04/17/96 | 2.0 | $3.209 \quad 1.2$ | 5.209 |  | 1706.0 | 1656.0 | 1681.0 | 91.33 | 9.30 |  |
|  | 04/18/96 | 3.0 | 3 | 6 |  | 1711.0 | 1706.0 | 1708.5 | 84.67 | 7.20 |  |
|  | 04/19/96 | 3.0 | 12 | 15 | 1.4 | 1679.0 | 1711.0 | 1695.0 | 86.20 | 6.30 |  |
|  | 04/20/96 | 0.0 | 1 | 1 | 1.4 | 1670.0 | 1679.0 | 1674.5 | 89.00 | 5.20 |  |
|  | 04/21/96 | 11.0 | 11 | 22 |  | 1675.0 | 1670.0 | 1672.5 | 89.77 | 5.60 |  |
|  | 04/22/96 | 15.0 | 21 | 36 |  | 1673.0 | 1675.0 | 1674.0 | 91.08 | 4.70 |  |
|  | 04/23/96 | 20.0 | $32.092 \quad 1.2$ | 52.092 |  | 1668.0 | 1673.0 | 1670.5 | 89.65 | 6.00 |  |
|  | 04/24/96 | 17.0 | 21 | 38 |  | 1673.0 | 1668.0 | 1670.5 | 89.66 | 6.40 |  |
|  | 04/25/96 | 18.0 | 21 | 39 |  | 1676.0 | 1673.0 | 1674.5 | 92.23 | 7.80 |  |
|  | 04/26/96 | 12.0 | 26 | 38 |  | 1676.0 | 1676.0 | 1676.0 | 91.19 | 5.70 |  |
|  | 04/27/96 | 36.0 | 59 | 95 |  | 1662.0 | 1676.0 | 1669.0 | 90.97 | 5.10 |  |
|  | 04/28/96 | 34.0 | 75 | 109 | 1.4 | 1668.0 | 1662.0 | 1665.0 | 91.68 | 5.90 |  |
|  | 04/29/96 | 26.0 | 63 | 89 |  | 1684.0 | 1668.0 | 1676.0 | 91.89 | 6.90 |  |
|  | 04/30/96 | $47.987 \quad 1.1$ | 77 | 124.987 |  | 1683.0 | 1684.0 | 1683.5 | 91.02 | 9.10 |  |
|  | 05/01/96 | 36 | $57.766 \underline{1.2}$ | 93.766 |  | 1684.0 | 1683.0 | 1683.5 | 91.21 | 9.40 |  |
|  | 05/02/96 | 25 | 59 | 84 |  | 1680.0 | 1684.0 | 1682.0 | 93.40 | $9.80 \quad 4$ |  |
|  | 05/03/96 | 11 | 33 | 75.237 | 1.3.b. | 1659.0 | 1680.0 | 1669.5 | 92.88 | 10.20 | 10.20 |
|  | 05/04/96 | 21 | 46 | 67 |  | 1674.0 | 1659.0 | 1666.5 | 90.47 | 9.80 |  |
|  | 05/05/96 | 32 | 75 | 107 |  | 1662.0 | 1674.0 | 1668.0 | 93.48 | 9.90 |  |
|  | 05/06/96 | 35 | 38 | 73 |  | 1640.0 | 1662.0 | 1651.0 | 90.80 | 9.20 |  |
|  | 05/07/96 | 20 | 22 | 42 |  | 1664.0 | 1640.0 | 1652.0 | 92.10 | 8.40 |  |
|  | 05/08/96 | 25 | 22 | 47 |  | 1650.0 | 1664.0 | 1657.0 | 91.91 | 9.20 |  |
|  | 05/09/96 | 19 | 28 | 47 |  | 1663.0 | 1650.0 | 1656.5 | 91.36 | 9.00 |  |
|  | 05/10/96 | 20 | $32.092 \quad 1.2$ | 52.092 |  | 1667.0 | 1663.0 | 1665.0 | 90.57 | 8.80 |  |
|  | 05/11/96 | 24 | 36 | 60 |  | 1653.0 | 1667.0 | 1660.0 | 91.84 | 8.70 |  |
|  | 05/12/96 | 8 | 12 | 20 |  | 1644.0 | 1653.0 | 1648.5 | 91.08 | 9.00 |  |
|  | 05/13/96 | 0 | 6 | 35.842 | 1.3.b | 1654.8 2.1 | 1644.0 | 1649.4 | 92.02 - | 8.80 |  |
|  | 05/14/96 | 13 | $20.860 \quad 1.2$ | 33.860 |  | $1665.6 \quad 2.1$ | 1654.8 2.2 | 1660.2 | 92.95 | 6.80 |  |
|  | 05/15/96 | 5 | 0 | 28.721 | 1.3.b | $1676.4 \underline{2.1}$ | 1665.6 2.2 | 1671.0 | 98.20 | 7.10 |  |
|  | 05/16/96 | 3 | 16 | 19 | 1.4 | 1687.2 2.1 | 1676.4 2.2 | 1681.8 | 91.21 | 6.90 |  |
|  | 05/17/96 | 6 | 4 | 10 | 1.4 | 1698.0 | 1687.2 2.2 | 1692.6 | 93.70 | 7.30 |  |
|  | 05/18/96 | 0 | 14 | 14 | 1.4 | 1658.0 | 1698.0 | 1678.0 | 95.79 | 7.10 |  |
|  | 05/19/96 | 1 | 9 | 10 | 1.4 | 1693.0 | 1658.0 | 1675.5 | 99.50 | 6.10 |  |
|  | 05/20/96 | 6 | 13 | 19 |  | 1697.0 | 1693.0 | 1695.0 | 95.00 | 6.20 |  |
|  | 05/21/96 | 3 | 20 | 23 |  | 1670.0 | 1697.0 | 1683.5 | 95.45 | 5.80 |  |
|  | 05/22/96 | 4 | $6.418 \quad 1.2$ | 10.418 |  | 1525.0 | 1670.0 | 1597.5 | 94.12 | 5.40 |  |
|  | 05/23/96 | 6 | 3 | 9 |  | 1151.0 | 1525.0 | 1338.0 | 95.89 | 6.40 |  |
|  | 05/24/96 | 6 | 12 | 18 |  | 936.0 | 1151.0 | 1043.5 | 94.61 | 7.90 |  |
|  | 05/25/96 | $12.464 \quad 1.1$ | 20 | 32.464 |  | 901.0 | 936.0 | 918.5 | 95.10 | 9.80 |  |
|  | 05/26/96 | 28 | 24 | 52 |  | 921.0 | 901.0 | 911.0 | 95.02 | 8.90 |  |
|  | 05/27/96 | 13 | 17 | 30 |  | 955.0 | 921.0 | 938.0 | 93.26 | 6.80 |  |
|  | 05/28/96 | 10 | 5 | 15 |  | 958.0 | 955.0 | 956.5 | 94.57 | 6.60 |  |
|  | 05/29/96 | 10 | 12 | 22 |  | 935.0 | 958.0 | 946.5 | 92.95 | 7.40 |  |
|  | 05/30/96 | 3 | 6 | 9 |  | 935.0 | 935.0 | 935.0 | 93.33 | 8.30 |  |
|  | 05/31/96 | 5 | 5 | 10 |  | 939.0 | 935.0 | 937.0 | 95.90 | 7.90 |  |
|  | 06/01/96 | 7 | 3 | 10 |  | 945.0 | 939.0 | 942.0 | 98.00 | 8.60 |  |
|  | 06/02/96 | 5 | 6 | 11 |  | 939.0 | 945.0 | 942.0 | 97.27 | 9.80 |  |
|  | 06/03/96 | 1 | 1 | 2 |  | 933.0 | 939.0 | 936.0 | 92.00 | 7.70 |  |
|  | 06/04/96 | 2 | $3.209 \quad 1.2$ | 5.209 |  | 936.0 | 933.0 | 934.5 | 99.00 | 6.80 |  |
|  | 06/05/96 | 3 | 4 | 7 | 1.4 | 933.0 | 936.0 | 934.5 | 102.00 | 6.60 |  |

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## Appendix A.4.b. Variables used to estimate 1997 Outmigration

|  |  | UNT |  |  | LOW (cfs) |  |  | SIZE | TURBIDIT | Y (NTU) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South Trap | Total | Current Day | Previous Day | Average | (Length, mm) | Current Day | Used for fit |
| Parr | 03/19/97 | 0 | 15 | 15 | 1618.0 | 1647.0 | 1632.5 | 64.47 | 11.80 | 11.80 |
|  | 03/20/97 | 2 | 15 | 17 | 1631.0 | 1618.0 | 1624.5 | 73.29 | 10.40 | 10.40 |
|  | 03/21/97 | 0 | 35 | 35 | 1645.0 | 1631.0 | 1638.0 | 71.77 | 12.80 | 12.80 |
|  | 03/22/97 | 0 | 36 | 36 | 1558.0 | 1645.0 | 1601.5 | 73.06 | 11.10 | 11.10 |
|  | 03/23/97 | 2 | 46 | 48 | 1362.0 | 1558.0 | 1460.0 | 74.85 | 10.80 | 10.80 |
|  | 03/24/97 | 3 | 39 | 42 | 1175.0 | 1362.0 | 1268.5 | 73.98 | 10.60 | 10.60 |
|  | 03/25/97 | 1 | 31 | 32 | 876.0 | 1175.0 | 1025.5 | 73.53 | 10.20 | 10.20 |
|  | 03/26/97 | 0 | 30 | 30 | 524.0 | 876.0 | 700.0 | 76.37 | 12.10 | 12.10 |
|  | 03/27/97 | 0 | 22 | 22 | 621.0 | 524.0 | 572.5 | 77.05 | 14.00 | 14.00 |
|  | 03/28/97 | 2 | 26 | 28 | 595.0 | 621.0 | 608.0 | 77.18 | 13.40 | 13.40 |
|  | 03/29/97 | 1 | 20 | 21 | 601.0 | 595.0 | 598.0 | 73.43 | 10.70 | 10.70 |
|  | 03/30/97 | 5 | $47.500 \quad 1.2$ | 52.500 | 605.0 | 601.0 | 603.0 | 81.78 | 8.70 |  |
|  | 03/31/97 | 9 | 21 | 30 | 616.0 | 605.0 | 610.5 | 79.73 | 10.10 | 10.10 |
|  | 04/01/97 | 6 | 39 | 45 | 618.0 | 616.0 | 617.0 | 76.27 | 10.80 | 10.80 |
|  | 04/02/97 | $2.316 \quad 1.1$ | 22 | 24.316 | 614.0 | 618.0 | 616.0 | 80.18 | 10.50 4 | 10.50 |
|  | 04/03/97 | 5 | 22 | 27 | 597.0 | 614.0 | 605.5 | 82.26 | 10.20 | 10.20 |
|  | 04/04/97 | 5 | 23 | 28 | 599.0 | 597.0 | 598.0 | 78.50 | 9.40 |  |

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| COUNT |  |  |  | FLOW (cfs) |  |  |  |  |  |  | SIZE | TURBIDITY (NTU) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North T |  | South T |  | Total |  | Current Day | Previous Day | Average | (Length, mm) | Current Day | Used for fit |
| Smolt | 04/05/97 | 12 |  | 36 |  | 48 |  | 602.0 | 599.0 | 600.5 | 79.19 | 8.70 |  |
|  | 04/06/97 | 12 |  | 39 |  | 51 |  | 597.0 | 602.0 | 599.5 | 81.02 | 9.30 |  |
|  | 04/07/97 | 11 |  | 28 |  | 39 |  | 590.0 | 597.0 | 593.5 | 83.18 | 6.30 |  |
|  | 04/08/97 | 2 |  | 24 |  | 26 |  | 602.0 | 590.0 | 596.0 | 83.54 | 7.80 |  |
|  | 04/09/97 | 8.444 | 1.1 | 38 |  | 46.444 |  | 599.0 | 602.0 | 600.5 | 80.76 | 8.40 |  |
|  | 04/10/97 | 21 |  | 39 |  | 60 |  | 598.0 | 599.0 | 598.5 | 80.42 | 8.80 |  |
|  | 04/11/97 | 8 |  | 36.000 | 1.2 | 44.000 |  | 589.0 | 598.0 | 593.5 | 83.84 | 8.40 |  |
|  | 04/12/97 | 11 |  | 38 |  | 49 |  | 730.0 | 589.0 | 659.5 | 83.37 | 7.80 |  |
|  | 04/13/97 | 11 |  | 34 |  | 45 |  | 1164.0 | 730.0 | 947.0 | 82.86 | 10.30 | 10.30 |
|  | 04/14/97 | 17 |  | 51 |  | 68 |  | 1711.0 | 1164.0 | 1437.5 | 82.78 | 12.50 | 12.50 |
|  | 04/15/97 | 23 |  | 103.500 | 1.2 | 126.500 |  | 1707.0 | 1711.0 | 1709.0 | 81.32 | 12.50 | 12.50 |
|  | 04/16/97 | 19 |  | 85.500 | $\underline{1.2}$ | 104.500 |  | 1651.0 | 1707.0 | 1679.0 | 84.22 | 11.20 | 11.20 |
|  | 04/17/97 | 11 |  | 70 |  | 81 |  | 1668.0 | 1651.0 | 1659.5 | 84.68 | 11.30 | 11.30 |
|  | 04/18/97 | 3 |  | 40 |  | 43 |  | 1684.0 | 1668.0 | 1676.0 | 83.63 | 12.00 | 12.00 |
|  | 04/19/97 | 2 |  | 20 |  | 22 | 1.4 | 1680.0 | 1684.0 | 1682.0 | 80.86 | 13.90 | 13.90 |
|  | 04/20/97 | 4 |  | 47 |  | 51 | 1.4 | 1695.0 | 1680.0 | 1687.5 | 85.02 | 13.40 | 13.40 |
|  | 04/21/97 | 3 |  | 25 |  | 28 |  | 1685.0 | 1695.0 | 1690.0 | 83.36 | 12.40 | 12.40 |
|  | 04/22/97 | 8.000 | 1.1 | 36 |  | 44.000 |  | 1668.0 | 1685.0 | 1676.5 | 85.39 | 11.40 | 11.40 |
|  | 04/23/97 | 2 |  | 8 |  | 10 |  | 1679.0 | 1668.0 | 1673.5 | 86.80 | 11.00 | 11.00 |
|  | 04/24/97 | 0 |  | 9 |  | 9 | 1.4 | 1680.0 | 1679.0 | 1679.5 | 85.00 | 10.50 | 10.50 |
|  | 04/25/97 | 3 |  | 23 |  | 26 |  | 1686.0 | 1680.0 | 1683.0 | 84.54 | 10.00 | 10.00 |
|  | 04/26/97 | 3 |  | 29 |  | 32 |  | 1691.0 | 1686.0 | 1688.5 | 85.16 | 10.30 | 10.30 |
|  | 04/27/97 | 2 |  | 13 |  | 15 |  | 1716.0 | 1691.0 | 1703.5 | 84.53 | 10.10 | 10.10 |
|  | 04/28/97 | 3 |  | 13.500 | 1.2 | 16.500 |  | 1685.0 | 1716.0 | 1700.5 | 90.00 | 9.90 |  |
|  | 04/29/97 | 6 |  | 15 |  | 21 |  | 1686.0 | 1685.0 | 1685.5 | 85.57 | 9.20 |  |
|  | 04/30/97 | 6 |  | 21 |  | 27 |  | 1680.0 | 1686.0 | 1683.0 | 87.56 | 8.90 |  |
|  | 05/01/97 | 0 |  | 3 |  | 3 |  | 1682.0 | 1680.0 | 1681.0 | 93.00 | 9.40 |  |
|  | 05/02/97 | 2 |  | 13 |  | 15 |  | 1672.0 | 1682.0 | 1677.0 | 86.60 | 9.70 |  |
|  | 05/03/97 | 11 |  | 31 |  | 42 |  | 1653.0 | 1672.0 | 1662.5 | 86.33 | 9.50 |  |
|  | 05/04/97 | 5 |  | 23 |  | 28 |  | 1648.0 | 1653.0 | 1650.5 | 88.71 | 9.30 |  |
|  | 05/05/97 | 13 |  | 34 |  | 47 |  | 1659.0 | 1648.0 | 1653.5 | 86.26 | 9.40 |  |
|  | 05/06/97 | 1 |  | 8 |  | 9 | 1.4 | 1633.0 | 1659.0 | 1646.0 | 91.00 | 8.90 |  |
|  | 05/07/97 | 3 |  | 29 |  | 32 |  | 1653.0 | 1633.0 | 1643.0 | 90.53 | 9.00 |  |
|  | 05/08/97 | 5 |  | 24 |  | 29 |  | 1639.0 | 1653.0 | 1646.0 | 88.52 | 9.20 |  |
|  | 05/09/97 | 4 |  | 27 |  | 31 |  | 1662.0 | 1639.0 | 1650.5 | 87.65 | 8.80 |  |
|  | 05/10/97 | 3 |  | 20 |  | 23 |  | 1652.0 | 1662.0 | 1657.0 | 86.13 | 9.10 |  |
|  | 05/11/97 | 1 |  | 20 |  | 21 | 1.4 | 1639.0 | 1652.0 | 1645.5 | 89.33 | 8.90 |  |
|  | 05/12/97 | 7 |  | 31.500 | 1.2 | 38.500 |  | 1642.0 | 1639.0 | 1640.5 | 86.04 | 8.80 |  |
|  | 05/13/97 | 7 |  | 31.500 | $\underline{1.2}$ | 38.500 |  | 1581.0 | 1642.0 | 1611.5 | 88.14 | 8.60 |  |
|  | 05/14/97 | 8 |  | 23 |  | 31 |  | 1038.0 | 1581.0 | 1309.5 | 89.61 | 8.70 |  |
|  | 05/15/97 | 0 |  | 19 |  | 35.566 | 1.3.a. | 1571.0 | 1038.0 | 1304.5 | 90.89 | 9.00 |  |
|  | 05/16/97 | 13 |  | 58.500 | 1.2 | 71.500 |  | 1613.0 | 1571.0 | 1592.0 | 90.73 | 9.40 |  |
|  | 05/17/97 | 5 |  | 22.500 | $\underline{1.2}$ | 27.500 |  | 1602.0 | 1613.0 | 1607.5 | 89.20 | 9.30 |  |
|  | 05/18/97 | 7 |  | 35 |  | 42 | 1.4 | 1616.0 | 1602.0 | 1609.0 | 89.78 | 8.90 |  |
|  | 05/19/97 | 7 |  | 55 |  | 62 |  | 1621.0 | 1616.0 | 1618.5 | 89.36 | 9.10 |  |
|  | 05/20/97 | 8 |  | 36.000 | 1.2 | 44.000 |  | 1598.0 | 1621.0 | 1609.5 | 88.95 | 9.20 |  |
|  | 05/21/97 | 0 |  | 23 |  | 23 |  | 1600.0 | 1598.0 | 1599.0 | 88.43 | 9.00 |  |
|  | 05/22/97 | 2 |  | 28 |  | 30 |  | 1607.0 | 1600.0 | 1603.5 | 91.07 | 8.90 |  |
|  | 05/23/97 | 0 |  | 0 |  | 33.785 | 1.3.a. | 1506.0 | 1607.0 | 1556.5 | 92.33 3 | 8.60 |  |
|  | 05/24/97 | 3 |  | 9 |  | 34.612 | 1.3.a. | 1218.0 | 1506.0 | 1362.0 | 93.58 | 9.10 |  |
|  | 05/25/97 | 6 |  | 25 |  | 31 |  | 1233.0 | 1218.0 | 1225.5 | 90.45 | 9.30 |  |
|  | 05/26/97 | 7 |  | 44 |  | 51 |  | 1224.0 | 1233.0 | 1228.5 | 88.58 | 9.40 |  |
|  | 05/27/97 | 1 |  | 10 |  | 11 | 1.4 | 1398.0 | 1224.0 | 1311.0 | 90.27 | 9.70 |  |
|  | 05/28/97 | 6 |  | 27.000 | 1.2 | 33.000 |  | 1608.0 | 1398.0 | 1503.0 | 90.17 | 9.60 |  |
|  | 05/29/97 | 6 |  | 36 |  | 42 |  | 1615.0 | 1608.0 | 1611.5 | 90.59 | 9.80 |  |
|  | 05/30/97 | 0 |  | 2 |  | 12.509 | 1.3.a. | 1468.0 | 1615.0 | 1541.5 | 87.00 | 9.50 |  |
|  | 05/31/97 | 2 |  | 5 |  | 7 |  | 1395.0 | 1468.0 | 1431.5 | 90.43 | 9.40 |  |
|  | 06/01/97 | 0 |  | 3 |  | 3 |  | 1300.0 | 1395.0 | 1347.5 | 94.00 | 9.50 |  |
|  | 06/02/97 | 1 |  | 10 |  | 11 |  | 1300.0 | 1300.0 | 1300.0 | 89.45 | 9.30 |  |
|  | 06/03/97 | 0 |  | 7 |  | 7 | 1.4 | 1602.8 | 1300.0 | 1451.4 | 89.29 | 9.70 |  |
|  | 06/04/97 | 1 |  | 1 |  | 2 |  | 1610.5 | 1602.8 | 1606.7 | 92.00 | 10.20 | 10.20 |
|  | 06/05/97 | 0 |  | 7 |  | 7 |  | 1609.3 | 1610.5 | 1609.9 | 86.57 | 10.50 | 10.50 |
|  | 06/06/97 | 1.556 | 1.1 | 7 |  | 8.556 |  | 1546.6 | 1609.3 | 1577.9 | 88.75 | 10.30 | 10.30 |
|  | 06/07/97 | 0 |  | 3 |  | 3 | 1.4 | 1193.7 | 1546.6 | 1370.2 | 86.00 | 11.10 | 11.10 |
|  | 06/08/97 | 0 |  | 2 |  | 2 | 1.4 | 948.6 | 1193.7 | 1071.1 | 92.50 | 11.50 | 11.50 |
|  | 06/09/97 | 4 |  | 18.000 | 1.2 | 22.000 |  | 906.6 | 948.6 | 927.6 | 90.17 | 12.60 | 12.60 |

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| COUNT |  |  |  | FLOW (cfs) |  |  |  |  |  | SIZE |  | TURBIDITY (NTU) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap |  | South Trap | Total | Current |  | Previous |  | Average | (Length, mm) | Current |  | Used for fit |
|  | 06/10/97 | 0.667 | 1.1 | 3 | 3.667 | 923.6 |  | 906.6 |  | 915.1 | 93.67 | 12.90 |  | 12.90 |
|  | 06/11/97 | 1 |  | 6 | 7 | 916.7 |  | 923.6 |  | 920.2 | 93.86 | 12.50 |  | 12.50 |
|  | 06/12/97 | 5 |  | 1 | 6 | 913.3 |  | 916.7 |  | 915.0 | 88.00 | 12.60 |  | 12.60 |
|  | 06/13/97 | 3 |  | 2 | 5 | 914.7 |  | 913.3 |  | 914.0 | 86.80 | 12.30 |  | 12.30 |
|  | 06/14/97 | 0 |  | 3 | 3 | 908.2 |  | 914.7 |  | 911.4 | 92.33 | 11.90 |  | 11.90 |
|  | 06/15/97 | 1 |  | 1 | 2 | 905.4 |  | 908.2 |  | 906.8 | 93.50 | 12.10 |  | 12.10 |
|  | 06/16/97 | 0 |  | 6 | 6 | 907.8 |  | 905.4 |  | 906.6 | 86.33 | 11.15 | $\underline{4}$ | 11.15 |
|  | 06/17/97 | 0.222 | 1.1 | 1 | 1.222 | 903.4 |  | 907.8 |  | 905.6 | 88.00 | 10.20 |  | 10.20 |
|  | 06/18/97 | 0 |  | 3 | 3 | 895.9 |  | 903.4 |  | 899.7 | 92.00 | 10.70 |  | 10.70 |
|  | 06/19/97 | 0 |  | 4 | 4 | 897.9 |  | 895.9 |  | 896.9 | 94.50 | 11.00 |  | 11.00 |
|  | 06/20/97 | 1 |  | 2 | 3 | 912.5 |  | 897.9 |  | 905.2 | 98.00 | 10.60 |  | 10.60 |
|  | 06/21/97 | 1 |  | 3 | 4 | 921.1 |  | 912.5 |  | 916.8 | 89.25 | 10.50 |  | 10.50 |
|  | 06/22/97 | 0 |  | 4 | 4 | 915.9 |  | 921.1 |  | 918.5 | 92.00 | 9.80 |  |  |
|  | 06/23/97 | 0 |  | 2 | 2 | 917.9 |  | 915.9 |  | 916.9 | 94.50 | 10.10 |  | 10.10 |
|  | 06/24/97 | 0 |  | 1 | 1 | 924.6 |  | 917.9 |  | 921.3 | 92.00 | 9.60 |  |  |
|  | 06/25/97 | 0 |  | 0 | 0 | 916.8 |  | 924.6 |  | 920.7 | 92.00 3 | 10.30 |  | 10.30 |
|  | 06/26/97 | 0 |  | 0 | 0 | 882.2 | $\underline{2.1}$ | 916.8 | $\underline{2.2}$ | 921.0 | 92.00 - | 10.70 |  | 10.70 |
|  | 06/27/97 | 0 |  | 0 | 0 | 792.3 | 2.1 | 882.2 | 2.2 | 921.0 | 92.00 3 | 11.40 |  | 11.40 |
| 1.1 North Trap = South Trap*(North-to-South-Trap Ratio) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1.2 South Trap = North Trap/(North-to-South-Trap Ratio) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1.3.a Missing value estimate for count (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1.3.b Missing value estimate for count because both traps stopped, total of north and south trap not the value used (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 1.4 Actual North + Actual South Trap even though there was trap stoppage (adjusted value for stoppage produced smaller count). |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 2.1 Missing value flow estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 2.2 Missing value flow estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 3. Missing value length estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| 4. Missing value turbidity estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |

## Appendix A.4.c. Variables used to estimate 1998 Outmigration

|  |  |  | COUNT | FLOW (cfs) |  |  |  |  | SIZE | TURBIDITY (NTU) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South Trap | Total |  | Current Day | Previous Day | Average | (Length, mm) | Current |  | Used for fit |
| Fry | 01/29/98 | 186 | 616 | 802 |  | 1806.0 | 1365.0 | 1585.5 | 35.41 | 11.50 |  | 11.50 |
|  | 01/30/98 | 32 | 254 | 286 | 1.4 | 2623.0 | 1806.0 | 2214.5 | 35.79 | 11.60 |  | 11.60 |
|  | 01/31/98 | 193 | 2 | 703.916 | 1.3.b | 2629.0 | 2623.0 | 2626.0 | 35.22 | 12.30 | 4 | 12.30 |
|  | 02/01/98 |  |  | 678.083 | 1.3.a | 2526.0 | 2629.0 | 2577.5 | 35.53 3 | 13.00 | $\underline{4}$ | 13.00 |
|  | 02/02/98 | 37 | 1048 | 1085 |  | 2524.0 | 2526.0 | 2525.0 | 35.84 | 13.70 |  | 13.70 |
|  | 02/03/98 | 259 | 73 | 332 |  | 3854.0 | 2524.0 | 3189.0 | 37.65 | 20.30 |  | 20.30 |
|  | 02/04/98 |  |  | 643.310 | 1.3.a | 3767.0 | 3854.0 | 3810.5 | 37.25 3 | 25.06 | 4 | 25.06 |
|  | 02/05/98 |  |  | 693.067 | 1.3.a | 5497.0 | 3767.0 | 4632.0 | 36.85 - | 29.82 | 4 | 29.82 |
|  | 02/06/98 |  |  | 759.130 | 1.3.a | 4915.0 | 5497.0 | 5206.0 | 36.45 - | 34.58 | 4 | 34.58 |
|  | 02/07/98 |  |  | 850.620 | 1.3.a | 4333.0 | 4915.0 | 4624.0 | 36.05 3 | 39.34 | $\underline{4}$ | 39.34 |
|  | 02/08/98 | 138 | 1042 | 1180 |  | 5434.0 | 4333.0 | 4883.5 | 35.65 | 44.10 |  | 44.10 |
|  | 02/09/98 |  |  | 1091.088 | 1.3.a | 5460.0 | 5434.0 | 5447.0 | 35.74 3 | 38.60 | 4 | 38.60 |
|  | 02/10/98 |  |  | 1045.738 | 1.3.a | 5095.0 | 5460.0 | 5277.5 | 35.83 - | 33.10 | 4 | 33.10 |
|  | 02/11/98 |  |  | 1011.772 | 1.3.a | 5004.0 | 5095.0 | 5049.5 | 35.91 3 | 27.60 | 4 | 27.60 |
|  | 02/12/98 |  |  | 862.420 | 1.3.a | 4850.0 | 5004.0 | 4927.0 | 36.00 3 | 22.10 | 4 | 22.10 |
|  | 02/13/98 | 64 | 833 | 897 |  | 4772.0 | 4850.0 | 4811.0 | 36.09 | 16.60 |  | 16.60 |
|  | 02/14/98 | 156 | 693 | 849 | 1.4 | 4508.0 | 4772.0 | 4640.0 | 37.40 | 14.70 |  | 14.70 |
|  | 02/15/98 | 104 | 918 | 1022 |  | 4358.0 | 4508.0 | 4433.0 | 36.51 | 12.10 |  | 12.10 |
|  | 02/16/98 | 158 | 2351 | 2509 | 1.4 | 5003.0 | 4358.0 | 4680.5 | 37.32 | 9.20 |  |  |
|  | 02/17/98 | 49 | 178 | 227 |  | 4468.0 | 5003.0 | 4735.5 | 37.86 | 10.00 |  | 10.00 |
|  | 02/18/98 | 1 | 61 | 62 | 1.4 | 5064.0 | 4468.0 | 4766.0 | 39.05 | 10.80 | 4 | 10.80 |
|  | 02/19/98 | 30 | 243 | 273 |  | 4481.0 | 5064.0 | 4772.5 | 37.04 | 11.60 |  | 11.60 |
|  | 02/20/98 | 29 | 323 | 352 |  | 4530.0 | 4481.0 | 4505.5 | 37.41 | 16.50 |  | 16.50 |
|  | 02/21/98 | 50 | 343 | 393 |  | 4566.0 | 4530.0 | 4548.0 | 35.55 | 18.90 |  | 18.90 |
|  | 02/22/98 | 22 | 294 | 316 |  | 4571.0 | 4566.0 | 4568.5 | 36.59 | 10.40 |  | 10.40 |
|  | 02/23/98 | 19 | $109.507 \quad 1.2$ | 128.507 |  | 4201.0 | 4571.0 | 4386.0 | 36.33 | 14.70 |  | 14.70 |
|  | 02/24/98 | 22 | 169 | 191 |  | 3746.0 | 4201.0 | 3973.5 | 36.51 | 10.10 |  | 10.10 |
|  | 02/25/98 | 26 | 162 | 188 |  | 3746.0 | 3746.0 | 3746.0 | 36.53 | 9.50 |  |  |
|  | 02/26/98 | 32 | 127 | 159 |  | 3751.0 | 3746.0 | 3748.5 | 37.96 | 8.25 | 4 |  |
|  | 02/27/98 | 25 | 124 | 149 |  | 3700.0 | 3751.0 | 3725.5 | 38.17 | 7.00 |  |  |
|  | 02/28/98 | 23 | 139 | 162 |  | 3709.0 | 3700.0 | 3704.5 | 39.16 | 7.30 |  |  |

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## Appendix A.4.d. Variables used to estimate 1999 Outmigration

|  |  | COUNT |  |  |  | FLOW (cfs) |  |  | SIZE | TURBIDITY (NTU) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South Trap | Total |  | Current Day | Previous Day | Average | (Length, mm) | Current Day | Used for fit |
| Fry | 01/18/99 | 13.0 | 0.0 | 13.0 |  | 1192.0 | 1157.0 | 1174.5 | 34.62 | 6.80 |  |
|  | 01/19/99 | 10.0 | 6.0 | 16.0 |  | 1428.0 | 1192.0 | 1310.0 | 35.38 | 8.90 |  |
|  | 01/20/99 | 111.0 | 1.0 | 112.0 |  | 2037.0 | 1428.0 | 1732.5 | 34.14 | 16.40 | 16.40 |
|  | 01/21/99 | 1535.0 | 71.0 | 1606.0 | 1.4 | 2471.0 | 2037.0 | 2254.0 | 34.65 | 25.50 | 25.50 |
|  | 01/22/99 | 1738.0 | 111.0 | 1849.0 |  | 2888.0 | 2471.0 | 2679.5 | 34.88 | 12.00 | 12.00 |
|  | 01/23/99 | 793.0 | 19.0 | 812.0 |  | 3052.0 | 2888.0 | 2970.0 | 35.10 | 5.40 |  |
|  | 01/24/99 | 185.0 | 0.0 | 185.0 |  | 2901.0 | 3052.0 | 2976.5 | 34.64 | 9.60 |  |
|  | 01/25/99 | 938.0 | 0.0 | 938.0 |  | 2876.0 | 2901.0 | 2888.5 | 34.48 | 6.40 |  |
|  | 01/26/99 | 688.0 | 78.0 | 766.0 |  | 3276.0 | 2876.0 | 3076.0 | 35.33 | 6.30 |  |
|  | 01/27/99 | 590.0 | 156.0 | 746.0 |  | 3607.0 | 3276.0 | 3441.5 | 34.89 | 7.50 |  |
|  | 01/28/99 | 909.0 | 0.0 | 909.0 |  | 3399.0 | 3607.0 | 3503.0 | 34.24 | 3.30 |  |
|  | 01/29/99 | 578.0 | 45.0 | 623.0 |  | 2930.0 | 3399.0 | 3164.5 | 34.85 | 7.10 |  |
|  | 01/30/99 | 486.0 | 105.0 | 591.0 |  | 2308.0 | 2930.0 | 2619.0 | 34.58 | 5.40 |  |
|  | 01/31/99 | 395.0 | 226.0 | 621.0 |  | 2057.0 | 2308.0 | 2182.5 | 34.49 | 2.10 |  |
|  | 02/01/99 | 187.0 | 123.0 | 310.0 |  | 1658.0 | 2057.0 | 1857.5 | 35.65 | 7.40 |  |
|  | 02/02/99 | 569.0 | 356.0 | 925.0 |  | 1719.0 | 1658.0 | 1688.5 | 34.97 | 4.50 |  |
|  | 02/03/99 | 330.0 | 203.0 | 533.0 |  | 2104.0 | 1719.0 | 1911.5 | 35.25 | 4.00 |  |
|  | 02/04/99 | 392.0 | 190.0 | 582.0 |  | 2205.0 | 2104.0 | 2154.5 | 34.98 | 5.80 |  |
|  | 02/05/99 | 358.0 | 228.0 | 586.0 |  | 2652.0 | 2205.0 | 2428.5 | 35.43 | 4.80 |  |
|  | 02/06/99 | 798.0 | 312.0 | 1110.0 |  | 2649.0 | 2652.0 | 2650.5 | 34.02 | 6.20 |  |
|  | 02/07/99 | 535.0 | 188.0 | 723.0 |  | 2901.0 | 2649.0 | 2775.0 | 35.30 | 4.30 |  |
|  | 02/08/99 | 379.0 | 32.0 | 411.0 |  | 3110.0 | 2901.0 | 3005.5 | 34.91 | 5.90 |  |
|  | 02/09/99 | 1382.0 | 8.0 | 1390.0 |  | 3278.0 | 3110.0 | 3194.0 | 35.17 | 10.70 | 10.70 |
|  | 02/10/99 | 1921.0 | 401.0 | 2322.0 |  | 3228.0 | 3278.0 | 3253.0 | 35.26 | 23.40 | 23.40 |
|  | 02/11/99 | 1903.0 | 0.0 | 1903.0 |  | 3896.0 | 3228.0 | 3562.0 | 34.36 | 16.50 | 16.50 |
|  | 02/12/99 | 1326.0 | 906.0 | 2232.0 |  | 4209.0 | 3896.0 | 4052.5 | 35.27 | 8.00 |  |
|  | 02/13/99 | 1261.0 | 175.0 | 1436.0 |  | 4183.0 | 4209.0 | 4196.0 | 35.57 | 6.50 |  |
|  | 02/14/99 | 855.0 | 288.0 | 1143.0 |  | 4166.0 | 4183.0 | 4174.5 | 35.60 | 11.50 | 11.50 |
|  | 02/15/99 | 1520.0 | 2.0 | 1522.0 |  | 3995.0 | 4166.0 | 4080.5 | 35.00 | 2.10 |  |
|  | 02/16/99 | 42.0 | 114.0 | 156.0 |  | 3557.0 | 3995.0 | 3776.0 | 35.82 | 6.80 |  |
|  | 02/17/99 | 613.0 | 130.0 | 743.0 |  | 3863.0 | 3557.0 | 3710.0 | 35.47 | 7.20 |  |
|  | 02/18/99 | 527.0 | 40.0 | 567.0 |  | 4296.0 | 3863.0 | 4079.5 | 35.78 | 5.40 |  |
|  | 02/19/99 | 978.0 | 0.0 | 978.0 |  | 4129.0 | 4296.0 | 4212.5 | 34.30 | 5.70 |  |
|  | 02/20/99 | 761.0 | 136.0 | 897.0 |  | 4316.0 | 4129.0 | 4222.5 | 36.45 | 10.60 | 10.60 |

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|  |  | COUNT |  |  |  |  | FLOW (cfs) |  |  |  | TURBIDITY (NTU) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South T |  | Total |  | Current Day | Previous Day | Average | (Length, mm) | Current |  | Used for fit |
|  | 02/21/99 | 740.0 | 287.0 |  | 1027.0 |  | 4291.0 | 4316.0 | 4303.5 | 35.31 | 7.90 |  |  |
|  | 02/22/99 | 799.0 | 335.0 |  | 1134.0 |  | 4158.0 | 4291.0 | 4224.5 | 35.31 | 5.80 |  |  |
|  | 02/23/99 | 767.0 | 35.0 |  | 802.0 |  | 4432.0 | 4158.0 | 4295.0 | 35.72 | 3.40 |  |  |
|  | 02/24/99 | 503.0 | 277.0 |  | 780.0 |  | 4325.0 | 4432.0 | 4378.5 | 35.28 | 4.40 |  |  |
|  | 02/25/99 | 353.0 | 138.0 |  | 491.0 |  | 4261.0 | 4325.0 | 4293.0 | 34.11 | 3.30 |  |  |
|  | 02/26/99 | 242.0 | 154.0 |  | 396.0 |  | 4284.0 | 4261.0 | 4272.5 | 35.82 | 5.00 |  |  |
|  | 02/27/99 | 268.0 | 86.0 |  | 354.0 |  | 4207.0 | 4284.0 | 4245.5 | 34.88 | 2.90 |  |  |
|  | 02/28/99 | 213.0 | 116.0 |  | 329.0 |  | 3842.0 | 4207.0 | 4024.5 | 35.27 | 3.00 |  |  |
|  | 03/01/99 | 162.0 | 75.0 |  | 237.0 |  | 3535.0 | 3842.0 | 3688.5 | 35.33 | 4.00 |  |  |
|  | 03/02/99 | 208.0 | 106.0 |  | 314.0 |  | 2800.0 | 3535.0 | 3167.5 | 37.08 | 4.70 |  |  |
|  | 03/03/99 | 113.0 | 31.0 |  | 144.0 | 1.4 | 2861.0 | 2800.0 | 2830.5 | 37.36 | 7.70 |  |  |
|  | 03/04/99 | 86.0 | 19.2 | 1.2 | 105.2 |  | 2840.0 | 2861.0 | 2850.5 | 36.76 | 6.00 |  |  |
|  | 03/05/99 | 179.0 | 75.0 | 1.2 | 254.0 | 1.4 | 2641.0 | 2840.0 | 2740.5 | 39.45 | 4.10 |  |  |
|  | 03/06/99 | 158.0 | 159.0 | 1.2 | 317.0 | 1.4 | 2135.0 | 2641.0 | 2388.0 | 40.33 | 3.80 |  |  |
|  | 03/07/99 | 54.0 | 24.0 | 1.2 | 78.0 | 1.4 | 1738.0 | 2135.0 | 1936.5 | 41.41 | 3.30 |  |  |
|  | 03/08/99 | 66.0 | 22.0 | 1.2 | 88.0 | 1.4 | 1727.0 | 1738.0 | 1732.5 | 40.93 | 5.00 |  |  |
|  | 03/09/99 | 43.0 | 50.0 |  | 93.0 |  | 1736.0 | 1727.0 | 1731.5 | 41.51 | 5.70 |  |  |
|  | 03/10/99 | 31.0 | 78.0 |  | 109.0 |  | 1734.0 | 1736.0 | 1735.0 | 41.02 | 5.10 |  |  |
|  | 03/11/99 | 19.0 | 20.0 |  | 39.0 |  | 1730.0 | 1734.0 | 1732.0 | 37.38 | 3.60 |  |  |
|  | 03/12/99 | 12.0 | 27.0 |  | 39.0 | 1.4 | 1727.0 | 1730.0 | 1728.5 | 37.36 | 4.40 |  |  |
|  | 03/13/99 |  |  |  | 56.2 | 1.3.a | 1724.0 | 1727.0 | 1725.5 | 38.98 - | 2.30 |  |  |
|  | 03/14/99 | 16.0 | 23.0 |  | 39.0 | 1.4 | 1722.0 | 1724.0 | 1723.0 | 40.59 | 4.10 |  |  |
|  | 03/15/99 | 14.0 | 24.0 |  | 38.0 |  | 1729.0 | 1722.0 | 1725.5 | 42.29 | 4.00 |  |  |
| Parr | 03/16/99 | 11 | 14 |  | 38.4 | 1.3.b | 1643.0 | 1729.0 | 1686.0 | 48.84 | 3.70 |  |  |
|  | 03/17/99 |  |  |  | 38.4 | 1.3.a | 1577.0 | 1643.0 | 1610.0 | 46.95 3 | 4.15 | 4 |  |
|  | 03/18/99 | 28.0 | 30.0 |  | 58.0 | 1.4 | 1602.0 | 1577.0 | 1589.5 | 45.05 | 4.60 |  |  |
|  | 03/19/99 | 13.0 | 23.4 | 1.2 | 36.4 |  | 1595.0 | 1602.0 | 1598.5 | 46.00 | 4.20 |  |  |
|  | 03/20/99 | 15.0 | 24.0 |  | 39.0 |  | 1450.0 | 1595.0 | 1522.5 | 46.19 | 4.80 |  |  |
|  | 03/21/99 | 10.0 | 18.0 | 1.2 | 28.0 |  | 1283.0 | 1450.0 | 1366.5 | 50.29 | 4.10 |  |  |
|  | 03/22/99 | 9.0 | 12.0 |  | 21.0 |  | 1172.0 | 1283.0 | 1227.5 | 45.86 | 4.40 |  |  |
|  | 03/23/99 | 8.0 | 10.0 |  | 18.0 |  | 1175.0 | 1172.0 | 1173.5 | 51.67 | 4.30 |  |  |
|  | 03/24/99 |  |  |  | 27.7 | 1.3.a | 1119.0 | 1175.0 | 1147.0 | 52.05 3 | 4.25 | $\underline{4}$ |  |
|  | 03/25/99 | 2.0 | 12.0 |  | 14.0 |  | 1124.0 | 1119.0 | 1121.5 | 52.43 | 4.20 |  |  |
|  | 03/26/99 | 19.0 | 20.0 |  | 39.0 |  | 1124.0 | 1124.0 | 1124.0 | 51.82 | 4.80 |  |  |
|  | 03/27/99 | 20.0 | 41.0 |  | 61.0 |  | 1121.0 | 1124.0 | 1122.5 | 53.57 | 3.70 |  |  |
|  | 03/28/99 | 22.0 | 35.0 |  | 57.0 |  | 1124.0 | 1121.0 | 1122.5 | 55.71 | 4.50 |  |  |
|  | 03/29/99 | 23.0 | 30.0 |  | 53.0 |  | 1124.0 | 1124.0 | 1124.0 | 52.96 | 4.50 |  |  |
|  | 03/30/99 | 5.0 | 15.0 |  | 20.0 |  | 1146.0 | 1124.0 | 1135.0 | 57.80 | 5.20 |  |  |
|  | 03/31/99 | 4.0 | 10.0 |  | 14.0 |  | 1116.0 | 1146.0 | 1131.0 | 57.64 | 4.00 |  |  |
|  | 04/01/99 | 14.0 | 49.0 |  | 63.0 |  | 1111.0 | 1116.0 | 1113.5 | 63.50 | 3.60 |  |  |
|  | 04/02/99 | 13.0 | 64.0 |  | 77.0 |  | 1123.0 | 1111.0 | 1117.0 | 61.70 | 3.20 |  |  |
|  | 04/03/99 | 9.0 | 38.0 |  | 47.0 |  | 1146.0 | 1123.0 | 1134.5 | 64.15 | 2.10 |  |  |
|  | 04/04/99 | 27.0 | 82.0 |  | 109.0 | 1.4 | 1116.0 | 1146.0 | 1131.0 | 64.29 | 2.20 |  |  |
|  | 04/05/99 | 10.0 | 16.0 |  | 26.0 |  | 1135.0 | 1116.0 | 1125.5 | 62.54 | 4.40 |  |  |
|  | 04/06/99 | 10.0 | 25.0 |  | 35.0 |  | 1117.0 | 1135.0 | 1126.0 | 64.97 | 3.60 |  |  |
|  | 04/07/99 | 8.0 | 15.0 |  | 23.0 |  | 1111.0 | 1117.0 | 1114.0 | 60.22 | 3.50 |  |  |
|  | 04/08/99 | 6.0 | 5.0 |  | 11.0 |  | 1121.0 | 1111.0 | 1116.0 | 66.45 | 3.80 |  |  |
|  | 04/09/99 | 11.0 | 19.8 | 1.2 | 30.8 |  | 1115.0 | 1121.0 | 1118.0 | 65.00 | 4.40 |  |  |
|  | 04/10/99 | 10.0 | 11.0 |  | 21.0 |  | 1108.0 | 1115.0 | 1111.5 | 65.14 | 3.10 |  |  |
|  | 04/11/99 | 3.0 | 4.0 |  | 7.0 |  | 1124.0 | 1108.0 | 1116.0 | 66.43 | 3.20 |  |  |
|  | 04/12/99 | 10.0 | 23.0 |  | 33.0 |  | 1113.0 | 1124.0 | 1118.5 | 73.58 | 2.83 | 4 |  |
|  | 04/13/99 | 8.0 | 10.0 |  | 18.0 |  | 1129.0 | 1113.0 | 1121.0 | 74.67 | 2.45 | 4 |  |
|  | 04/14/99 | 20.0 | 18.0 |  | 38.0 |  | 1169.0 | 1129.0 | 1149.0 | 73.84 | 2.08 | 4 |  |
|  | 04/15/99 | 23.0 | 23.0 |  | 46.0 |  | 1348.0 | 1169.0 | 1258.5 | 69.49 | 1.70 |  |  |
|  | 04/16/99 | 28.0 | 50.3 | 1.2 | 78.3 |  | 1368.0 | 1348.0 | 1358.0 | 70.21 | 3.30 |  |  |
|  | 04/17/99 | 12.0 | 45.0 |  | 57.0 |  | 1366.0 | 1368.0 | 1367.0 | 74.39 | 3.80 |  |  |
|  | 04/18/99 | 21.0 | 31.0 |  | 52.0 |  | 1363.0 | 1366.0 | 1364.5 | 75.58 | 4.60 |  |  |
|  | 04/19/99 | 16.0 | 43.0 |  | 59.0 |  | 1369.0 | 1363.0 | 1366.0 | 74.43 | 5.30 |  |  |
|  | 04/20/99 | 27.0 | 11.0 |  | 38.0 |  | 1372.0 | 1369.0 | 1370.5 | 74.39 | 4.20 |  |  |
|  | 04/21/99 | 19.0 | 8.0 |  | 27.0 |  | 1377.0 | 1372.0 | 1374.5 | 77.15 | 4.30 |  |  |
|  | 04/22/99 | 15.0 | 13.0 |  | 28.0 |  | 1366.0 | 1377.0 | 1371.5 | 73.14 | 4.00 |  |  |
|  | 04/23/99 | 12.0 | 21.6 | 1.2 | 33.6 |  | 1364.0 | 1366.0 | 1365.0 | 77.92 | 3.70 |  |  |
|  | 04/24/99 | 27.0 | 22.0 |  | 49.0 |  | 1380.0 | 1364.0 | 1372.0 | 80.02 | 5.70 |  |  |

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|  |  | COUNT |  |  |  | FLOW (cfs) |  |  | SIZE | TURBIDITY (NTU) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South Trap | Total |  | Current Day | Previous Day | Average | (Length, mm) | Current | Day | Used for fit |
|  | 04/25/99 | 15.0 | 25.0 | 40.0 |  | 1382.0 | 1380.0 | 1381.0 | 78.22 | 3.00 |  |  |
|  | 04/26/99 | 10.0 | 35.0 | 45.0 |  | 1373.0 | 1382.0 | 1377.5 | 77.71 | 5.00 |  |  |
|  | 04/27/99 | 12.0 | 24.0 | 36.0 | 1.4 | 883.0 | 1373.0 | 1128.0 | 76.53 | 3.40 |  |  |
|  | 04/28/99 | 20.0 | 34.0 | 54.0 |  | 1010.0 | 883.0 | 946.5 | 76.72 | 3.80 |  |  |
|  | 04/29/99 | 5 | 30 | 52.2 | 1.3.b | 1399.0 | 1010.0 | 1204.5 | 78.69 | 4.20 |  |  |
|  | 04/30/99 | 12.0 | 23.0 | 35.0 | 1.4 | 1372.0 | 1399.0 | 1385.5 | 76.74 | 5.00 |  |  |
|  | 05/01/99 | 17.0 | 42.0 | 59.0 | 1.4 | 1364.0 | 1372.0 | 1368.0 | 78.98 | 4.60 |  |  |
|  | 05/02/99 | 24.0 | 68.0 | 92.0 |  | 1384.0 | 1364.0 | 1374.0 | 81.80 | 3.10 |  |  |
|  | 05/03/99 | 9.0 | 25.0 | 34.0 |  | 1416.0 | 1384.0 | 1400.0 | 77.85 | 2.30 |  |  |
|  | 05/04/99 | 0.0 | 15.0 | 15.0 | 1.4 | 1392.0 | 1416.0 | 1404.0 | 80.93 | 4.00 |  |  |
|  | 05/05/99 | 5 | 4 | 67.1 | 1.3.b | 1389.0 | 1392.0 | 1390.5 | 81.00 | 4.20 |  |  |
|  | 05/06/99 | 45.0 | 74.0 | 119.0 |  | 1372.0 | 1389.0 | 1380.5 | 78.29 | 3.97 | $\underline{4}$ |  |
|  | 05/07/99 | 16.0 | 39.0 | 55.0 |  | 1355.0 | 1372.0 | 1363.5 | 78.53 | 3.73 | $\underline{4}$ |  |
|  | 05/08/99 | 24.0 | 40.0 | 64.0 |  | 1348.0 | 1355.0 | 1351.5 | 79.31 | 3.50 |  |  |
|  | 05/09/99 | 23.0 | 45.0 | 68.0 |  | 1348.0 | 1348.0 | 1348.0 | 79.78 | 4.60 |  |  |
| Smolt | 05/10/99 | 16.0 | 39.0 | 55.0 |  | 1352.0 | 1348.0 | 1350.0 | 80.02 | 4.40 |  |  |
|  | 05/11/99 | 5.0 | 18.0 | 23.0 | 1.4 | 1345.0 | 1352.0 | 1348.5 | 83.05 | 4.90 |  |  |
|  | 05/12/99 | 15.0 | 47.0 | 62.0 |  | 1339.0 | 1345.0 | 1342.0 | 81.48 | 3.00 |  |  |
|  | 05/13/99 | 10.0 | 37.0 | 47.0 |  | 1344.0 | 1339.0 | 1341.5 | 78.83 | 6.00 |  |  |
|  | 05/14/99 | 24.0 | 48.0 | 72.0 |  | 1349.0 | 1344.0 | 1346.5 | 81.53 | 6.00 |  |  |
|  | 05/15/99 | 25.0 | 15.0 | 40.0 |  | 1347.0 | 1349.0 | 1348.0 | 80.92 | 2.30 |  |  |
|  | 05/16/99 | 19.0 | 31.0 | 50.0 |  | 1342.0 | 1347.0 | 1344.5 | 80.74 | 5.00 |  |  |
|  | 05/17/99 | 15.0 | 34.0 | 49.0 |  | 1341.0 | 1342.0 | 1341.5 | 81.78 | 5.00 |  |  |
|  | 05/18/99 | 34.0 | 61.0 | 95.0 |  | 1339.0 | 1341.0 | 1340.0 | 82.99 | 5.80 |  |  |
|  | 05/19/99 | 18.0 | 20.0 | 38.0 |  | 1412.0 | 1339.0 | 1375.5 | 83.16 | 5.00 |  |  |
|  | 05/20/99 | 11.0 | 29.0 | 40.0 |  | 1534.0 | 1412.0 | 1473.0 | 82.83 | 4.70 |  |  |
|  | 05/21/99 | 42.0 | 84.0 | 126.0 |  | 1533.0 | 1534.0 | 1533.5 | 82.74 | 5.00 |  |  |
|  | 05/22/99 | 37.0 | 84.7 1.2 | 121.7 |  | 1523.0 | 1533.0 | 1528.0 | 83.88 | 4.00 |  |  |
|  | 05/23/99 | 15.0 | 60.0 | 75.0 |  | 1527.0 | 1523.0 | 1525.0 | 84.54 | 4.10 |  |  |
|  | 05/24/99 | 11.0 | 54.0 | 65.0 |  | 1525.0 | 1527.0 | 1526.0 | 82.85 | 4.20 | 4 |  |
|  | 05/25/99 | 13.0 | 0.0 | 13.0 | 1.4 | 1532.0 | 1525.0 | 1528.5 | 83.38 | 4.30 |  |  |
|  | 05/26/99 | 18.0 | 100.0 | 118.0 |  | 1521.0 | 1532.0 | 1526.5 | 85.01 | 0.00 |  |  |
|  | 05/27/99 | 19.0 | 54.0 | 73.0 |  | 1520.0 | 1521.0 | 1520.5 | 84.93 | 5.50 |  |  |
|  | 05/28/99 | 5.0 | 22.0 | 27.0 |  | 1371.0 | 1520.0 | 1445.5 | 86.31 | 5.60 |  |  |
|  | 05/29/99 |  |  | 54.7 | 1.3.a | 1124.0 | 1371.0 | 1247.5 | 85.80 - | 5.65 | 4 |  |
|  | 05/30/99 |  |  | 52.6 | 1.3.a | 1122.0 | 1124.0 | 1123.0 | 85.30 - | 5.70 | $\underline{4}$ |  |
|  | 05/31/99 |  |  | 49.3 | 1.3.a | 1114.0 | 1122.0 | 1118.0 | 84.79 了 | 5.75 | $\underline{4}$ |  |
|  | 06/01/99 | 32.0 | 49.0 | 81.0 |  | 1229.0 | 1114.0 | 1171.5 | 84.28 | 5.80 |  |  |
|  | 06/02/99 | 14.0 | 13.0 | 27.0 |  | 1365.0 | 1229.0 | 1297.0 | 87.07 | 6.40 |  |  |
|  | 06/03/99 | 22.0 | 37.0 | 59.0 |  | 1369.0 | 1365.0 | 1367.0 | 86.64 | 6.30 |  |  |
|  | 06/04/99 | 18.0 | 37.0 | 55.0 |  | 1360.0 | 1369.0 | 1364.5 | 84.33 | 4.50 |  |  |
|  | 06/05/99 | 5.0 | 18.0 | 23.0 |  | 1356.0 | 1360.0 | 1358.0 | 84.96 | 1.40 |  |  |
|  | 06/06/99 | 4.0 | $9.2 \quad 1.2$ | 13.2 |  | 1362.0 | 1356.0 | 1359.0 | 85.00 | 5.40 |  |  |
|  | 06/07/99 | 5.0 | 26.0 | 31.0 |  | 1433.0 | 1362.0 | 1397.5 | 87.29 | 4.40 |  |  |
|  | 06/08/99 | 7.0 | 23.0 | 30.0 |  | 1516.0 | 1433.0 | 1474.5 | 85.43 | 4.80 |  |  |
|  | 06/09/99 | 17.0 | 17.0 | 34.0 |  | 1522.0 | 1516.0 | 1519.0 | 88.94 | 5.10 |  |  |
|  | 06/10/99 | 6.0 | 12.0 | 18.0 |  | 1518.0 | 1522.0 | 1520.0 | 89.17 | 2.00 |  |  |
|  | 06/11/99 | 4.0 | 25.0 | 29.0 |  | 1525.0 | 1518.0 | 1521.5 | 88.14 | 5.20 |  |  |
|  | 06/12/99 |  |  | 23.7 | 1.3.a | 1521.0 | 1525.0 | 1523.0 | 88.09 - | 5.17 | $\underline{4}$ |  |
|  | 06/13/99 |  |  | 19.5 | 1.3.a | 1522.0 | 1521.0 | 1521.5 | 88.05 - | 5.13 | $\underline{4}$ |  |
|  | 06/14/99 | 2.0 | $4.6 \quad 1.2$ | 6.6 |  | 1521.0 | 1522.0 | 1521.5 | 88.00 | 5.10 |  |  |
|  | 06/15/99 | 2.0 | 4.6 1.2 | 6.6 |  | 1527.0 | 1521.0 | 1524.0 | 92.00 | 5.20 |  |  |
|  | 06/16/99 | 3.0 | 9.0 | 12.0 |  | 1535.0 | 1527.0 | 1531.0 | 92.08 | 4.00 |  |  |
|  | 06/17/99 | 8.0 | 7.0 | 15.0 |  | 1531.0 | 1535.0 | 1533.0 | 91.86 | 4.40 |  |  |
|  | 06/18/99 | 3.0 | 4.0 | 7.0 |  | 1528.0 | 1531.0 | 1529.5 | 86.86 | 7.30 |  |  |
|  | 06/19/99 |  |  | 7.0 | 1.3.a | 1529.0 | 1528.0 | 1528.5 | 89.24 了 | 6.53 | $\underline{4}$ |  |
|  | 06/20/99 |  |  | 6.0 | 1.3.a | 1535.0 | 1529.0 | 1532.0 | 91.62 3 | 5.77 | $\underline{4}$ |  |
|  | 06/21/99 | 0.0 | 2.0 | 2.0 |  | 1525.0 | 1535.0 | 1530.0 | 94.00 | 5.00 |  |  |
|  | 06/22/99 | $3.9 \quad 1.1$ | 9.0 | 12.9 |  | 1530.0 | 1525.0 | 1527.5 | 90.44 | 4.60 |  |  |
|  | 06/23/99 | 2.0 | 4.0 | 6.0 |  | 1386.0 | 1530.0 | 1458.0 | 93.00 | 3.40 |  |  |
|  | 06/24/99 | 2.0 | 4.0 | 6.0 |  | 1130.0 | 1386.0 | 1258.0 | 89.60 | 4.80 |  |  |
|  | 06/25/99 | 2.0 | 3.0 | 5.0 |  | 992.0 | 1130.0 | 1061.0 | 93.40 | 4.00 |  |  |
|  | 06/26/99 |  |  | 4.0 | 1.3.a | 994.0 | 992.0 | 993.0 | 94.27 3 | 4.63 | 4 |  |

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|  |  | COUNT |  |  |  | FLOW (cfs) |  |  | SIZE | TURBIDITY (NTU) |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| COHORT | DATE | North Trap | South Trap | Tot |  | Current Day | Previous Day | Average | (Length, mm) | Current Day | Used for fit |
|  | 06/27/99 |  |  | 3.6 | 1.3.a | 992.0 | 994.0 | 993.0 | 95.13 了 | 5.27 4 |  |
|  | 06/28/99 | $0.4 \quad 1.1$ | 1.0 | 1.4 |  | 953.0 | 992.0 | 972.5 | 96.00 | 5.90 |  |
|  | 06/29/99 | 0.0 | 1.0 | 1.0 |  | 846.0 | 953.0 | 899.5 | 103.00 | 4.60 |  |
|  | 06/30/99 | 2.0 | 2.0 | 4.0 |  | 841.0 | 846.0 | 843.5 | 92.75 | 4.60 |  |
| 1.1 North Trap = South Trap*(North-to-South-Trap Ratio) |  |  |  |  |  |  |  |  |  |  |  |
| 1.2 South Trap = North Trap/(North-to-South-Trap Ratio) |  |  |  |  |  |  |  |  |  |  |  |
| 1.3.a Missing value estimate for count (see text) |  |  |  |  |  |  |  |  |  |  |  |
| 1.3.b Missing value estimate for count because both traps stopped, total of north and south trap not the value used (see text) |  |  |  |  |  |  |  |  |  |  |  |
| 1.4 Actual North + Actual South Trap even though there was trap stoppage (adjusted value for stoppage produced smaller count). |  |  |  |  |  |  |  |  |  |  |  |
| 2.1 Missing value flow estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |
| 2.2 Missing value flow estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |
| 3. Missing value length estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |
| 4. Missing value turbidity estimate for predictor variable (see text) |  |  |  |  |  |  |  |  |  |  |  |

Appendix A.5.a. 1996 Outmigration index estimates based on efficiency predictor that included turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 02/06/96 | 89.0 | 0.0998 | 892 | 233 | 892 | 233 | Fry | 892 | 233 |
| 02/07/96 | 99.2 | 0.1129 | 879 | 322 | 1,771 | 510 |  | 1,771 | 510 |
| 02/08/96 | 49.5 | 0.1154 | 429 | 379 | 2,200 | 711 |  | 2,200 | 711 |
| 02/09/96 | 13.0 | 0.1159 | 112 | 209 | 2,312 | 764 |  | 2,312 | 764 |
| 02/10/96 | 2.0 | 0.2116 | 10 | 33 | 2,321 | 766 |  | 2,321 | 766 |
| 02/11/96 | 0.0 | 0.2252 | 0 | 13 | 2,321 | 766 |  | 2,321 | 766 |
| 02/12/96 | 6.0 | 0.1890 | 32 | 17 | 2,353 | 770 |  | 2,353 | 770 |
| 02/13/96 | 2.0 | 0.1915 | 10 | 72 | 2,364 | 774 |  | 2,364 | 774 |
| 02/14/96 | 28.0 | 0.1754 | 160 | 110 | 2,523 | 793 |  | 2,523 | 793 |
| 02/15/96 | 39.0 | 0.1374 | 284 | 75 | 2,807 | 813 |  | 2,807 | 813 |
| 02/16/96 | 21.3 | 0.1238 | 172 | 98 | 2,979 | 826 |  | 2,979 | 826 |
| 02/17/96 | 44.0 | 0.0929 | 474 | 202 | 3,453 | 878 |  | 3,453 | 878 |
| 02/18/96 | 57.0 | 0.0410 | 1,391 | 337 | 4,844 | 1,083 |  | 4,844 | 1,083 |
| 02/19/96 | 52.0 | 0.0355 | 1,467 | 433 | 6,311 | 1,350 |  | 6,311 | 1,350 |
| 02/20/96 | 37.0 | 0.0468 | 790 | 221 | 7,101 | 1,420 |  | 7,101 | 1,420 |
| 02/21/96 | 48.4 | 0.0134 | 3,616 | 1,276 | 10,717 | 2,286 |  | 10,717 | 2,286 |
| 02/22/96 | 43.1 | 0.0136 | 3,160 | 2,889 | 13,877 | 4,183 |  | 13,877 | 4,183 |
| 02/23/96 | 113.0 | 0.0144 | 7,864 | 4,120 | 21,741 | 7,031 |  | 21,741 | 7,031 |
| 02/24/96 | 18.0 | 0.0328 | 551 | 1,568 | 22,292 | 7,290 |  | 22,292 | 7,290 |
| 02/25/96 | 24.0 | 0.0305 | 788 | 306 | 23,079 | 7,429 |  | 23,079 | 7,429 |
| 02/26/96 | 11.0 | 0.0276 | 399 | 255 | 23,478 | 7,506 |  | 23,478 | 7,506 |
| 02/27/96 | 16.0 | 0.0254 | 630 | 218 | 24,107 | 7,628 |  | 24,107 | 7,628 |
| 02/28/96 | 11.0 | 0.0209 | 527 | 301 | 24,634 | 7,743 |  | 24,634 | 7,743 |
| 02/29/96 | 5.0 | 0.0074 | 676 | 485 | 25,310 | 7,966 |  | 25,310 | 7,966 |
| 03/01/96 | 6.0 | 0.0066 | 909 | 454 | 26,219 | 8,292 |  | 26,219 | 8,292 |
| 03/02/96 | 7.7 | 0.0066 | 1,177 | 543 | 27,396 | 8,717 |  | 27,396 | 8,717 |
| 03/03/96 | 6.8 | 0.0066 | 1,031 | 502 | 28,427 | 9,094 |  | 28,427 | 9,094 |
| 03/04/96 | 4.8 | 0.0066 | 732 | 424 | 29,159 | 9,364 |  | 29,159 | 9,364 |
| 03/05/96 | 2.8 | 0.0061 | 448 | 404 | 29,607 | 9,536 |  | 29,607 | 9,536 |
| 03/06/96 | 0.0 | 0.0060 | 0 | 305 | 29,607 | 9,541 |  | 29,607 | 9,541 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/07/96 | 4.0 | 0.0062 | 641 | 434 | 30,248 | 9,776 |  | 30,248 | 9,776 |
| 03/08/96 | 4.0 | 0.0063 | 631 | 368 | 30,879 | 10,009 |  | 30,879 | 10,009 |
| 03/09/96 | 1.0 | 0.0064 | 156 | 300 | 31,035 | 10,071 |  | 31,035 | 10,071 |
| 03/10/96 | 0.0 | 0.0067 | 0 | 78 | 31,035 | 10,071 |  | 31,035 | 10,071 |
| 03/11/96 | 0.0 | 0.0073 | 0 | 71 | 31,035 | 10,072 |  | 31,035 | 10,072 |
| 03/12/96 | 1.0 | 0.0081 | 124 | 83 | 31,159 | 10,116 |  | 31,159 | 10,116 |
| 03/13/96 | 0.0 | 0.0176 | , | 31 | 31,159 | 10,116 |  | 31,159 | 10,116 |
| 03/14/96 | 1.0 | 0.0082 | 122 | 80 | 31,281 | 10,156 |  | 31,281 | 10,156 |
| 03/15/96 | 0.0 | 0.0080 | 0 | 67 | 31,281 | 10,156 |  | 31,281 | 10,156 |
| 03/16/96 | 1.0 | 0.0076 | 132 | 86 | 31,413 | 10,198 |  | 31,413 | 10,198 |
| 03/17/96 | 0.0 | 0.0081 | 0 | 114 | 31,413 | 10,198 |  | 31,413 | 10,198 |
| 03/18/96 | 2.0 | 0.0087 | 231 | 153 | 31,644 | 10,279 |  | 31,644 | 10,279 |
| 03/19/96 | 0.0 | 0.0083 | 0 | 110 | 31,644 | 10,279 |  | 31,644 | 10,279 |
| 03/20/96 | 1.0 | 0.0081 | 123 | 81 | 31,767 | 10,320 |  | 31,767 | 10,320 |
| 03/21/96 | 0.0 | 0.0077 | 0 | 69 | 31,767 | 10,320 | Parr | 0 | 0 |
| 03/22/96 | 0.0 | 0.0076 | 0 | 0 | 31,767 | 10,320 |  | 0 | 0 |
| 03/23/96 | 0.0 | 0.0088 | 0 | 0 | 31,767 | 10,320 |  | 0 | 0 |
| 03/24/96 | 0.0 | 0.0109 | 0 | 0 | 31,767 | 10,320 |  | 0 | 0 |
| 03/25/96 | 0.0 | 0.0125 | 0 | 180 | 31,767 | 10,322 |  | 0 | 180 |
| 03/26/96 | 4.0 | 0.0147 | 272 | 142 | 32,038 | 10,354 |  | 272 | 230 |
| 03/27/96 | 2.0 | 0.0176 | 114 | 142 | 32,152 | 10,366 |  | 385 | 273 |
| 03/28/96 | 7.0 | 0.0184 | 381 | 224 | 32,533 | 10,397 |  | 766 | 363 |
| 03/29/96 | 10.0 | 0.0199 | 502 | 187 | 33,035 | 10,429 |  | 1,268 | 426 |
| 03/30/96 | 3.0 | 0.0237 | 127 | 152 | 33,162 | 10,439 |  | 1,395 | 458 |
| 03/31/96 | 5.0 | 0.0249 | 201 | 50 | 33,362 | 10,450 |  | 1,596 | 470 |
| 04/01/96 | 3.0 | 0.0218 | 138 | 55 | 33,500 | 10,454 | Smolt | 138 | 17 |
| 04/02/96 | 3.0 | 0.0210 | 143 | 138 | 33,643 | 10,458 |  | 281 | 141 |
| 04/03/96 | 8.0 | 0.0223 | 359 | 343 | 34,002 | 10,477 |  | 640 | 375 |
| 04/04/96 | 18.0 | 0.0226 | 798 | 259 | 34,800 | 10,512 |  | 1,438 | 470 |
| 04/05/96 | 9.0 | 0.0225 | 401 | 205 | 35,201 | 10,530 |  | 1,839 | 527 |
| 04/06/96 | 14.0 | 0.0214 | 655 | 148 | 35,856 | 10,552 |  | 2,494 | 578 |
| 04/07/96 | 13.0 | 0.0241 | 540 | 304 | 36,396 | 10,586 |  | 3,034 | 676 |
| 04/08/96 | 1.0 | 0.0232 | 43 | 258 | 36,439 | 10,591 |  | 3,077 | 726 |
| 04/09/96 | 8.0 | 0.0221 | 362 | 164 | 36,802 | 10,604 |  | 3,440 | 764 |
| 04/10/96 | 4.0 | 0.0233 | 171 | 131 | 36,973 | 10,613 |  | 3,611 | 784 |
| 04/11/96 | 2.0 | 0.0222 | 90 | 162 | 37,063 | 10,618 |  | 3,701 | 806 |
| 04/12/96 | 9.0 | 0.0232 | 388 | 178 | 37,451 | 10,635 |  | 4,089 | 846 |
| 04/13/96 | 2.0 | 0.0256 | 78 | 184 | 37,529 | 10,639 |  | 4,167 | 870 |
| 04/14/96 | 0.0 | 0.0264 | 0 | 200 | 37,529 | 10,641 |  | 4,167 | 893 |
| 04/15/96 | 10.0 | 0.0257 | 389 | 209 | 37,918 | 10,649 |  | 4,556 | 936 |
| 04/16/96 | 2.0 | 0.0222 | 90 | 180 | 38,008 | 10,650 |  | 4,646 | 960 |
| 04/17/96 | 5.2 | 0.0226 | 230 | 97 | 38,238 | 10,653 |  | 4,876 | 980 |
| 04/18/96 | 6.0 | 0.0237 | 253 | 230 | 38,492 | 10,663 |  | 5,129 | 1,021 |
| 04/19/96 | 15.0 | 0.0236 | 636 | 307 | 39,128 | 10,684 |  | 5,766 | 1,104 |
| 04/20/96 | 1.0 | 0.0233 | 43 | 456 | 39,171 | 10,695 |  | 5,808 | 1,197 |
| 04/21/96 | 22.0 | 0.0232 | 950 | 764 | 40,121 | 10,737 |  | 6,758 | 1,471 |
| 04/22/96 | 36.0 | 0.0228 | 1,577 | 684 | 41,698 | 10,782 |  | 8,335 | 1,711 |
| 04/23/96 | 52.1 | 0.0232 | 2,243 | 463 | 43,940 | 10,835 |  | 10,578 | 1,913 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 04/24/96 | 38.0 | 0.0232 | 1,636 | 390 | 45,577 | 10,879 |  | 12,214 | 2,073 |
| 04/25/96 | 39.0 | 0.0226 | 1,729 | 228 | 47,306 | 10,915 |  | 13,944 | 2,233 |
| 04/26/96 | 38.0 | 0.0228 | 1,669 | 1,437 | 48,975 | 11,050 |  | 15,613 | 2,783 |
| 04/27/96 | 95.0 | 0.0229 | 4,140 | 1,707 | 53,115 | 11,289 |  | 19,753 | 3,548 |
| 04/28/96 | 109.0 | 0.0229 | 4,769 | 756 | 57,885 | 11,456 |  | 24,522 | 4,007 |
| 04/29/96 | 89.0 | 0.0226 | 3,937 | 941 | 61,822 | 11,640 |  | 28,460 | 4,470 |
| 04/30/96 | 125.0 | 0.0227 | 5,514 | 1,105 | 67,336 | 11,934 |  | 33,974 | 5,108 |
| 05/01/96 | 93.8 | 0.0226 | 4,145 | 1,077 | 71,481 | 12,190 |  | 38,119 | 5,630 |
| 05/02/96 | 84.0 | 0.0222 | 3,792 | 661 | 75,273 | 12,408 |  | 41,911 | 6,080 |
| 05/03/96 | 75.2 | 0.0470 | 1,602 | 368 | 76,875 | 12,411 |  | 43,513 | 6,279 |
| 05/04/96 | 67.0 | 0.0231 | 2,899 | 975 | 79,774 | 12,621 |  | 46,412 | 6,656 |
| 05/05/96 | 107.0 | 0.0224 | 4,780 | 1,153 | 84,554 | 12,964 |  | 51,192 | 7,289 |
| 05/06/96 | 73.0 | 0.0233 | 3,130 | 1,437 | 87,684 | 13,244 |  | 54,322 | 7,760 |
| 05/07/96 | 42.0 | 0.0230 | 1,826 | 755 | 89,510 | 13,388 |  | 56,148 | 8,000 |
| 05/08/96 | 47.0 | 0.0229 | 2,048 | 291 | 91,558 | 13,532 |  | 58,196 | 8,235 |
| 05/09/96 | 47.0 | 0.0231 | 2,036 | 287 | 93,594 | 13,679 |  | 60,232 | 8,466 |
| 05/10/96 | 52.1 | 0.0231 | 2,254 | 397 | 95,848 | 13,850 |  | 62,485 | 8,723 |
| 05/11/96 | 60.0 | 0.0229 | 2,619 | 977 | 98,467 | 14,079 |  | 65,104 | 9,074 |
| 05/12/96 | 20.0 | 0.0233 | 858 | 864 | 99,325 | 14,170 |  | 65,963 | 9,212 |
| 05/13/96 | 35.8 | 0.0231 | 1,554 | 421 | 100,879 | 14,293 |  | 67,516 | 9,399 |
| 05/14/96 | 33.9 | 0.0227 | 1,495 | 255 | 102,374 | 14,413 |  | 69,011 | 9,581 |
| 05/15/96 | 28.7 | 0.0213 | 1,349 | 407 | 103,723 | 14,531 |  | 70,361 | 9,775 |
| 05/16/96 | 19.0 | 0.0227 | 839 | 424 | 104,562 | 14,606 |  | 71,199 | 9,881 |
| 05/17/96 | 10.0 | 0.0219 | 457 | 213 | 105,018 | 14,647 |  | 71,656 | 9,941 |
| 05/18/96 | 14.0 | 0.0217 | 645 | 141 | 105,664 | 14,703 |  | 72,301 | 10,027 |
| 05/19/96 | 10.0 | 0.0209 | 478 | 226 | 106,141 | 14,747 |  | 72,779 | 10,098 |
| 05/20/96 | 19.0 | 0.0216 | 881 | 330 | 107,022 | 14,828 |  | 73,660 | 10,218 |
| 05/21/96 | 23.0 | 0.0217 | 1,061 | 331 | 108,084 | 14,925 |  | 74,722 | 10,362 |
| 05/22/96 | 10.4 | 0.0235 | 442 | 330 | 108,526 | 14,963 |  | 75,164 | 10,421 |
| 05/23/96 | 9.0 | 0.0285 | 316 | 173 | 108,842 | 14,979 |  | 75,480 | 10,458 |
| 05/24/96 | 18.0 | 0.0366 | 492 | 328 | 109,334 | 14,989 |  | 75,971 | 10,507 |
| 05/25/96 | 32.5 | 0.0402 | 807 | 436 | 110,141 | 14,994 |  | 76,779 | 10,583 |
| 05/26/96 | 52.0 | 0.0405 | 1,284 | 349 | 111,425 | 14,997 |  | 78,063 | 10,697 |
| 05/27/96 | 30.0 | 0.0403 | 744 | 468 | 112,168 | 15,006 |  | 78,806 | 10,767 |
| 05/28/96 | 15.0 | 0.0392 | 382 | 197 | 112,551 | 15,010 |  | 79,189 | 10,803 |
| 05/29/96 | 22.0 | 0.0402 | 547 | 177 | 113,098 | 15,014 |  | 79,736 | 10,849 |
| 05/30/96 | 9.0 | 0.0404 | 223 | 180 | 113,321 | 15,016 |  | 79,959 | 10,869 |
| 05/31/96 | 10.0 | 0.0393 | 254 | 40 | 113,575 | 15,018 |  | 80,213 | 10,893 |
| 06/01/96 | 10.0 | 0.0384 | 261 | 43 | 113,836 | 15,021 |  | 80,474 | 10,919 |
| 06/02/96 | 11.0 | 0.0386 | 285 | 133 | 114,121 | 15,024 |  | 80,758 | 10,947 |
| 06/03/96 | 2.0 | 0.0409 | 49 | 111 | 114,170 | 15,025 |  | 80,807 | 10,952 |
| 06/04/96 | 5.2 | 0.0382 | 136 | 69 | 114,306 | 15,026 |  | 80,944 | 10,966 |
| 06/05/96 | 7.0 | 0.0371 | 189 | 62 | 114,495 | 15,029 |  | 81,133 | 10,988 |
| 06/06/96 | 3.0 | 0.0379 | 79 | 81 | 114,574 | 15,030 |  | 81,212 | 10,997 |
| 06/07/96 | 1.0 | 0.0408 | 25 | 37 | 114,598 | 15,030 |  | 81,236 | 10,999 |
| 06/08/96 | 4.0 | 0.0326 | 123 | 50 | 114,721 | 15,034 |  | 81,359 | 11,013 |
| 06/09/96 | 2.0 | 0.0308 | 65 | 43 | 114,786 | 15,037 |  | 81,424 | 11,020 |
| 06/10/96 | 1.5 | 0.0315 | 48 | 33 | 114,833 | 15,040 |  | 81,471 | 11,025 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 06/11/96 | 0.0 | 0.0318 | 0 | 47 | 114,833 | 15,040 |  | 81,471 | 11,025 |
| 06/12/96 | 3.0 | 0.0321 | 94 | 48 | 114,927 | 15,044 |  | 81,565 | 11,033 |
| 06/13/96 | 2.0 | 0.0312 | 64 | 22 | 114,991 | 15,047 |  | 81,629 | 11,039 |
| 06/14/96 | 3.2 | 0.0336 | 97 | 50 | 115,088 | 15,050 |  | 81,726 | 11,046 |
| 06/15/96 | 0.0 | 0.0364 | 0 | 51 | 115,088 | 15,050 |  | 81,726 | 11,047 |
| 06/16/96 | 0.0 | 0.0402 | 0 | 14 | 115,088 | 15,050 |  | 81,726 | 11,047 |
| 06/17/96 | 1.0 | 0.0438 | 23 | 13 | 115,111 | 15,050 |  | 81,748 | 11,048 |
| 06/18/96 | 0.0 | 0.0444 | 0 | 13 | 115,111 | 15,050 |  | 81,748 | 11,048 |
| 06/19/96 | 0.0 | 0.0437 | 0 | 0 | 115,111 | 15,050 |  | 81,748 | 11,048 |
| 06/20/96 | 0.0 | 0.0433 | 0 | 13 | 115,111 | 15,050 |  | 81,748 | 11,048 |
| 06/21/96 | 1.0 | 0.0429 | 23 | 14 | 115,134 | 15,050 |  | 81,772 | 11,050 |
| 06/22/96 | 0.0 | 0.0398 | 0 | 14 | 115,134 | 15,050 |  | 81,772 | 11,050 |
| 06/23/96 | 1.0 | 0.0371 | 27 | 16 | 115,161 | 15,051 |  | 81,799 | 11,054 |
| 06/24/96 | 1.0 | 0.0395 | 25 | 15 | 115,186 | 15,051 |  | 81,824 | 11,057 |
| 06/25/96 | 0.0 | 0.0404 | 0 | 14 | 115,186 | 15,051 |  | 81,824 | 11,057 |
| 06/26/96 | 0.0 | 0.0415 | 0 | 14 | 115,186 | 15,051 |  | 81,824 | 11,057 |
| 06/27/96 | 1.0 | 0.0432 | 23 | 14 | 115,209 | 15,051 |  | 81,847 | 11,059 |
| 06/28/96 | 0.0 | 0.0417 | 0 | 14 | 115,209 | 15,051 |  | 81,847 | 11,059 |
| 06/29/96 | 0.0 | 0.0406 | 0 | 14 | 115,209 | 15,051 |  | 81,847 | 11,059 |
| 06/30/96 | 1.0 | 0.0395 | 25 | 15 | 115,235 | 15,051 |  | 81,872 | 11,062 |
| 07/01/96 | 1.0 | 0.0432 | 23 | 4 | 115,258 | 15,051 |  | 81,896 | 11,065 |

Appendix A.5.b. 1997 Outmigration index estimates based on efficiency predictor that included turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/19/97 | 15.0 | 0.0714 | 210 | 30 | 210 | 30 | Parr | 210 | 30 |
| 03/20/97 | 17.0 | 0.0598 | 284 | 186 | 494 | 192 |  | 494 | 192 |
| 03/21/97 | 35.0 | 0.0710 | 493 | 164 | 987 | 268 |  | 987 | 268 |
| 03/22/97 | 36.0 | 0.0641 | 562 | 133 | 1,549 | 327 |  | 1,549 | 327 |
| 03/23/97 | 48.0 | 0.0688 | 697 | 125 | 2,246 | 397 |  | 2,246 | 397 |
| 03/24/97 | 42.0 | 0.0793 | 530 | 122 | 2,775 | 457 |  | 2,775 | 457 |
| 03/25/97 | 32.0 | 0.0932 | 344 | 83 | 3,119 | 494 |  | 3,119 | 494 |
| 03/26/97 | 30.0 | 0.1309 | 229 | 58 | 3,348 | 524 |  | 3,348 | 524 |
| 03/27/97 | 22.0 | 0.1607 | 137 | 38 | 3,485 | 544 |  | 3,485 | 544 |
| 03/28/97 | 28.0 | 0.1508 | 186 | 45 | 3,671 | 572 |  | 3,671 | 572 |
| 03/29/97 | 21.0 | 0.1321 | 159 | 127 | 3,830 | 605 |  | 3,830 | 605 |
| 03/30/97 | 52.5 | 0.0588 | 892 | 303 | 4,722 | 716 |  | 4,722 | 716 |
| 03/31/97 | 30.0 | 0.1188 | 252 | 106 | 4,975 | 758 |  | 4,975 | 758 |
| 04/01/97 | 45.0 | 0.1278 | 352 | 105 | 5,327 | 813 |  | 5,327 | 813 |
| 04/02/97 | 24.3 | 0.1210 | 201 | 99 | 5,528 | 849 |  | 5,528 | 849 |
| 04/03/97 | 27.0 | 0.1173 | 230 | 48 | 5,758 | 886 |  | 5,758 | 886 |
| 04/04/97 | 28.0 | 0.0610 | 459 | 204 | 6,217 | 941 |  | 6,217 | 941 |
| 04/05/97 | 48.0 | 0.0605 | 794 | 235 | 7,011 | 1,031 |  | 7,011 | 1,037 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 04/06/97 | 51.0 | 0.0594 | 858 | 164 | 7,869 | 1,122 | Smolt | 858 | 190 |
| 04/07/97 | 39.0 | 0.0585 | 667 | 234 | 8,536 | 1,215 |  | 1,525 | 341 |
| 04/08/97 | 26.0 | 0.0581 | 447 | 188 | 8,983 | 1,279 |  | 1,972 | 427 |
| 04/09/97 | 46.4 | 0.0595 | 780 | 307 | 9,763 | 1,396 |  | 2,752 | 587 |
| 04/10/97 | 60.0 | 0.0598 | 1,003 | 205 | 10,766 | 1,519 |  | 3,755 | 712 |
| 04/11/97 | 44.0 | 0.0581 | 758 | 181 | 11,524 | 1,621 |  | 4,513 | 817 |
| 04/12/97 | 49.0 | 0.0554 | 884 | 135 | 12,408 | 1,729 |  | 5,397 | 925 |
| 04/13/97 | 45.0 | 0.0911 | 494 | 158 | 12,901 | 1,802 |  | 5,891 | 983 |
| 04/14/97 | 68.0 | 0.0730 | 931 | 590 | 13,833 | 1,983 |  | 6,822 | 1,190 |
| 04/15/97 | 126.5 | 0.0600 | 2,108 | 610 | 15,941 | 2,226 |  | 8,930 | 1,417 |
| 04/16/97 | 104.5 | 0.0545 | 1,918 | 529 | 17,858 | 2,467 |  | 10,847 | 1,658 |
| 04/17/97 | 81.0 | 0.0555 | 1,460 | 608 | 19,318 | 2,701 |  | 12,308 | 1,911 |
| 04/18/97 | 43.0 | 0.0581 | 740 | 523 | 20,058 | 2,841 |  | 13,047 | 2,066 |
| 04/19/97 | 22.0 | 0.0679 | 324 | 225 | 20,382 | 2,892 |  | 13,371 | 2,118 |
| 04/20/97 | 51.0 | 0.0627 | 813 | 289 | 21,196 | 3,020 |  | 14,185 | 2,247 |
| 04/21/97 | 28.0 | 0.0593 | 472 | 214 | 21,668 | 3,090 |  | 14,657 | 2,319 |
| 04/22/97 | 44.0 | 0.0547 | 804 | 338 | 22,471 | 3,215 |  | 15,461 | 2,448 |
| 04/23/97 | 10.0 | 0.0526 | 190 | 374 | 22,662 | 3,263 |  | 15,651 | 2,502 |
| 04/24/97 | 9.0 | 0.0514 | 175 | 185 | 22,836 | 3,290 |  | 15,826 | 2,531 |
| 04/25/97 | 26.0 | 0.0497 | 523 | 251 | 23,359 | 3,364 |  | 16,348 | 2,607 |
| 04/26/97 | 32.0 | 0.0503 | 636 | 200 | 23,995 | 3,451 |  | 16,985 | 2,696 |
| 04/27/97 | 15.0 | 0.0493 | 304 | 195 | 24,300 | 3,495 |  | 17,289 | 2,742 |
| 04/28/97 | 16.5 | 0.0226 | 731 | 165 | 25,030 | 3,538 |  | 18,019 | 2,790 |
| 04/29/97 | 21.0 | 0.0239 | 878 | 239 | 25,908 | 3,583 |  | 18,898 | 2,842 |
| 04/30/97 | 27.0 | 0.0235 | 1,150 | 545 | 27,058 | 3,683 |  | 20,047 | 2,958 |
| 05/01/97 | 3.0 | 0.0223 | 135 | 535 | 27,193 | 3,731 |  | 20,182 | 3,016 |
| 05/02/97 | 15.0 | 0.0238 | 630 | 836 | 27,822 | 3,855 |  | 20,811 | 3,164 |
| 05/03/97 | 42.0 | 0.0242 | 1,737 | 587 | 29,559 | 3,986 |  | 22,549 | 3,312 |
| 05/04/97 | 28.0 | 0.0238 | 1,175 | 433 | 30,734 | 4,083 |  | 23,724 | 3,418 |
| 05/05/97 | 47.0 | 0.0244 | 1,929 | 803 | 32,663 | 4,268 |  | 25,652 | 3,626 |
| 05/06/97 | 9.0 | 0.0234 | 385 | 814 | 33,048 | 4,374 |  | 26,037 | 3,746 |
| 05/07/97 | 32.0 | 0.0235 | 1,360 | 553 | 34,407 | 4,509 |  | 27,397 | 3,891 |
| 05/08/97 | 29.0 | 0.0240 | 1,210 | 154 | 35,618 | 4,595 |  | 28,607 | 3,983 |
| 05/09/97 | 31.0 | 0.0241 | 1,287 | 225 | 36,904 | 4,689 |  | 29,894 | 4,082 |
| 05/10/97 | 23.0 | 0.0243 | 945 | 239 | 37,850 | 4,756 |  | 30,839 | 4,155 |
| 05/11/97 | 21.0 | 0.0238 | 883 | 413 | 38,733 | 4,843 |  | 31,722 | 4,248 |
| 05/12/97 | 38.5 | 0.0247 | 1,560 | 440 | 40,292 | 4,968 |  | 33,282 | 4,382 |
| 05/13/97 | 38.5 | 0.0247 | 1,556 | 247 | 41,849 | 5,093 |  | 34,838 | 4,513 |
| 05/14/97 | 31.0 | 0.0311 | 997 | 163 | 42,846 | 5,181 |  | 35,835 | 4,599 |
| 05/15/97 | 35.6 | 0.0308 | 1,154 | 727 | 44,000 | 5,336 |  | 36,989 | 4,757 |
| 05/16/97 | 71.5 | 0.0245 | 2,921 | 1,013 | 46,921 | 5,684 |  | 39,910 | 5,123 |
| 05/17/97 | 27.5 | 0.0246 | 1,120 | 916 | 48,041 | 5,852 |  | 41,030 | 5,302 |
| 05/18/97 | 42.0 | 0.0244 | 1,723 | 734 | 49,764 | 6,048 |  | 42,753 | 5,507 |
| 05/19/97 | 62.0 | 0.0243 | 2,552 | 540 | 52,316 | 6,292 |  | 45,305 | 5,760 |
| 05/20/97 | 44.0 | 0.0246 | 1,790 | 815 | 54,106 | 6,500 |  | 47,095 | 5,977 |
| 05/21/97 | 23.0 | 0.0249 | 923 | 439 | 55,029 | 6,595 |  | 48,018 | 6,076 |
| 05/22/97 | 30.0 | 0.0242 | 1,241 | 271 | 56,270 | 6,721 |  | 49,260 | 6,204 |
| 05/23/97 | 33.8 | 0.0248 | 1,363 | 196 | 57,633 | 6,863 |  | 50,622 | 6,348 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 05/24/97 | 34.6 | 0.0286 | 1,209 | 164 | 58,842 | 6,992 |  | 51,831 | 6,476 |
| 05/25/97 | 31.0 | 0.0330 | 940 | 339 | 59,782 | 7,088 |  | 52,771 | 6,569 |
| 05/26/97 | 51.0 | 0.0335 | 1,521 | 615 | 61,303 | 7,247 |  | 54,293 | 6,726 |
| 05/27/97 | 11.0 | 0.0308 | 357 | 647 | 61,660 | 7,309 |  | 54,649 | 6,790 |
| 05/28/97 | 33.0 | 0.0265 | 1,247 | 616 | 62,907 | 7,454 |  | 55,897 | 6,938 |
| 05/29/97 | 42.0 | 0.0241 | 1,741 | 656 | 64,648 | 7,653 |  | 57,637 | 7,142 |
| 05/30/97 | 12.5 | 0.0265 | 472 | 709 | 65,120 | 7,726 |  | 58,109 | 7,218 |
| 05/31/97 | 7.0 | 0.0280 | 250 | 172 | 65,371 | 7,752 |  | 58,360 | 7,244 |
| 06/01/97 | 3.0 | 0.0288 | 104 | 138 | 65,475 | 7,765 |  | 58,464 | 7,257 |
| 06/02/97 | 11.0 | 0.0314 | 351 | 132 | 65,825 | 7,798 |  | 58,814 | 7,290 |
| 06/03/97 | 7.0 | 0.0278 | 252 | 163 | 66,077 | 7,823 |  | 59,066 | 7,315 |
| 06/04/97 | 2.0 | 0.0498 | 40 | 57 | 66,117 | 7,830 |  | 59,106 | 7,321 |
| 06/05/97 | 7.0 | 0.0535 | 131 | 67 | 66,248 | 7,847 |  | 59,237 | 7,338 |
| 06/06/97 | 8.6 | 0.0529 | 162 | 61 | 66,409 | 7,870 |  | 59,399 | 7,359 |
| 06/07/97 | 3.0 | 0.0676 | 44 | 52 | 66,454 | 7,875 |  | 59,443 | 7,365 |
| 06/08/97 | 2.0 | 0.0822 | 24 | 134 | 66,478 | 7,880 |  | 59,467 | 7,370 |
| 06/09/97 | 22.0 | 0.1008 | 218 | 118 | 66,696 | 7,912 |  | 59,685 | 7,398 |
| 06/10/97 | 3.7 | 0.1005 | 37 | 95 | 66,733 | 7,918 |  | 59,722 | 7,404 |
| 06/11/97 | 7.0 | 0.0973 | 72 | 24 | 66,805 | 7,929 |  | 59,794 | 7,414 |
| 06/12/97 | 6.0 | 0.1039 | 58 | 16 | 66,862 | 7,937 |  | 59,852 | 7,421 |
| 06/13/97 | 5.0 | 0.1030 | 49 | 18 | 66,911 | 7,943 |  | 59,900 | 7,426 |
| 06/14/97 | 3.0 | 0.0954 | 31 | 17 | 66,942 | 7,948 |  | 59,932 | 7,431 |
| 06/15/97 | 2.0 | 0.0960 | 21 | 22 | 66,963 | 7,951 |  | 59,952 | 7,434 |
| 06/16/97 | 6.0 | 0.0963 | 62 | 29 | 67,026 | 7,959 |  | 60,015 | 7,441 |
| 06/17/97 | 1.2 | 0.0890 | 14 | 27 | 67,039 | 7,961 |  | 60,029 | 7,442 |
| 06/18/97 | 3.0 | 0.0890 | 34 | 17 | 67,073 | 7,966 |  | 60,062 | 7,447 |
| 06/19/97 | 4.0 | 0.0889 | 45 | 12 | 67,118 | 7,973 |  | 60,107 | 7,453 |
| 06/20/97 | 3.0 | 0.0832 | 36 | 11 | 67,154 | 7,979 |  | 60,143 | 7,458 |
| 06/21/97 | 4.0 | 0.0890 | 45 | 11 | 67,199 | 7,985 |  | 60,188 | 7,464 |
| 06/22/97 | 4.0 | 0.0415 | 96 | 31 | 67,295 | 7,994 |  | 60,285 | 7,473 |
| 06/23/97 | 2.0 | 0.0824 | 24 | 19 | 67,320 | 7,998 |  | 60,309 | 7,476 |
| 06/24/97 | 1.0 | 0.0414 | 24 | 24 | 67,344 | 8,000 |  | 60,333 | 7,478 |
| 06/25/97 | 0.0 | 0.0853 | 0 | 7 | 67,344 | 8,000 |  | 60,333 | 7,478 |
| 06/26/97 | 0.0 | 0.0876 | 0 | 0 | 67,344 | 8,000 |  | 60,333 | 7,478 |
| 06/27/97 | 0.0 | 0.0919 | 0 | 0 | 67,344 | 8,000 |  | 60,333 | 7,478 |

Appendix A.5.c. 1998 Outmigration index estimates based on efficiency predictor that included turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 01/29/98 | 802.0 | 0.1018 | 7,879 | 3,852 | 7,879 | 3,852 | Fry | 7,879 | 3,852 |
| 01/30/98 | 286.0 | 0.1017 | 2,812 | 2,699 | 10,691 | 4,890 |  | 10,691 | 4,890 |
| 01/31/98 | 703.9 | 0.1078 | 6,531 | 2,521 | 17,222 | 6,009 |  | 17,222 | 6,009 |
| 02/01/98 | 678.1 | 0.1122 | 6,042 | 2,364 | 23,263 | 7,120 |  | 23,263 | 7,120 |
| 02/02/98 | 1085.0 | 0.1168 | 9,287 | 3,744 | 32,550 | 9,170 |  | 32,550 | 9,170 |
| 02/03/98 | 332.0 | 0.1722 | 1,928 | 2,179 | 34,478 | 9,782 |  | 34,478 | 9,782 |
| 02/04/98 | 643.3 | 0.2304 | 2,792 | 1,161 | 37,270 | 10,424 |  | 37,270 | 10,424 |
| 02/05/98 | 693.1 | 0.3012 | 2,301 | 749 | 39,572 | 10,973 |  | 39,572 | 10,973 |
| 02/06/98 | 759.1 | 0.3828 | 1,983 | 662 | 41,555 | 11,461 |  | 41,555 | 11,461 |
| 02/07/98 | 850.6 | 0.4717 | 1,803 | 712 | 43,358 | 11,903 |  | 43,358 | 11,903 |
| 02/08/98 | 1180.0 | 0.5624 | 2,098 | 663 | 45,456 | 12,379 |  | 45,456 | 12,379 |
| 02/09/98 | 1091.1 | 0.4597 | 2,374 | 748 | 47,830 | 12,982 |  | 47,830 | 12,982 |
| 02/10/98 | 1045.7 | 0.3603 | 2,903 | 929 | 50,733 | 13,764 |  | 50,733 | 13,764 |
| 02/11/98 | 1011.8 | 0.2715 | 3,726 | 1,197 | 54,459 | 14,773 |  | 54,459 | 14,773 |
| 02/12/98 | 862.4 | 0.1979 | 4,357 | 1,280 | 58,816 | 15,879 |  | 58,816 | 15,879 |
| 02/13/98 | 897.0 | 0.1404 | 6,388 | 1,550 | 65,204 | 17,292 |  | 65,204 | 17,292 |
| 02/14/98 | 849.0 | 0.1210 | 7,015 | 1,731 | 72,219 | 18,765 |  | 72,219 | 18,765 |
| 02/15/98 | 1022.0 | 0.1037 | 9,858 | 8,849 | 82,076 | 22,357 |  | 82,076 | 22,357 |
| 02/16/98 | 2509.0 | 0.0440 | 57,012 | 27,239 | 139,088 | 37,423 |  | 139,088 | 37,423 |
| 02/17/98 | 227.0 | 0.0876 | 2,592 | 15,356 | 141,680 | 40,741 |  | 141,680 | 40,741 |
| 02/18/98 | 62.0 | 0.0903 | 687 | 1,214 | 142,367 | 40,838 |  | 142,367 | 40,838 |
| 02/19/98 | 273.0 | 0.0992 | 2,752 | 1,579 | 145,119 | 41,203 |  | 145,119 | 41,203 |
| 02/20/98 | 352.0 | 0.1360 | 2,587 | 757 | 147,706 | 41,581 |  | 147,706 | 41,581 |
| 02/21/98 | 393.0 | 0.1641 | 2,395 | 661 | 150,102 | 41,959 |  | 150,102 | 41,959 |
| 02/22/98 | 316.0 | 0.0923 | 3,423 | 1,587 | 153,525 | 42,404 |  | 153,525 | 42,404 |
| 02/23/98 | 128.5 | 0.1236 | 1,040 | 788 | 154,565 | 42,561 |  | 154,565 | 42,561 |
| 02/24/98 | 191.0 | 0.0906 | 2,108 | 551 | 156,673 | 42,823 |  | 156,673 | 42,823 |
| 02/25/98 | 188.0 | 0.0448 | 4,201 | 710 | 160,873 | 43,108 |  | 160,873 | 43,108 |
| 02/26/98 | 159.0 | 0.0434 | 3,662 | 681 | 164,535 | 43,354 |  | 164,535 | 43,354 |
| 02/27/98 | 149.0 | 0.0432 | 3,447 | 494 | 167,982 | 43,587 |  | 167,982 | 43,587 |
| 02/28/98 | 162.0 | 0.0423 | 3,827 | 953 | 171,810 | 43,851 |  | 171,810 | 43,851 |
| 03/01/98 | 97.0 | 0.0421 | 2,302 | 828 | 174,112 | 44,013 |  | 174,112 | 44,013 |
| 03/02/98 | 123.0 | 0.0432 | 2,850 | 683 | 176,962 | 44,216 |  | 176,962 | 44,216 |
| 03/03/98 | 74.0 | 0.0425 | 1,741 | 659 | 178,702 | 44,342 |  | 178,702 | 44,342 |
| 03/04/98 | 81.2 | 0.0425 | 1,909 | 469 | 180,611 | 44,478 |  | 180,611 | 44,478 |
| 03/05/98 | 49.0 | 0.0382 | 1,283 | 487 | 181,894 | 44,560 |  | 181,894 | 44,560 |
| 03/06/98 | 52.0 | 0.0760 | 684 | 695 | 182,578 | 44,645 |  | 182,578 | 44,645 |
| 03/07/98 | 142.0 | 0.0412 | 3,451 | 1,231 | 186,029 | 44,897 |  | 186,029 | 44,908 |
| 03/08/98 | 124.0 | 0.0328 | 3,786 | 4,740 | 189,815 | 45,337 | Parr | 3,786 | 5,982 |
| 03/09/98 | 402.0 | 0.0330 | 12,198 | 4,945 | 202,013 | 46,227 |  | 15,984 | 7,827 |
| 03/10/98 | 394.0 | 0.0326 | 12,086 | 3,028 | 214,099 | 46,956 |  | 28,070 | 8,643 |
| 03/11/98 | 242.0 | 0.0331 | 7,317 | 2,484 | 221,416 | 47,428 |  | 35,387 | 9,243 |
| 03/12/98 | 352.0 | 0.0326 | 10,804 | 4,515 | 232,221 | 48,238 |  | 46,191 | 10,689 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/13/98 | 68.0 | 0.0304 | 2,234 | 5,280 | 234,454 | 48,641 |  | 48,425 | 12,013 |
| 03/14/98 | 77.0 | 0.0298 | 2,583 | 318 | 237,037 | 48,770 |  | 51,008 | 12,125 |
| 03/15/98 | 78.0 | 0.0333 | 2,342 | 582 | 239,379 | 48,913 |  | 53,350 | 12,247 |
| 03/16/98 | 108.0 | 0.0320 | 3,374 | 2,666 | 242,753 | 49,176 |  | 56,724 | 12,689 |
| 03/17/98 | 238.0 | 0.0308 | 7,721 | 3,626 | 250,474 | 49,722 |  | 64,445 | 13,543 |
| 03/18/98 | 20.0 | 0.0323 | 619 | 3,800 | 251,093 | 49,904 |  | 65,064 | 14,097 |
| 03/19/98 | 29.0 | 0.0304 | 953 | 406 | 252,047 | 49,957 |  | 66,017 | 14,149 |
| 03/20/98 | 43.9 | 0.0268 | 1,634 | 511 | 253,681 | 50,030 |  | 67,651 | 14,232 |
| 03/21/98 | 55.0 | 0.0325 | 1,692 | 783 | 255,372 | 50,137 |  | 69,343 | 14,340 |
| 03/22/98 | 7.3 | 0.0434 | 169 | 653 | 255,541 | 50,155 |  | 69,511 | 14,365 |
| 03/23/98 | 58.5 | 0.0249 | 2,347 | 1,157 | 257,888 | 50,258 |  | 71,859 | 14,516 |
| 03/24/98 | 54.0 | 0.0837 | 645 | 169 | 258,532 | 50,338 |  | 72,503 | 14,544 |
| 03/25/98 | 48.0 | 0.0234 | 2,053 | 11,124 | 260,586 | 51,619 |  | 74,556 | 18,383 |
| 03/26/98 | 504.0 | 0.0267 | 18,887 | 8,711 | 279,472 | 53,150 |  | 93,443 | 21,002 |
| 03/27/98 | 279.2 | 0.0243 | 11,500 | 8,674 | 290,973 | 54,288 |  | 104,944 | 23,164 |
| 03/28/98 | 85.0 | 0.1618 | 526 | 649 | 291,498 | 54,373 |  | 105,469 | 23,194 |
| 03/29/98 | 102.3 | 0.0719 | 1,423 | 414 | 292,921 | 54,534 |  | 106,892 | 23,260 |
| 03/30/98 | 123.0 | 0.0231 | 5,315 | 1,319 | 298,236 | 54,745 |  | 112,207 | 23,528 |
| 03/31/98 | 68.4 | 0.0227 | 3,014 | 1,391 | 301,250 | 54,872 |  | 115,221 | 23,706 |
| 04/01/98 | 71.0 | 0.0233 | 3,048 | 385 | 304,298 | 54,991 |  | 118,269 | 23,853 |
| 04/02/98 | 62.0 | 0.0209 | 2,968 | 1,135 | 307,266 | 55,097 |  | 121,236 | 24,020 |
| 04/03/98 | 105.0 | 0.0233 | 4,505 | 3,683 | 311,770 | 55,399 |  | 125,741 | 24,521 |
| 04/04/98 | 227.0 | 0.0580 | 3,915 | 1,858 | 315,685 | 55,828 |  | 129,656 | 24,798 |
| 04/05/98 | 302.0 | 0.0233 | 12,976 | 2,741 | 328,661 | 56,426 |  | 142,632 | 25,610 |
| 04/06/98 | 194.7 | 0.0263 | 7,389 | 2,163 | 336,050 | 56,845 |  | 150,021 | 26,109 |
| 04/07/98 | 254.0 | 0.0235 | 10,796 | 2,738 | 346,846 | 57,396 |  | 160,817 | 26,870 |
| 04/08/98 | 312.0 | 0.0231 | 13,496 | 4,187 | 360,342 | 58,162 |  | 174,312 | 27,999 |
| 04/09/98 | 133.3 | 0.0238 | 5,595 | 5,099 | 365,937 | 58,662 |  | 179,908 | 28,811 |
| 04/10/98 | 80.4 | 0.0236 | 3,404 | 1,354 | 369,341 | 58,847 |  | 183,312 | 29,062 |
| 04/11/98 | 79.0 | 0.0218 | 3,621 | 476 | 372,962 | 59,015 |  | 186,933 | 29,306 |
| 04/12/98 | 71.0 | 0.0209 | 3,391 | 1,466 | 376,353 | 59,183 |  | 190,324 | 29,574 |
| 04/13/98 | 24.0 | 0.0200 | 1,200 | 1,341 | 377,553 | 59,249 |  | 191,524 | 29,688 |
| 04/14/98 | 25.0 | 0.0167 | 1,500 | 548 | 379,052 | 59,301 |  | 193,023 | 29,804 |
| 04/15/98 | 39.0 | 0.0164 | 2,374 | 586 | 381,427 | 59,382 |  | 195,398 | 29,987 |
| 04/16/98 | 27.0 | 0.0183 | 1,479 | 657 | 382,906 | 59,443 |  | 196,877 | 30,103 |
| 04/17/98 | 16.0 | 0.0182 | 879 | 573 | 383,785 | 59,480 |  | 197,756 | 30,175 |
| 04/18/98 | 36.5 | 0.0201 | 1,821 | 1,472 | 385,606 | 59,580 |  | 199,577 | 30,343 |
| 04/19/98 | 74.0 | 0.0180 | 4,119 | 1,567 | 389,726 | 59,763 |  | 203,696 | 30,696 |
| 04/20/98 | 23.0 | 0.0198 | 1,160 | 1,510 | 390,886 | 59,835 |  | 204,856 | 30,820 |
| 04/21/98 | 21.0 | 0.0182 | 1,154 | 231 | 392,040 | 59,883 |  | 206,011 | 30,911 |
| 04/22/98 | 27.0 | 0.0158 | 1,711 | 637 | 393,751 | 59,947 |  | 207,722 | 31,058 |
| 04/23/98 | 39.0 | 0.0178 | 2,190 | 736 | 395,940 | 60,042 |  | 209,911 | 31,238 |
| 04/24/98 | 51.0 | 0.0168 | 3,026 | 689 | 398,967 | 60,166 | Smolt | 3,026 | 504 |
| 04/25/98 | 56.1 | 0.0169 | 3,312 | 649 | 402,279 | 60,305 |  | 6,338 | 1,063 |
| 04/26/98 | 42.0 | 0.0175 | 2,395 | 552 | 404,673 | 60,411 |  | 8,733 | 1,448 |
| 04/27/98 | 44.0 | 0.0177 | 2,487 | 1,095 | 407,161 | 60,532 |  | 11,220 | 2,055 |
| 04/28/98 | 75.0 | 0.0161 | 4,671 | 1,238 | 411,831 | 60,738 |  | 15,891 | 2,866 |
| 04/29/98 | 67.0 | 0.0175 | 3,830 | 601 | 415,661 | 60,919 |  | 19,721 | 3,353 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 04/30/98 | 72.0 | 0.0162 | 4,443 | 1,320 | 420,104 | 61,133 |  | 24,164 | 4,138 |
| 05/01/98 | 101.0 | 0.0168 | 6,006 | 1,599 | 426,110 | 61,440 |  | 30,169 | 5,131 |
| 05/02/98 | 57.0 | 0.0168 | 3,388 | 1,807 | 429,498 | 61,636 |  | 33,557 | 5,854 |
| 05/03/98 | 45.0 | 0.0166 | 2,718 | 689 | 432,215 | 61,777 |  | 36,275 | 6,245 |
| 05/04/98 | 39.0 | 0.0302 | 1,293 | 1,161 | 433,508 | 61,909 |  | 37,568 | 6,542 |
| 05/05/98 | 102.0 | 0.0158 | 6,453 | 2,236 | 439,961 | 62,274 |  | 44,021 | 7,729 |
| 05/06/98 | 65.0 | 0.0158 | 4,111 | 2,805 | 444,072 | 62,553 |  | 48,132 | 8,753 |
| 05/07/98 | 15.0 | 0.0341 | 440 | 981 | 444,512 | 62,606 |  | 48,572 | 8,866 |
| 05/08/98 | 0.0 | 0.0159 | 0 | 1,667 | 444,512 | 62,628 |  | 48,572 | 9,021 |
| 05/09/98 | 52.2 | 0.0156 | 3,348 | 3,062 | 447,861 | 62,884 |  | 51,920 | 9,950 |
| 05/10/98 | 95.0 | 0.0153 | 6,222 | 1,815 | 454,082 | 63,250 |  | 58,142 | 10,913 |
| 05/11/98 | 88.0 | 0.0150 | 5,880 | 1,040 | 459,963 | 63,592 |  | 64,022 | 11,768 |
| 05/12/98 | 94.0 | 0.0154 | 6,104 | 1,991 | 466,067 | 63,988 |  | 70,126 | 12,765 |
| 05/13/98 | 45.0 | 0.0142 | 3,175 | 3,113 | 469,242 | 64,256 |  | 73,301 | 13,616 |
| 05/14/98 | 133.0 | 0.0146 | 9,083 | 4,298 | 478,325 | 64,965 |  | 82,385 | 15,526 |
| 05/15/98 | 158.0 | 0.0153 | 10,355 | 1,984 | 488,680 | 65,682 |  | 92,739 | 17,064 |
| 05/16/98 | 132.0 | 0.0153 | 8,604 | 2,044 | 497,283 | 66,316 |  | 101,343 | 18,400 |
| 05/17/98 | 113.0 | 0.0161 | 7,023 | 1,730 | 504,306 | 66,852 |  | 108,366 | 19,450 |
| 05/18/98 | 89.0 | 0.0150 | 5,943 | 5,046 | 510,250 | 67,489 |  | 114,309 | 20,965 |
| 05/19/98 | 229.4 | 0.0156 | 14,670 | 5,806 | 524,919 | 68,858 |  | 128,979 | 23,719 |
| 05/20/98 | 80.0 | 0.0151 | 5,305 | 6,662 | 530,224 | 69,610 |  | 134,284 | 25,395 |
| 05/21/98 | 37.0 | 0.0151 | 2,458 | 1,469 | 532,682 | 69,830 |  | 136,742 | 25,795 |
| 05/22/98 | 59.0 | 0.0150 | 3,929 | 2,803 | 536,611 | 70,215 |  | 140,670 | 26,518 |
| 05/23/98 | 117.3 | 0.0152 | 7,722 | 2,648 | 544,333 | 70,918 |  | 148,392 | 27,753 |
| 05/24/98 | 53.0 | 0.0148 | 3,573 | 2,815 | 547,905 | 71,287 |  | 151,965 | 28,429 |
| 05/25/98 | 40.0 | 0.0155 | 2,579 | 1,078 | 550,485 | 71,521 |  | 154,544 | 28,820 |
| 05/26/98 | 71.0 | 0.0152 | 4,677 | 2,285 | 555,162 | 71,972 |  | 159,222 | 29,595 |
| 05/27/98 | 5.0 | 0.0156 | 320 | 2,088 | 555,483 | 72,031 |  | 159,542 | 29,715 |
| 05/28/98 | 41.0 | 0.0145 | 2,824 | 1,714 | 558,306 | 72,309 |  | 162,366 | 30,200 |
| 05/29/98 | 51.0 | 0.0137 | 3,716 | 835 | 562,022 | 72,662 |  | 166,081 | 30,818 |
| 05/30/98 | 39.0 | 0.0138 | 2,817 | 918 | 564,839 | 72,934 |  | 168,899 | 31,292 |
| 05/31/98 | 29.8 | 0.0130 | 2,294 | 1,362 | 567,133 | 73,171 |  | 171,192 | 31,720 |
| 06/01/98 | 6.0 | 0.0122 | 492 | 1,925 | 567,624 | 73,245 |  | 171,684 | 31,869 |
| 06/02/98 | 54.0 | 0.0144 | 3,746 | 1,770 | 571,371 | 73,624 |  | 175,430 | 32,508 |
| 06/03/98 | 29.0 | 0.0144 | 2,009 | 1,877 | 573,380 | 73,842 |  | 177,439 | 32,880 |
| 06/04/98 | 0.0 | 0.0138 | 0 | 2,730 | 573,380 | 73,892 |  | 177,439 | 32,993 |
| 06/05/98 | 76.0 | 0.0132 | 5,761 | 3,352 | 579,140 | 74,542 |  | 183,200 | 34,141 |
| 06/06/98 | 5.1 | 0.0112 | 454 | 3,281 | 579,595 | 74,663 |  | 183,654 | 34,388 |
| 06/07/98 | 17.9 | 0.0118 | 1,515 | 830 | 581,109 | 74,829 |  | 185,169 | 34,686 |
| 06/08/98 | 0.0 | 0.0124 | 0 | 2,693 | 581,109 | 74,877 |  | 185,169 | 34,790 |
| 06/09/98 | 66.0 | 0.0130 | 5,064 | 3,015 | 586,174 | 75,459 |  | 190,233 | 35,792 |
| 06/10/98 | 1.0 | 0.0124 | 81 | 2,702 | 586,255 | 75,516 |  | 190,314 | 35,908 |
| 06/11/98 | 15.0 | 0.0130 | 1,157 | 777 | 587,412 | 75,641 |  | 191,471 | 36,119 |
| 06/12/98 | 19.9 | 0.0131 | 1,521 | 479 | 588,933 | 75,803 |  | 192,993 | 36,387 |
| 06/13/98 | 25.0 | 0.0128 | 1,955 | 704 | 590,888 | 76,015 |  | 194,947 | 36,741 |
| 06/14/98 | 10.0 | 0.0126 | 795 | 993 | 591,683 | 76,107 |  | 195,742 | 36,899 |
| 06/15/98 | 0.0 | 0.0262 | 0 | 186 | 591,683 | 76,107 |  | 195,742 | 36,899 |
| 06/16/98 | 6.0 | 0.0310 | 193 | 113 | 591,876 | 76,138 |  | 195,936 | 36,935 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 06/17/98 | 1.0 | 0.0310 | 32 | 82 | 591,909 | 76,143 |  | 195,968 | 36,942 |
| 06/18/98 | 2.0 | 0.0299 | 67 | 38 | 591,976 | 76,154 |  | 196,035 | 36,955 |
| 06/19/98 | 0.0 | 0.0274 | 0 | 38 | 591,976 | 76,154 |  | 196,035 | 36,955 |
| 06/20/98 | 1.7 | 0.0251 | 70 | 44 | 592,045 | 76,165 |  | 196,105 | 36,969 |
| 06/21/98 | 1.8 | 0.0118 | 154 | 50 | 592,199 | 76,182 |  | 196,259 | 36,998 |
| 06/22/98 | 1.0 | 0.0121 | 83 | 46 | 592,282 | 76,191 |  | 196,341 | 37,014 |
| 06/23/98 | 2.0 | 0.0135 | 148 | 78 | 592,430 | 76,207 |  | 196,489 | 37,039 |
| 06/24/98 | 3.0 | 0.0102 | 293 | 161 | 592,722 | 76,241 |  | 196,782 | 37,102 |
| 06/25/98 | 0.0 | 0.0099 | 0 | 147 | 592,722 | 76,241 |  | 196,782 | 37,103 |
| 06/26/98 | 1.6 | 0.0096 | 170 | 99 | 592,892 | 76,262 |  | 196,952 | 37,141 |
| 06/27/98 | 1.4 | 0.0093 | 150 | 45 | 593,042 | 76,280 |  | 197,102 | 37,175 |
| 06/28/98 | 1.2 | 0.0090 | 137 | 43 | 593,179 | 76,297 |  | 197,238 | 37,208 |
| 06/29/98 | 1.0 | 0.0087 | 115 | 66 | 593,294 | 76,311 |  | 197,353 | 37,235 |
| 06/30/98 | 2.0 | 0.0093 | 215 | 118 | 593,508 | 76,337 |  | 197,568 | 37,285 |
| 07/01/98 | 0.0 | 0.0099 | 0 | 112 | 593,508 | 76,337 |  | 197,568 | 37,285 |
| 07/02/98 | 0.0 | 0.0106 | 0 | 106 | 593,508 | 76,337 |  | 197,568 | 37,285 |
| 07/03/98 | 2.0 | 0.0113 | 176 | 95 | 593,684 | 76,357 |  | 197,744 | 37,320 |
| 07/04/98 | 0.7 | 0.0113 | 66 | 70 | 593,750 | 76,365 |  | 197,810 | 37,333 |
| 07/05/98 | 0.5 | 0.0113 | 46 | 21 | 593,797 | 76,370 |  | 197,856 | 37,343 |
| 07/06/98 | 0.3 | 0.0113 | 28 | 24 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/07/98 | 0.0 | 0.0113 | 0 | 16 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/08/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/09/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/10/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/11/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/12/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/13/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/14/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/15/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |
| 07/16/98 | 0.0 | 0.0113 | 0 | 0 | 593,825 | 76,373 |  | 197,884 | 37,348 |

Appendix A.5.d. 1999 Outmigration index estimates based on efficiency predictor that included turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 01/18/99 | 13 | 0.0260 | 501 | 97 | 501 | 97 | Fry | 501 | 97 |
| 01/19/99 | 16 | 0.0260 | 616 | 2,158 | 1,117 | 2,162 |  | 1,117 | 2,162 |
| 01/20/99 | 112 | 0.0831 | 1,347 | 10,424 | 2,464 | 10,647 |  | 2,464 | 10,647 |
| 01/21/99 | 1,606 | 0.1517 | 10,585 | 6,789 | 13,049 | 12,721 |  | 13,049 | 12,721 |
| 01/22/99 | 1,849 | 0.0613 | 30,168 | 10,407 | 43,216 | 17,668 |  | 43,216 | 17,668 |
| 01/23/99 | 812 | 0.0260 | 31,272 | 32,353 | 74,488 | 37,236 |  | 74,488 | 37,236 |
| 01/24/99 | 185 | 0.0260 | 7,125 | 15,466 | 81,613 | 40,458 |  | 81,613 | 40,458 |
| 01/25/99 | 938 | 0.0260 | 36,125 | 15,575 | 117,737 | 44,062 |  | 117,737 | 44,062 |
| 01/26/99 | 766 | 0.0260 | 29,501 | 5,079 | 147,238 | 45,178 |  | 147,238 | 45,178 |
| 01/27/99 | 746 | 0.0260 | 28,730 | 4,534 | 175,968 | 46,388 |  | 175,968 | 46,388 |
| 01/28/99 | 909 | 0.0260 | 35,008 | 6,596 | 210,976 | 48,241 |  | 210,976 | 48,241 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 01/29/99 | 623 | 0.0260 | 23,993 | 7,157 | 234,969 | 49,870 |  | 234,969 | 49,870 |
| 01/30/99 | 591 | 0.0260 | 22,761 | 2,469 | 257,730 | 51,068 |  | 257,730 | 51,068 |
| 01/31/99 | 621 | 0.0260 | 23,916 | 7,028 | 281,646 | 52,819 |  | 281,646 | 52,819 |
| 02/01/99 | 310 | 0.0260 | 11,939 | 11,844 | 293,585 | 54,795 |  | 293,585 | 54,795 |
| 02/02/99 | 925 | 0.0260 | 35,624 | 12,490 | 329,209 | 58,169 |  | 329,209 | 58,169 |
| 02/03/99 | 533 | 0.0260 | 20,527 | 8,456 | 349,736 | 60,006 |  | 349,736 | 60,006 |
| 02/04/99 | 582 | 0.0260 | 22,414 | 2,595 | 372,150 | 61,452 |  | 372,150 | 61,452 |
| 02/05/99 | 586 | 0.0260 | 22,568 | 11,868 | 394,718 | 64,018 |  | 394,718 | 64,018 |
| 02/06/99 | 1,110 | 0.0260 | 42,749 | 11,322 | 437,467 | 67,754 |  | 437,467 | 67,754 |
| 02/07/99 | 723 | 0.0260 | 27,844 | 13,723 | 465,311 | 71,006 |  | 465,311 | 71,006 |
| 02/08/99 | 411 | 0.0260 | 15,829 | 19,227 | 481,140 | 74,636 |  | 481,140 | 74,636 |
| 02/09/99 | 1,390 | 0.0559 | 24,848 | 17,376 | 505,988 | 78,634 |  | 505,988 | 78,634 |
| 02/10/99 | 2,322 | 0.1326 | 17,507 | 6,303 | 523,495 | 80,718 |  | 523,495 | 80,718 |
| 02/11/99 | 1,903 | 0.0837 | 22,733 | 5,975 | 546,227 | 83,427 |  | 546,227 | 83,427 |
| 02/12/99 | 2,232 | 0.0260 | 85,960 | 17,747 | 632,187 | 91,007 |  | 632,187 | 91,007 |
| 02/13/99 | 1,436 | 0.0260 | 55,304 | 22,341 | 687,491 | 97,635 |  | 687,491 | 97,635 |
| 02/14/99 | 1,143 | 0.0592 | 19,314 | 4,859 | 706,805 | 99,818 |  | 706,805 | 99,818 |
| 02/15/99 | 1,522 | 0.0260 | 58,616 | 27,693 | 765,421 | 107,800 |  | 765,421 | 107,800 |
| 02/16/99 | 156 | 0.0260 | 6,008 | 26,254 | 771,429 | 111,396 |  | 771,429 | 111,396 |
| 02/17/99 | 743 | 0.0260 | 28,615 | 11,917 | 800,044 | 114,131 |  | 800,044 | 114,131 |
| 02/18/99 | 567 | 0.0260 | 21,837 | 8,219 | 821,880 | 116,058 |  | 821,880 | 116,058 |
| 02/19/99 | 978 | 0.0260 | 37,665 | 9,216 | 859,545 | 119,250 |  | 859,545 | 119,250 |
| 02/20/99 | 897 | 0.0555 | 16,149 | 3,061 | 875,694 | 120,977 |  | 875,694 | 120,977 |
| 02/21/99 | 1,027 | 0.0260 | 39,552 | 6,136 | 915,246 | 124,171 |  | 915,246 | 124,171 |
| 02/22/99 | 1,134 | 0.0260 | 43,673 | 7,927 | 958,919 | 127,836 |  | 958,919 | 127,836 |
| 02/23/99 | 802 | 0.0260 | 30,887 | 8,250 | 989,806 | 130,568 |  | 989,806 | 130,568 |
| 02/24/99 | 780 | 0.0260 | 30,040 | 7,348 | 1,019,845 | 133,201 |  | 1,019,845 | 133,201 |
| 02/25/99 | 491 | 0.0260 | 18,910 | 7,910 | 1,038,755 | 134,984 |  | 1,038,755 | 134,984 |
| 02/26/99 | 396 | 0.0260 | 15,251 | 3,123 | 1,054,006 | 136,278 |  | 1,054,006 | 136,278 |
| 02/27/99 | 354 | 0.0260 | 13,633 | 1,923 | 1,067,639 | 137,423 |  | 1,067,639 | 137,423 |
| 02/28/99 | 329 | 0.0260 | 12,671 | 2,704 | 1,080,310 | 138,507 |  | 1,080,310 | 138,507 |
| 03/01/99 | 237 | 0.0260 | 9,127 | 2,116 | 1,089,437 | 139,288 |  | 1,089,437 | 139,288 |
| 03/02/99 | 314 | 0.0260 | 12,093 | 3,496 | 1,101,530 | 140,347 |  | 1,101,530 | 140,347 |
| 03/03/99 | 144 | 0.0260 | 5,546 | 4,292 | 1,107,076 | 140,881 |  | 1,107,076 | 140,881 |
| 03/04/99 | 105 | 0.0260 | 4,053 | 2,986 | 1,111,129 | 141,256 |  | 1,111,129 | 141,256 |
| 03/05/99 | 254 | 0.0260 | 9,782 | 4,288 | 1,120,911 | 142,148 |  | 1,120,911 | 142,148 |
| 03/06/99 | 317 | 0.0260 | 12,208 | 4,912 | 1,133,119 | 143,266 |  | 1,133,119 | 143,266 |
| 03/07/99 | 78 | 0.0260 | 3,004 | 5,188 | 1,136,123 | 143,616 |  | 1,136,123 | 143,616 |
| 03/08/99 | 88 | 0.0260 | 3,389 | 459 | 1,139,512 | 143,905 |  | 1,139,512 | 143,905 |
| 03/09/99 | 93 | 0.0260 | 3,582 | 562 | 1,143,094 | 144,212 |  | 1,143,094 | 144,212 |
| 03/10/99 | 109 | 0.0260 | 4,198 | 1,471 | 1,147,292 | 144,578 |  | 1,147,292 | 144,578 |
| 03/11/99 | 39 | 0.0260 | 1,502 | 1,556 | 1,148,794 | 144,714 |  | 1,148,794 | 144,714 |
| 03/12/99 | 39 | 0.0260 | 1,502 | 411 | 1,150,296 | 144,843 |  | 1,150,296 | 144,843 |
| 03/13/99 | 56 | 0.0260 | 2,163 | 442 | 1,152,459 | 145,029 |  | 1,152,459 | 145,029 |
| 03/14/99 | 39 | 0.0260 | 1,502 | 421 | 1,153,961 | 145,159 |  | 1,153,961 | 145,159 |
| 03/15/99 | 38 | 0.0260 | 1,464 | 154 | 1,155,424 | 145,284 |  | 1,155,424 | 145,284 |
| 03/16/99 | 38 | 0.0260 | 1,478 | 154 | 1,156,902 | 145,411 | Parr | 1,478 | 154 |
| 03/17/99 | 38 | 0.0260 | 1,480 | 460 | 1,158,382 | 145,538 |  | 2,958 | 532 |
| 03/18/99 | 58 | 0.0260 | 2,234 | 513 | 1,160,616 | 145,731 |  | 5,191 | 830 |
| 03/19/99 | 36 | 0.0260 | 1,401 | 475 | 1,162,017 | 145,852 |  | 6,592 | 1,036 |
| 03/20/99 | 39 | 0.0260 | 1,502 | 270 | 1,163,519 | 145,981 |  | 8,094 | 1,167 |
| 03/21/99 | 28 | 0.0260 | 1,078 | 365 | 1,164,596 | 146,074 |  | 9,172 | 1,298 |
| 03/22/99 | 21 | 0.0260 | 809 | 214 | 1,165,405 | 146,144 |  | 9,981 | 1,375 |
| 03/23/99 | 18 | 0.0260 | 693 | 203 | 1,166,098 | 146,204 |  | 10,674 | 1,443 |
| 03/24/99 | 28 | 0.0260 | 1,066 | 291 | 1,167,164 | 146,296 |  | 11,739 | 1,554 |
| 03/25/99 | 14 | 0.0260 | 539 | 483 | 1,167,703 | 146,343 |  | 12,279 | 1,669 |
| 03/26/99 | 39 | 0.0260 | 1,502 | 914 | 1,169,205 | 146,475 |  | 13,781 | 2,005 |
| 03/27/99 | 61 | 0.0260 | 2,349 | 511 | 1,171,554 | 146,678 |  | 16,130 | 2,233 |
| 03/28/99 | 57 | 0.0260 | 2,195 | 275 | 1,173,749 | 146,867 |  | 18,325 | 2,414 |
| 03/29/99 | 53 | 0.0260 | 2,041 | 806 | 1,175,790 | 147,045 |  | 20,366 | 2,700 |
| 03/30/99 | 20 | 0.0260 | 770 | 808 | 1,176,561 | 147,114 |  | 21,136 | 2,879 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/31/99 | 14 | 0.0260 | 539 | 1,025 | 1,177,100 | 147,164 |  | 21,676 | 3,096 |
| 04/01/99 | 63 | 0.0260 | 2,426 | 1,292 | 1,179,526 | 147,379 |  | 24,102 | 3,521 |
| 04/02/99 | 77 | 0.0260 | 2,966 | 653 | 1,182,492 | 147,636 |  | 27,067 | 3,791 |
| 04/03/99 | 47 | 0.0260 | 1,810 | 1,203 | 1,184,302 | 147,797 |  | 28,877 | 4,109 |
| 04/04/99 | 109 | 0.0260 | 4,198 | 1,710 | 1,188,500 | 148,169 |  | 33,075 | 4,737 |
| 04/05/99 | 26 | 0.0260 | 1,001 | 1,748 | 1,189,501 | 148,266 |  | 34,077 | 5,120 |
| 04/06/99 | 35 | 0.0260 | 1,348 | 277 | 1,190,849 | 148,383 |  | 35,424 | 5,224 |
| 04/07/99 | 23 | 0.0260 | 886 | 469 | 1,191,735 | 148,460 |  | 36,310 | 5,309 |
| 04/08/99 | 11 | 0.0260 | 424 | 384 | 1,192,158 | 148,498 |  | 36,734 | 5,354 |
| 04/09/99 | 31 | 0.0260 | 1,185 | 398 | 1,193,344 | 148,601 |  | 37,919 | 5,457 |
| 04/10/99 | 21 | 0.0260 | 809 | 466 | 1,194,152 | 148,671 |  | 38,728 | 5,537 |
| 04/11/99 | 7 | 0.0260 | 270 | 499 | 1,194,422 | 148,696 |  | 38,998 | 5,580 |
| 04/12/99 | 33 | 0.0260 | 1,271 | 517 | 1,195,693 | 148,807 |  | 40,268 | 5,699 |
| 04/13/99 | 18 | 0.0260 | 693 | 405 | 1,196,386 | 148,867 |  | 40,962 | 5,766 |
| 04/14/99 | 38 | 0.0260 | 1,464 | 573 | 1,197,849 | 148,995 |  | 42,425 | 5,906 |
| 04/15/99 | 46 | 0.0260 | 1,772 | 839 | 1,199,621 | 149,151 |  | 44,197 | 6,100 |
| 04/16/99 | 78 | 0.0260 | 3,017 | 704 | 1,202,638 | 149,414 |  | 47,214 | 6,372 |
| 04/17/99 | 57 | 0.0260 | 2,195 | 583 | 1,204,833 | 149,606 |  | 49,409 | 6,572 |
| 04/18/99 | 52 | 0.0260 | 2,003 | 250 | 1,206,836 | 149,780 |  | 51,412 | 6,738 |
| 04/19/99 | 59 | 0.0260 | 2,272 | 473 | 1,209,108 | 149,978 |  | 53,684 | 6,940 |
| 04/20/99 | 38 | 0.0260 | 1,464 | 641 | 1,210,572 | 150,107 |  | 55,147 | 7,091 |
| 04/21/99 | 27 | 0.0260 | 1,040 | 257 | 1,211,612 | 150,197 |  | 56,187 | 7,183 |
| 04/22/99 | 28 | 0.0260 | 1,078 | 176 | 1,212,690 | 150,291 |  | 57,266 | 7,276 |
| 04/23/99 | 34 | 0.0260 | 1,293 | 438 | 1,213,983 | 150,404 |  | 58,559 | 7,399 |
| 04/24/99 | 49 | 0.0260 | 1,887 | 356 | 1,215,870 | 150,569 |  | 60,446 | 7,567 |
| 04/25/99 | 40 | 0.0260 | 1,541 | 236 | 1,217,411 | 150,703 |  | 61,986 | 7,703 |
| 04/26/99 | 45 | 0.0260 | 1,733 | 250 | 1,219,144 | 150,854 |  | 63,719 | 7,857 |
| 04/27/99 | 36 | 0.0260 | 1,386 | 374 | 1,220,530 | 150,976 |  | 65,106 | 7,987 |
| 04/28/99 | 54 | 0.0260 | 2,080 | 437 | 1,222,610 | 151,158 |  | 67,185 | 8,181 |
| 04/29/99 | 52 | 0.0260 | 2,009 | 453 | 1,224,618 | 151,333 |  | 69,194 | 8,370 |
| 04/30/99 | 35 | 0.0260 | 1,348 | 494 | 1,225,966 | 151,452 |  | 70,542 | 8,505 |
| 05/01/99 | 59 | 0.0260 | 2,272 | 1,121 | 1,228,238 | 151,654 |  | 72,814 | 8,779 |
| 05/02/99 | 92 | 0.0260 | 3,543 | 1,174 | 1,231,782 | 151,968 |  | 76,357 | 9,168 |
| 05/03/99 | 34 | 0.0260 | 1,309 | 1,543 | 1,233,091 | 152,090 |  | 77,667 | 9,413 |
| 05/04/99 | 15 | 0.0260 | 578 | 1,013 | 1,233,669 | 152,144 |  | 78,244 | 9,518 |
| 05/05/99 | 67 | 0.0260 | 2,586 | 2,010 | 1,236,254 | 152,384 |  | 80,830 | 9,951 |
| 05/06/99 | 119 | 0.0260 | 4,583 | 1,387 | 1,240,837 | 152,790 |  | 85,413 | 10,440 |
| 05/07/99 | 55 | 0.0260 | 2,118 | 1,345 | 1,242,956 | 152,982 |  | 87,531 | 10,711 |
| 05/08/99 | 64 | 0.0260 | 2,465 | 362 | 1,245,420 | 153,198 |  | 89,996 | 10,934 |
| 05/09/99 | 68 | 0.0260 | 2,619 | 374 | 1,248,039 | 153,428 |  | 92,615 | 11,169 |
| 05/10/99 | 55 | 0.0260 | 2,118 | 914 | 1,250,157 | 153,617 | Smolt | 2,118 | 894 |
| 05/11/99 | 23 | 0.0260 | 886 | 802 | 1,251,043 | 153,697 |  | 3,004 | 1,218 |
| 05/12/99 | 62 | 0.0260 | 2,388 | 794 | 1,253,431 | 153,908 |  | 5,392 | 1,506 |
| 05/13/99 | 47 | 0.0260 | 1,810 | 518 | 1,255,241 | 154,068 |  | 7,202 | 1,658 |
| 05/14/99 | 72 | 0.0260 | 2,773 | 706 | 1,258,014 | 154,313 |  | 9,975 | 1,919 |
| 05/15/99 | 40 | 0.0260 | 1,541 | 647 | 1,259,554 | 154,450 |  | 11,515 | 2,106 |
| 05/16/99 | 50 | 0.0260 | 1,926 | 291 | 1,261,480 | 154,620 |  | 13,441 | 2,236 |
| 05/17/99 | 49 | 0.0260 | 1,887 | 1,025 | 1,263,367 | 154,789 |  | 15,328 | 2,569 |
| 05/18/99 | 95 | 0.0260 | 3,659 | 1,219 | 1,267,026 | 155,116 |  | 18,987 | 3,051 |
| 05/19/99 | 38 | 0.0260 | 1,464 | 1,248 | 1,268,489 | 155,250 |  | 20,450 | 3,386 |
| 05/20/99 | 40 | 0.0260 | 1,541 | 1,931 | 1,270,030 | 155,397 |  | 21,991 | 3,985 |
| 05/21/99 | 126 | 0.0260 | 4,853 | 1,924 | 1,274,882 | 155,837 |  | 26,843 | 4,680 |
| 05/22/99 | 122 | 0.0260 | 4,687 | 1,189 | 1,279,570 | 156,254 |  | 31,531 | 5,103 |
| 05/23/99 | 75 | 0.0260 | 2,888 | 1,198 | 1,282,458 | 156,514 |  | 34,419 | 5,427 |
| 05/24/99 | 65 | 0.0260 | 2,503 | 1,301 | 1,284,961 | 156,741 |  | 36,922 | 5,746 |
| 05/25/99 | 13 | 0.0260 | 501 | 2,012 | 1,285,462 | 156,798 |  | 37,423 | 6,121 |
| 05/26/99 | 118 | 0.0260 | 4,545 | 2,073 | 1,290,007 | 157,213 |  | 41,967 | 6,742 |
| 05/27/99 | 73 | 0.0260 | 2,811 | 1,767 | 1,292,818 | 157,472 |  | 44,779 | 7,151 |
| 05/28/99 | 27 | 0.0260 | 1,040 | 894 | 1,293,858 | 157,566 |  | 45,819 | 7,277 |
| 05/29/99 | 55 | 0.0260 | 2,107 | 630 | 1,295,965 | 157,754 |  | 47,926 | 7,446 |
| 05/30/99 | 53 | 0.0260 | 2,024 | 236 | 1,297,989 | 157,934 |  | 49,950 | 7,590 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 05/31/99 | 49 | 0.0260 | 1,897 | 697 | 1,299,886 | 158,103 |  | 51,847 | 7,756 |
| 06/01/99 | 81 | 0.0260 | 3,120 | 1,089 | 1,303,006 | 158,384 |  | 54,966 | 8,053 |
| 06/02/99 | 27 | 0.0260 | 1,040 | 1,046 | 1,304,045 | 158,479 |  | 56,006 | 8,196 |
| 06/03/99 | 59 | 0.0260 | 2,272 | 709 | 1,306,318 | 158,683 |  | 58,278 | 8,393 |
| 06/04/99 | 55 | 0.0260 | 2,118 | 787 | 1,308,436 | 158,873 |  | 60,397 | 8,588 |
| 06/05/99 | 23 | 0.0260 | 886 | 843 | 1,309,322 | 158,954 |  | 61,282 | 8,696 |
| 06/06/99 | 13 | 0.0260 | 507 | 346 | 1,309,828 | 158,999 |  | 61,789 | 8,742 |
| 06/07/99 | 31 | 0.0260 | 1,194 | 404 | 1,311,022 | 159,106 |  | 62,983 | 8,842 |
| 06/08/99 | 30 | 0.0260 | 1,155 | 144 | 1,312,178 | 159,208 |  | 64,138 | 8,932 |
| 06/09/99 | 34 | 0.0260 | 1,309 | 347 | 1,313,487 | 159,325 |  | 65,448 | 9,040 |
| 06/10/99 | 18 | 0.0260 | 693 | 322 | 1,314,180 | 159,387 |  | 66,141 | 9,100 |
| 06/11/99 | 29 | 0.0260 | 1,117 | 241 | 1,315,297 | 159,487 |  | 67,258 | 9,191 |
| 06/12/99 | 24 | 0.0260 | 911 | 206 | 1,316,208 | 159,568 |  | 68,169 | 9,265 |
| 06/13/99 | 20 | 0.0260 | 750 | 350 | 1,316,958 | 159,635 |  | 68,919 | 9,332 |
| 06/14/99 | 7 | 0.0260 | 253 | 286 | 1,317,211 | 159,658 |  | 69,172 | 9,356 |
| 06/15/99 | 7 | 0.0260 | 253 | 123 | 1,317,464 | 159,680 |  | 69,425 | 9,377 |
| 06/16/99 | 12 | 0.0260 | 462 | 170 | 1,317,927 | 159,722 |  | 69,887 | 9,416 |
| 06/17/99 | 15 | 0.0260 | 578 | 166 | 1,318,504 | 159,773 |  | 70,465 | 9,464 |
| 06/18/99 | 7 | 0.0260 | 270 | 180 | 1,318,774 | 159,797 |  | 70,735 | 9,487 |
| 06/19/99 | 7 | 0.0260 | 268 | 36 | 1,319,042 | 159,821 |  | 71,003 | 9,509 |
| 06/20/99 | 6 | 0.0260 | 229 | 103 | 1,319,271 | 159,842 |  | 71,232 | 9,528 |
| 06/21/99 | 2 | 0.0260 | 77 | 212 | 1,319,348 | 159,849 |  | 71,309 | 9,537 |
| 06/22/99 | 13 | 0.0260 | 498 | 218 | 1,319,846 | 159,893 |  | 71,807 | 9,580 |
| 06/23/99 | 6 | 0.0260 | 231 | 155 | 1,320,077 | 159,914 |  | 72,038 | 9,600 |
| 06/24/99 | 6 | 0.0260 | 231 | 33 | 1,320,308 | 159,934 |  | 72,269 | 9,619 |
| 06/25/99 | 5 | 0.0260 | 193 | 43 | 1,320,501 | 159,952 |  | 72,462 | 9,634 |
| 06/26/99 | 4 | 0.0260 | 154 | 32 | 1,320,655 | 159,965 |  | 72,616 | 9,647 |
| 06/27/99 | 4 | 0.0260 | 140 | 55 | 1,320,794 | 159,978 |  | 72,755 | 9,659 |
| 06/28/99 | 1 | 0.0260 | 55 | 54 | 1,320,850 | 159,983 |  | 72,811 | 9,663 |
| 06/29/99 | 1 | 0.0260 | 39 | 62 | 1,320,888 | 159,986 |  | 72,849 | 9,667 |
| 06/30/99 | 4 | 0.0260 | 154 | 83 | 1,321,042 | 160,000 |  | 73,003 | 9,680 |

## Appendix A.6.a. 1996 Outmigration index estimates based on efficiency predictor that excluded turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life- <br> Stage <br> Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 02/06/96 | 89.0 | 0.2032 | 438 | 172 | 438 | 172 | Fry | 438 | 172 |
| 02/07/96 | 99.2 | 0.2340 | 424 | 194 | 862 | 350 |  | 862 | 350 |
| 02/08/96 | 49.5 | 0.2407 | 206 | 184 | 1,068 | 457 |  | 1,068 | 457 |
| 02/09/96 | 13.0 | 0.2412 | 54 | 97 | 1,122 | 485 |  | 1,122 | 485 |
| 02/10/96 | 2.0 | 0.2204 | 9 | 30 | 1,131 | 489 |  | 1,131 | 489 |
| 02/11/96 | 0.0 | 0.1870 | 0 | 15 | 1,131 | 490 |  | 1,131 | 490 |
| 02/12/96 | 6.0 | 0.1731 | 35 | 21 | 1,165 | 502 |  | 1,165 | 502 |
| 02/13/96 | 2.0 | 0.1567 | 13 | 83 | 1,178 | 513 |  | 1,178 | 513 |
| 02/14/96 | 28.0 | 0.1256 | 223 | 164 | 1,401 | 604 |  | 1,401 | 604 |
| 02/15/96 | 39.0 | 0.0915 | 426 | 186 | 1,827 | 752 |  | 1,827 | 752 |
| 02/16/96 | 21.3 | 0.0730 | 292 | 188 | 2,119 | 864 |  | 2,119 | 864 |
| 02/17/96 | 44.0 | 0.0710 | 620 | 333 | 2,739 | 1,103 |  | 2,739 | 1,103 |
| 02/18/96 | 57.0 | 0.0709 | 804 | 318 | 3,543 | 1,389 |  | 3,543 | 1,389 |

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| Date | Count | Efficiency | Outmigration |  |  |  | $\begin{gathered} \hline \text { Life- } \\ \text { Stage } \\ \text { Cohort } \end{gathered}$ | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 02/19/96 | 52.0 | 0.0594 | 876 | 373 | 4,419 | 1,711 |  | 4,419 | 1,711 |
| 02/20/96 | 37.0 | 0.0326 | 1,134 | 506 | 5,553 | 2,127 |  | 5,553 | 2,127 |
| 02/21/96 | 48.4 | 0.0175 | 2,763 | 1,256 | 8,316 | 3,199 |  | 8,316 | 3,199 |
| 02/22/96 | 43.1 | 0.0179 | 2,406 | 2,220 | 10,722 | 4,620 |  | 10,722 | 4,620 |
| 02/23/96 | 113.0 | 0.0191 | 5,909 | 3,457 | 16,631 | 7,337 |  | 16,631 | 7,337 |
| 02/24/96 | 18.0 | 0.0151 | 1,195 | 3,188 | 17,826 | 8,425 |  | 17,826 | 8,425 |
| 02/25/96 | 24.0 | 0.0140 | 1,717 | 887 | 19,543 | 9,092 |  | 19,543 | 9,092 |
| 02/26/96 | 11.0 | 0.0137 | 802 | 561 | 20,344 | 9,415 |  | 20,344 | 9,415 |
| 02/27/96 | 16.0 | 0.0124 | 1,289 | 618 | 21,634 | 9,916 |  | 21,634 | 9,916 |
| 02/28/96 | 11.0 | 0.0106 | 1,033 | 657 | 22,667 | 10,325 |  | 22,667 | 10,325 |
| 02/29/96 | 5.0 | 0.0076 | 655 | 484 | 23,322 | 10,590 |  | 23,322 | 10,590 |
| 03/01/96 | 6.0 | 0.0072 | 837 | 475 | 24,159 | 10,972 |  | 24,159 | 10,972 |
| 03/02/96 | 7.7 | 0.0070 | 1,109 | 589 | 25,267 | 11,470 |  | 25,267 | 11,470 |
| 03/03/96 | 6.8 | 0.0068 | 993 | 541 | 26,261 | 11,910 |  | 26,261 | 11,910 |
| 03/04/96 | 4.8 | 0.0067 | 721 | 447 | 26,982 | 12,226 |  | 26,982 | 12,226 |
| 03/05/96 | 2.8 | 0.0060 | 459 | 417 | 27,441 | 12,428 |  | 27,441 | 12,428 |
| 03/06/96 | 0.0 | 0.0057 | 0 | 312 | 27,441 | 12,432 |  | 27,441 | 12,432 |
| 03/07/96 | 4.0 | 0.0058 | 684 | 477 | 28,125 | 12,714 |  | 28,125 | 12,714 |
| 03/08/96 | 4.0 | 0.0061 | 659 | 405 | 28,785 | 12,991 |  | 28,785 | 12,991 |
| 03/09/96 | 1.0 | 0.0063 | 160 | 300 | 28,944 | 13,062 |  | 28,944 | 13,062 |
| 03/10/96 | 0.0 | 0.0067 | 0 | 75 | 28,944 | 13,062 |  | 28,944 | 13,062 |
| 03/11/96 | 0.0 | 0.0077 | 0 | 66 | 28,944 | 13,062 |  | 28,944 | 13,062 |
| 03/12/96 | 1.0 | 0.0088 | 113 | 79 | 29,058 | 13,112 |  | 29,058 | 13,112 |
| 03/13/96 | 0.0 | 0.0085 | 0 | 60 | 29,058 | 13,112 |  | 29,058 | 13,112 |
| 03/14/96 | 1.0 | 0.0082 | 122 | 82 | 29,179 | 13,159 |  | 29,179 | 13,159 |
| 03/15/96 | 0.0 | 0.0076 | 0 | 69 | 29,179 | 13,159 |  | 29,179 | 13,159 |
| 03/16/96 | 1.0 | 0.0067 | 148 | 99 | 29,328 | 13,209 |  | 29,328 | 13,209 |
| 03/17/96 | 0.0 | 0.0081 | 0 | 111 | 29,328 | 13,210 |  | 29,328 | 13,210 |
| 03/18/96 | 2.0 | 0.0097 | 207 | 143 | 29,534 | 13,299 |  | 29,534 | 13,299 |
| 03/19/96 | 0.0 | 0.0088 | 0 | 102 | 29,534 | 13,300 |  | 29,534 | 13,300 |
| 03/20/96 | 1.0 | 0.0081 | 124 | 84 | 29,658 | 13,347 |  | 29,658 | 13,347 |
| 03/21/96 | 0.0 | 0.0070 | 0 | 75 | 29,658 | 13,347 | Parr | 0 | 0 |
| 03/22/96 | 0.0 | 0.0063 | 0 | 0 | 29,658 | 13,347 |  | 0 | 0 |
| 03/23/96 | 0.0 | 0.0071 | 0 | 0 | 29,658 | 13,347 |  | 0 | 0 |
| 03/24/96 | 0.0 | 0.0085 | 0 | 0 | 29,658 | 13,347 |  | 0 | 0 |
| 03/25/96 | 0.0 | 0.0094 | 0 | 239 | 29,658 | 13,349 |  | 0 | 239 |
| 03/26/96 | 4.0 | 0.0107 | 374 | 200 | 30,032 | 13,369 |  | 374 | 311 |
| 03/27/96 | 2.0 | 0.0136 | 147 | 184 | 30,179 | 13,377 |  | 521 | 367 |
| 03/28/96 | 7.0 | 0.0136 | 516 | 308 | 30,695 | 13,383 |  | 1,037 | 499 |
| 03/29/96 | 10.0 | 0.0148 | 677 | 264 | 31,372 | 13,378 |  | 1,714 | 605 |
| 03/30/96 | 3.0 | 0.0206 | 146 | 175 | 31,518 | 13,385 |  | 1,860 | 638 |
| 03/31/96 | 5.0 | 0.0217 | 230 | 58 | 31,749 | 13,391 |  | 2,090 | 654 |
| 04/01/96 | 3.0 | 0.0151 | 199 | 86 | 31,947 | 13,378 | Smolt | 199 | 42 |
| 04/02/96 | 3.0 | 0.0140 | 215 | 207 | 32,162 | 13,362 |  | 413 | 221 |
| 04/03/96 | 8.0 | 0.0165 | 485 | 464 | 32,647 | 13,352 |  | 898 | 529 |
| 04/04/96 | 18.0 | 0.0170 | 1,059 | 365 | 33,706 | 13,326 |  | 1,957 | 690 |
| 04/05/96 | 9.0 | 0.0169 | 531 | 277 | 34,237 | 13,316 |  | 2,489 | 784 |
| 04/06/96 | 14.0 | 0.0149 | 942 | 264 | 35,179 | 13,265 |  | 3,431 | 927 |

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| Date | Count | Efficiency | Outmigration |  |  |  | $\begin{gathered} \text { Life- } \\ \text { Stage } \\ \text { Cohort } \end{gathered}$ | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 04/07/96 | 13.0 | 0.0201 | 646 | 365 | 35,825 | 13,281 |  | 4,076 | 1,042 |
| 04/08/96 | 1.0 | 0.0181 | 55 | 329 | 35,880 | 13,285 |  | 4,131 | 1,098 |
| 04/09/96 | 8.0 | 0.0158 | 507 | 240 | 36,387 | 13,264 |  | 4,638 | 1,185 |
| 04/10/96 | 4.0 | 0.0183 | 218 | 168 | 36,605 | 13,264 |  | 4,856 | 1,218 |
| 04/11/96 | 2.0 | 0.0162 | 124 | 220 | 36,728 | 13,261 |  | 4,980 | 1,254 |
| 04/12/96 | 9.0 | 0.0177 | 508 | 240 | 37,236 | 13,253 |  | 5,488 | 1,331 |
| 04/13/96 | 2.0 | 0.0207 | 97 | 227 | 37,333 | 13,254 |  | 5,585 | 1,359 |
| 04/14/96 | 0.0 | 0.0207 | 0 | 252 | 37,333 | 13,256 |  | 5,585 | 1,382 |
| 04/15/96 | 10.0 | 0.0194 | 517 | 284 | 37,850 | 13,232 |  | 6,101 | 1,469 |
| 04/16/96 | 2.0 | 0.0135 | 148 | 288 | 37,998 | 13,216 |  | 6,249 | 1,526 |
| 04/17/96 | 5.2 | 0.0151 | 344 | 159 | 38,342 | 13,188 |  | 6,593 | 1,590 |
| 04/18/96 | 6.0 | 0.0177 | 340 | 309 | 38,682 | 13,180 |  | 6,933 | 1,662 |
| 04/19/96 | 15.0 | 0.0172 | 874 | 438 | 39,555 | 13,147 |  | 7,807 | 1,832 |
| 04/20/96 | 1.0 | 0.0162 | 62 | 643 | 39,617 | 13,159 |  | 7,868 | 1,951 |
| 04/21/96 | 22.0 | 0.0159 | 1,380 | 1,120 | 40,997 | 13,113 |  | 9,249 | 2,430 |
| 04/22/96 | 36.0 | 0.0154 | 2,345 | 1,100 | 43,342 | 12,991 |  | 11,594 | 2,994 |
| 04/23/96 | 52.1 | 0.0160 | 3,251 | 889 | 46,593 | 12,844 |  | 14,845 | 3,597 |
| 04/24/96 | 38.0 | 0.0160 | 2,372 | 706 | 48,965 | 12,762 |  | 17,217 | 4,062 |
| 04/25/96 | 39.0 | 0.0149 | 2,623 | 644 | 51,588 | 12,656 |  | 19,840 | 4,622 |
| 04/26/96 | 38.0 | 0.0153 | 2,488 | 2,157 | 54,076 | 12,765 |  | 22,327 | 5,551 |
| 04/27/96 | 95.0 | 0.0155 | 6,137 | 2,758 | 60,213 | 12,948 |  | 28,465 | 7,195 |
| 04/28/96 | 109.0 | 0.0152 | 7,150 | 1,827 | 67,363 | 13,108 |  | 35,615 | 8,715 |
| 04/29/96 | 89.0 | 0.0150 | 5,939 | 1,850 | 73,302 | 13,453 |  | 41,554 | 10,091 |
| 04/30/96 | 125.0 | 0.0152 | 8,208 | 2,287 | 81,510 | 14,146 |  | 49,761 | 11,939 |
| 05/01/96 | 93.8 | 0.0151 | 6,190 | 2,003 | 87,699 | 14,839 |  | 55,951 | 13,384 |
| 05/02/96 | 84.0 | 0.0143 | 5,878 | 1,640 | 93,578 | 15,584 |  | 61,829 | 14,833 |
| 05/03/96 | 75.2 | 0.0147 | 5,124 | 1,404 | 98,702 | 16,302 |  | 66,953 | 16,075 |
| 05/04/96 | 67.0 | 0.0157 | 4,259 | 1,625 | 102,960 | 16,949 |  | 71,212 | 17,053 |
| 05/05/96 | 107.0 | 0.0145 | 7,396 | 2,387 | 110,357 | 18,222 |  | 78,608 | 18,943 |
| 05/06/96 | 73.0 | 0.0158 | 4,608 | 2,258 | 114,964 | 19,064 |  | 83,216 | 20,061 |
| 05/07/96 | 42.0 | 0.0153 | 2,750 | 1,249 | 117,714 | 19,564 |  | 85,965 | 20,730 |
| 05/08/96 | 47.0 | 0.0153 | 3,077 | 762 | 120,791 | 20,102 |  | 89,042 | 21,445 |
| 05/09/96 | 47.0 | 0.0155 | 3,030 | 734 | 123,820 | 20,635 |  | 92,072 | 22,135 |
| 05/10/96 | 52.1 | 0.0157 | 3,315 | 856 | 127,135 | 21,221 |  | 95,387 | 22,870 |
| 05/11/96 | 60.0 | 0.0153 | 3,933 | 1,645 | 131,068 | 21,995 |  | 99,320 | 23,825 |
| 05/12/96 | 20.0 | 0.0158 | 1,269 | 1,277 | 132,337 | 22,261 |  | 100,588 | 24,142 |
| 05/13/96 | 35.8 | 0.0154 | 2,335 | 783 | 134,672 | 22,711 |  | 102,923 | 24,694 |
| 05/14/96 | 33.9 | 0.0148 | 2,288 | 623 | 136,960 | 23,168 |  | 105,212 | 25,255 |
| 05/15/96 | 28.7 | 0.0127 | 2,267 | 897 | 139,226 | 23,715 |  | 107,478 | 25,936 |
| 05/16/96 | 19.0 | 0.0152 | 1,252 | 668 | 140,478 | 23,967 |  | 108,730 | 26,230 |
| 05/17/96 | 10.0 | 0.0140 | 713 | 362 | 141,192 | 24,121 |  | 109,443 | 26,414 |
| 05/18/96 | 14.0 | 0.0134 | 1,042 | 337 | 142,233 | 24,358 |  | 110,485 | 26,700 |
| 05/19/96 | 10.0 | 0.0122 | 822 | 439 | 143,055 | 24,569 |  | 111,307 | 26,959 |
| 05/20/96 | 19.0 | 0.0135 | 1,408 | 613 | 144,463 | 24,893 |  | 112,715 | 27,341 |
| 05/21/96 | 23.0 | 0.0135 | 1,705 | 660 | 146,168 | 25,288 |  | 114,420 | 27,809 |
| 05/22/96 | 10.4 | 0.0153 | 681 | 518 | 146,850 | 25,436 |  | 115,101 | 27,985 |
| 05/23/96 | 9.0 | 0.0190 | 473 | 275 | 147,322 | 25,525 |  | 115,574 | 28,103 |
| 05/24/96 | 18.0 | 0.0267 | 675 | 460 | 147,997 | 25,624 |  | 116,249 | 28,247 |

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| Date | Count | Efficiency | Outmigration |  |  |  | $\begin{gathered} \hline \text { Life- } \\ \text { Stage } \\ \text { Cohort } \end{gathered}$ | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 05/25/96 | 32.5 | 0.0299 | 1,086 | 613 | 149,083 | 25,769 |  | 117,334 | 28,473 |
| 05/26/96 | 52.0 | 0.0302 | 1,722 | 568 | 150,804 | 25,994 |  | 119,056 | 28,825 |
| 05/27/96 | 30.0 | 0.0308 | 974 | 628 | 151,778 | 26,119 |  | 120,030 | 29,014 |
| 05/28/96 | 15.0 | 0.0292 | 514 | 278 | 152,292 | 26,190 |  | 120,544 | 29,120 |
| 05/29/96 | 22.0 | 0.0308 | 714 | 260 | 153,007 | 26,278 |  | 121,258 | 29,254 |
| 05/30/96 | 9.0 | 0.0308 | 292 | 238 | 153,298 | 26,315 |  | 121,550 | 29,310 |
| 05/31/96 | 10.0 | 0.0287 | 348 | 89 | 153,647 | 26,366 |  | 121,898 | 29,385 |
| 06/01/96 | 10.0 | 0.0270 | 370 | 103 | 154,017 | 26,425 |  | 122,268 | 29,473 |
| 06/02/96 | 11.0 | 0.0275 | 400 | 202 | 154,416 | 26,488 |  | 122,668 | 29,566 |
| 06/03/96 | 2.0 | 0.0319 | 63 | 140 | 154,479 | 26,496 |  | 122,731 | 29,577 |
| 06/04/96 | 5.2 | 0.0265 | 197 | 107 | 154,676 | 26,529 |  | 122,927 | 29,626 |
| 06/05/96 | 7.0 | 0.0244 | 287 | 119 | 154,963 | 26,584 |  | 123,214 | 29,705 |
| 06/06/96 | 3.0 | 0.0259 | 116 | 118 | 155,078 | 26,605 |  | 123,330 | 29,735 |
| 06/07/96 | 1.0 | 0.0323 | 31 | 47 | 155,109 | 26,608 |  | 123,361 | 29,741 |
| 06/08/96 | 4.0 | 0.0216 | 186 | 86 | 155,295 | 26,645 |  | 123,547 | 29,790 |
| 06/09/96 | 2.0 | 0.0219 | 91 | 62 | 155,386 | 26,660 |  | 123,638 | 29,810 |
| 06/10/96 | 1.5 | 0.0232 | 65 | 46 | 155,451 | 26,670 |  | 123,702 | 29,823 |
| 06/11/96 | 0.0 | 0.0241 | 0 | 61 | 155,451 | 26,671 |  | 123,702 | 29,823 |
| 06/12/96 | 3.0 | 0.0251 | 120 | 63 | 155,570 | 26,686 |  | 123,822 | 29,841 |
| 06/13/96 | 2.0 | 0.0232 | 86 | 33 | 155,656 | 26,699 |  | 123,908 | 29,858 |
| 06/14/96 | 3.2 | 0.0266 | 122 | 64 | 155,778 | 26,713 |  | 124,030 | 29,876 |
| 06/15/96 | 0.0 | 0.0292 | 0 | 63 | 155,778 | 26,714 |  | 124,030 | 29,876 |
| 06/16/96 | 0.0 | 0.0329 | 0 | 17 | 155,778 | 26,714 |  | 124,030 | 29,876 |
| 06/17/96 | 1.0 | 0.0363 | 28 | 16 | 155,806 | 26,716 |  | 124,057 | 29,880 |
| 06/18/96 | 0.0 | 0.0361 | 0 | 16 | 155,806 | 26,716 |  | 124,057 | 29,880 |
| 06/19/96 | 0.0 | 0.0345 | 0 | 0 | 155,806 | 26,716 |  | 124,057 | 29,880 |
| 06/20/96 | 0.0 | 0.0333 | 0 | 17 | 155,806 | 26,716 |  | 124,057 | 29,880 |
| 06/21/96 | 1.0 | 0.0321 | 31 | 19 | 155,837 | 26,720 |  | 124,088 | 29,887 |
| 06/22/96 | 0.0 | 0.0262 | 0 | 21 | 155,837 | 26,720 |  | 124,088 | 29,887 |
| 06/23/96 | 1.0 | 0.0215 | 47 | 31 | 155,884 | 26,732 |  | 124,135 | 29,904 |
| 06/24/96 | 1.0 | 0.0254 | 39 | 25 | 155,923 | 26,740 |  | 124,175 | 29,915 |
| 06/25/96 | 0.0 | 0.0276 | 0 | 20 | 155,923 | 26,740 |  | 124,175 | 29,915 |
| 06/26/96 | 0.0 | 0.0300 | 0 | 19 | 155,923 | 26,740 |  | 124,175 | 29,915 |
| 06/27/96 | 1.0 | 0.0333 | 30 | 18 | 155,953 | 26,743 |  | 124,205 | 29,921 |
| 06/28/96 | 0.0 | 0.0297 | 0 | 19 | 155,953 | 26,743 |  | 124,205 | 29,921 |
| 06/29/96 | 0.0 | 0.0267 | 0 | 20 | 155,953 | 26,743 |  | 124,205 | 29,921 |
| 06/30/96 | 1.0 | 0.0240 | 42 | 27 | 155,995 | 26,753 |  | 124,246 | 29,935 |
| 07/01/96 | 1.0 | 0.0301 | 33 | 10 | 156,028 | 26,758 |  | 124,279 | 29,943 |

Appendix A.6.b. 1997 Outmigration index estimates based on efficiency predictor that excluded turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/19/97 | 15.0 | 0.0330 | 455 | 66 | 455 | 66 | Parr | 455 | 66 |
| 03/20/97 | 17.0 | 0.0262 | 648 | 422 | 1,103 | 431 |  | 1,103 | 431 |
| 03/21/97 | 35.0 | 0.0269 | 1,299 | 410 | 2,402 | 610 |  | 2,402 | 610 |
| 03/22/97 | 36.0 | 0.0270 | 1,333 | 288 | 3,735 | 705 |  | 3,735 | 705 |
| 03/23/97 | 48.0 | 0.0297 | 1,615 | 235 | 5,350 | 785 |  | 5,350 | 785 |
| 03/24/97 | 42.0 | 0.0370 | 1,136 | 232 | 6,486 | 848 |  | 6,486 | 848 |
| 03/25/97 | 32.0 | 0.0478 | 670 | 146 | 7,156 | 874 |  | 7,156 | 874 |
| 03/26/97 | 30.0 | 0.0612 | 490 | 105 | 7,646 | 889 |  | 7,646 | 889 |
| 03/27/97 | 22.0 | 0.0682 | 323 | 76 | 7,968 | 899 |  | 7,968 | 899 |
| 03/28/97 | 28.0 | 0.0656 | 427 | 82 | 8,395 | 916 |  | 8,395 | 916 |
| 03/29/97 | 21.0 | 0.0730 | 288 | 228 | 8,682 | 954 |  | 8,682 | 954 |
| 03/30/97 | 52.5 | 0.0585 | 897 | 306 | 9,579 | 1,041 |  | 9,579 | 1,041 |
| 03/31/97 | 30.0 | 0.0613 | 490 | 198 | 10,069 | 1,088 |  | 10,069 | 1,088 |
| 04/01/97 | 45.0 | 0.0666 | 675 | 183 | 10,744 | 1,144 |  | 10,744 | 1,144 |
| 04/02/97 | 24.3 | 0.0602 | 404 | 194 | 11,148 | 1,189 |  | 11,148 | 1,189 |
| 04/03/97 | 27.0 | 0.0576 | 468 | 79 | 11,616 | 1,230 |  | 11,616 | 1,230 |
| 04/04/97 | 28.0 | 0.0641 | 437 | 193 | 12,053 | 1,280 |  | 12,053 | 1,280 |
| 04/05/97 | 48.0 | 0.0628 | 765 | 225 | 12,818 | 1,364 |  | 12,818 | 1,368 |
| 04/06/97 | 51.0 | 0.0599 | 851 | 163 | 13,669 | 1,454 | Smolt | 851 | 189 |
| 04/07/97 | 39.0 | 0.0569 | 685 | 243 | 14,354 | 1,546 |  | 1,536 | 349 |
| 04/08/97 | 26.0 | 0.0563 | 462 | 196 | 14,816 | 1,610 |  | 1,998 | 442 |
| 04/09/97 | 46.4 | 0.0603 | 771 | 303 | 15,587 | 1,720 |  | 2,769 | 598 |
| 04/10/97 | 60.0 | 0.0609 | 985 | 201 | 16,572 | 1,838 |  | 3,754 | 721 |
| 04/11/97 | 44.0 | 0.0560 | 786 | 193 | 17,358 | 1,944 |  | 4,540 | 838 |
| 04/12/97 | 49.0 | 0.0531 | 923 | 151 | 18,282 | 2,061 |  | 5,464 | 960 |
| 04/13/97 | 45.0 | 0.0403 | 1,116 | 335 | 19,397 | 2,200 |  | 6,579 | 1,122 |
| 04/14/97 | 68.0 | 0.0246 | 2,770 | 1,736 | 22,167 | 2,943 |  | 9,349 | 2,159 |
| 04/15/97 | 126.5 | 0.0193 | 6,544 | 1,783 | 28,711 | 3,692 |  | 15,893 | 2,993 |
| 04/16/97 | 104.5 | 0.0184 | 5,673 | 1,542 | 34,384 | 4,443 |  | 21,566 | 3,791 |
| 04/17/97 | 81.0 | 0.0186 | 4,364 | 1,806 | 38,748 | 5,231 |  | 25,930 | 4,629 |
| 04/18/97 | 43.0 | 0.0188 | 2,290 | 1,614 | 41,038 | 5,715 |  | 28,219 | 5,144 |
| 04/19/97 | 22.0 | 0.0201 | 1,093 | 752 | 42,131 | 5,867 |  | 29,313 | 5,301 |
| 04/20/97 | 51.0 | 0.0179 | 2,855 | 980 | 44,985 | 6,288 |  | 32,167 | 5,733 |
| 04/21/97 | 28.0 | 0.0186 | 1,502 | 669 | 46,487 | 6,498 |  | 33,669 | 5,948 |
| 04/22/97 | 44.0 | 0.0179 | 2,460 | 1,033 | 48,946 | 6,898 |  | 36,128 | 6,357 |
| 04/23/97 | 10.0 | 0.0173 | 579 | 1,138 | 49,526 | 7,076 |  | 36,708 | 6,543 |
| 04/24/97 | 9.0 | 0.0180 | 499 | 528 | 50,025 | 7,162 |  | 37,207 | 6,632 |
| 04/25/97 | 26.0 | 0.0182 | 1,430 | 691 | 51,455 | 7,380 |  | 38,636 | 6,854 |
| 04/26/97 | 32.0 | 0.0178 | 1,800 | 573 | 53,254 | 7,646 |  | 40,436 | 7,124 |
| 04/27/97 | 15.0 | 0.0178 | 842 | 540 | 54,096 | 7,779 |  | 41,278 | 7,259 |
| 04/28/97 | 16.5 | 0.0154 | 1,072 | 312 | 55,169 | 7,975 |  | 42,351 | 7,458 |
| 04/29/97 | 21.0 | 0.0176 | 1,191 | 364 | 56,360 | 8,156 |  | 43,541 | 7,641 |
| 04/30/97 | 27.0 | 0.0167 | 1,612 | 798 | 57,972 | 8,453 |  | 45,154 | 7,943 |
| 05/01/97 | 3.0 | 0.0145 | 207 | 804 | 58,179 | 8,534 |  | 45,361 | 8,027 |
| 05/02/97 | 15.0 | 0.0173 | 867 | 1,146 | 59,046 | 8,744 |  | 46,228 | 8,243 |
| 05/03/97 | 42.0 | 0.0177 | 2,374 | 869 | 61,421 | 9,143 |  | 48,603 | 8,646 |
| 05/04/97 | 28.0 | 0.0168 | 1,668 | 670 | 63,089 | 9,453 |  | 50,271 | 8,960 |
| 05/05/97 | 47.0 | 0.0179 | 2,627 | 1,150 | 65,716 | 9,923 |  | 52,898 | 9,434 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 05/06/97 | 9.0 | 0.0158 | 568 | 1,183 | 66,284 | 10,103 |  | 53,466 | 9,619 |
| 05/07/97 | 32.0 | 0.0161 | 1,988 | 879 | 68,273 | 10,515 |  | 55,455 | 10,034 |
| 05/08/97 | 29.0 | 0.0169 | 1,711 | 362 | 69,984 | 10,820 |  | 57,166 | 10,341 |
| 05/09/97 | 31.0 | 0.0173 | 1,794 | 425 | 71,778 | 11,133 |  | 58,960 | 10,656 |
| 05/10/97 | 23.0 | 0.0179 | 1,286 | 374 | 73,064 | 11,344 |  | 60,246 | 10,868 |
| 05/11/97 | 21.0 | 0.0166 | 1,266 | 625 | 74,330 | 11,597 |  | 61,512 | 11,123 |
| 05/12/97 | 38.5 | 0.0182 | 2,111 | 665 | 76,441 | 11,950 |  | 63,623 | 11,478 |
| 05/13/97 | 38.5 | 0.0177 | 2,169 | 492 | 78,610 | 12,339 |  | 65,792 | 11,869 |
| 05/14/97 | 31.0 | 0.0233 | 1,333 | 304 | 79,943 | 12,570 |  | 67,125 | 12,096 |
| 05/15/97 | 35.6 | 0.0226 | 1,575 | 1,015 | 81,518 | 12,895 |  | 68,700 | 12,420 |
| 05/16/97 | 71.5 | 0.0169 | 4,238 | 1,651 | 85,756 | 13,821 |  | 72,938 | 13,353 |
| 05/17/97 | 27.5 | 0.0173 | 1,589 | 1,309 | 87,345 | 14,182 |  | 74,527 | 13,716 |
| 05/18/97 | 42.0 | 0.0170 | 2,469 | 1,127 | 89,813 | 14,701 |  | 76,995 | 14,238 |
| 05/19/97 | 62.0 | 0.0170 | 3,638 | 995 | 93,451 | 15,421 |  | 80,633 | 14,961 |
| 05/20/97 | 44.0 | 0.0174 | 2,530 | 1,216 | 95,981 | 15,943 |  | 83,163 | 15,485 |
| 05/21/97 | 23.0 | 0.0178 | 1,290 | 642 | 97,271 | 16,194 |  | 84,452 | 15,737 |
| 05/22/97 | 30.0 | 0.0165 | 1,816 | 524 | 99,087 | 16,580 |  | 86,269 | 16,124 |
| 05/23/97 | 33.8 | 0.0168 | 2,016 | 498 | 101,103 | 17,024 |  | 88,285 | 16,569 |
| 05/24/97 | 34.6 | 0.0198 | 1,749 | 426 | 102,852 | 17,406 |  | 90,034 | 16,950 |
| 05/25/97 | 31.0 | 0.0248 | 1,251 | 490 | 104,103 | 17,637 |  | 91,285 | 17,179 |
| 05/26/97 | 51.0 | 0.0260 | 1,962 | 836 | 106,066 | 17,976 |  | 93,248 | 17,514 |
| 05/27/97 | 11.0 | 0.0228 | 482 | 866 | 106,548 | 18,086 |  | 93,730 | 17,624 |
| 05/28/97 | 33.0 | 0.0188 | 1,757 | 909 | 108,305 | 18,453 |  | 95,487 | 17,992 |
| 05/29/97 | 42.0 | 0.0166 | 2,530 | 1,052 | 110,836 | 19,005 |  | 98,018 | 18,546 |
| 05/30/97 | 12.5 | 0.0197 | 636 | 948 | 111,471 | 19,137 |  | 98,653 | 18,679 |
| 05/31/97 | 7.0 | 0.0201 | 349 | 244 | 111,820 | 19,207 |  | 99,002 | 18,749 |
| 06/01/97 | 3.0 | 0.0199 | 151 | 199 | 111,971 | 19,242 |  | 99,153 | 18,784 |
| 06/02/97 | 11.0 | 0.0236 | 466 | 189 | 112,438 | 19,325 |  | 99,620 | 18,867 |
| 06/03/97 | 7.0 | 0.0203 | 345 | 228 | 112,783 | 19,391 |  | 99,965 | 18,932 |
| 06/04/97 | 2.0 | 0.0161 | 125 | 177 | 112,907 | 19,420 |  | 100,089 | 18,961 |
| 06/05/97 | 7.0 | 0.0186 | 377 | 194 | 113,285 | 19,486 |  | 100,467 | 19,027 |
| 06/06/97 | 8.6 | 0.0181 | 474 | 182 | 113,758 | 19,577 |  | 100,940 | 19,118 |
| 06/07/97 | 3.0 | 0.0241 | 124 | 146 | 113,883 | 19,596 |  | 101,065 | 19,138 |
| 06/08/97 | 2.0 | 0.0275 | 73 | 401 | 113,955 | 19,614 |  | 101,137 | 19,156 |
| 06/09/97 | 22.0 | 0.0338 | 650 | 346 | 114,606 | 19,720 |  | 101,788 | 19,260 |
| 06/10/97 | 3.7 | 0.0312 | 118 | 306 | 114,723 | 19,745 |  | 101,905 | 19,284 |
| 06/11/97 | 7.0 | 0.0309 | 227 | 75 | 114,950 | 19,789 |  | 102,132 | 19,328 |
| 06/12/97 | 6.0 | 0.0363 | 165 | 40 | 115,115 | 19,812 |  | 102,297 | 19,350 |
| 06/13/97 | 5.0 | 0.0375 | 133 | 46 | 115,248 | 19,829 |  | 102,430 | 19,367 |
| 06/14/97 | 3.0 | 0.0325 | 92 | 50 | 115,341 | 19,846 |  | 102,523 | 19,383 |
| 06/15/97 | 2.0 | 0.0316 | 63 | 66 | 115,404 | 19,858 |  | 102,586 | 19,395 |
| 06/16/97 | 6.0 | 0.0383 | 157 | 71 | 115,561 | 19,877 |  | 102,743 | 19,414 |
| 06/17/97 | 1.2 | 0.0367 | 33 | 65 | 115,594 | 19,882 |  | 102,776 | 19,418 |
| 06/18/97 | 3.0 | 0.0331 | 91 | 46 | 115,685 | 19,898 |  | 102,867 | 19,434 |
| 06/19/97 | 4.0 | 0.0311 | 129 | 35 | 115,813 | 19,924 |  | 102,995 | 19,459 |
| 06/20/97 | 3.0 | 0.0280 | 107 | 35 | 115,920 | 19,948 |  | 103,102 | 19,484 |
| 06/21/97 | 4.0 | 0.0351 | 114 | 27 | 116,035 | 19,966 |  | 103,217 | 19,501 |
| 06/22/97 | 4.0 | 0.0325 | 123 | 43 | 116,158 | 19,987 |  | 103,340 | 19,522 |
| 06/23/97 | 2.0 | 0.0304 | 66 | 51 | 116,223 | 20,001 |  | 103,405 | 19,535 |
| 06/24/97 | 1.0 | 0.0324 | 31 | 31 | 116,254 | 20,006 |  | 103,436 | 19,540 |
| 06/25/97 | 0.0 | 0.0324 | 0 | 17 | 116,254 | 20,006 |  | 103,436 | 19,540 |
| 06/26/97 | 0.0 | 0.0324 | 0 | 0 | 116,254 | 20,006 |  | 103,436 | 19,540 |
| 06/27/97 | 0.0 | 0.0324 | 0 | 0 | 116,254 | 20,006 |  | 103,436 | 19,540 |

## Appendix A.6.c. 1998 Outmigration index estimates based on efficiency predictor that excluded turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 01/29/98 | 802.0 | 0.0458 | 17,500 | 8,225 | 17,500 | 8,225 | Fry | 17,500 | 8,225 |
| 01/30/98 | 286.0 | 0.0455 | 6,291 | 6,032 | 23,791 | 10,386 |  | 23,791 | 10,386 |
| 01/31/98 | 703.9 | 0.0460 | 15,298 | 5,439 | 39,090 | 12,257 |  | 39,090 | 12,257 |
| 02/01/98 | 678.1 | 0.0457 | 14,834 | 5,315 | 53,923 | 14,099 |  | 53,923 | 14,099 |
| 02/02/98 | 1085.0 | 0.0454 | 23,892 | 8,804 | 77,815 | 17,917 |  | 77,815 | 17,917 |
| 02/03/98 | 332.0 | 0.0437 | 7,596 | 8,642 | 85,411 | 20,379 |  | 85,411 | 20,379 |
| 02/04/98 | 643.3 | 0.0441 | 14,595 | 4,778 | 100,006 | 21,903 |  | 100,006 | 21,903 |
| 02/05/98 | 693.1 | 0.0445 | 15,591 | 2,382 | 115,597 | 23,188 |  | 115,597 | 23,188 |
| 02/06/98 | 759.1 | 0.0448 | 16,933 | 2,804 | 132,531 | 24,731 |  | 132,531 | 24,731 |
| 02/07/98 | 850.6 | 0.0452 | 18,814 | 5,440 | 151,345 | 26,945 |  | 151,345 | 26,945 |
| 02/08/98 | 1180.0 | 0.0456 | 25,879 | 5,034 | 177,224 | 29,763 |  | 177,224 | 29,763 |
| 02/09/98 | 1091.1 | 0.0455 | 23,974 | 3,485 | 201,198 | 32,307 |  | 201,198 | 32,307 |
| 02/10/98 | 1045.7 | 0.0454 | 23,020 | 3,143 | 224,218 | 34,820 |  | 224,218 | 34,820 |
| 02/11/98 | 1011.8 | 0.0453 | 22,314 | 3,617 | 246,532 | 37,373 |  | 246,532 | 37,373 |
| 02/12/98 | 862.4 | 0.0453 | 19,056 | 3,022 | 265,588 | 39,577 |  | 265,588 | 39,577 |
| 02/13/98 | 897.0 | 0.0452 | 19,857 | 2,646 | 285,445 | 41,871 |  | 285,445 | 41,871 |
| 02/14/98 | 849.0 | 0.0439 | 19,323 | 3,170 | 304,767 | 44,109 |  | 304,767 | 44,109 |
| 02/15/98 | 1022.0 | 0.0448 | 22,826 | 20,425 | 327,593 | 50,963 |  | 327,593 | 50,963 |
| 02/16/98 | 2509.0 | 0.0440 | 57,012 | 27,091 | 384,605 | 62,836 |  | 384,605 | 62,836 |
| 02/17/98 | 227.0 | 0.0435 | 5,217 | 31,188 | 389,821 | 70,614 |  | 389,821 | 70,614 |
| 02/18/98 | 62.0 | 0.0424 | 1,461 | 2,602 | 391,283 | 70,789 |  | 391,283 | 70,789 |
| 02/19/98 | 273.0 | 0.0443 | 6,166 | 3,450 | 397,449 | 71,435 |  | 397,449 | 71,435 |
| 02/20/98 | 352.0 | 0.0439 | 8,013 | 1,711 | 405,462 | 72,183 |  | 405,462 | 72,183 |
| 02/21/98 | 393.0 | 0.0457 | 8,601 | 1,410 | 414,063 | 73,019 |  | 414,063 | 73,019 |
| 02/22/98 | 316.0 | 0.0447 | 7,070 | 3,153 | 421,132 | 73,753 |  | 421,132 | 73,753 |
| 02/23/98 | 128.5 | 0.0449 | 2,859 | 2,139 | 423,992 | 74,058 |  | 423,992 | 74,058 |
| 02/24/98 | 191.0 | 0.0448 | 4,266 | 956 | 428,258 | 74,472 |  | 428,258 | 74,472 |
| 02/25/98 | 188.0 | 0.0448 | 4,201 | 669 | 432,458 | 74,879 |  | 432,458 | 74,879 |
| 02/26/98 | 159.0 | 0.0434 | 3,662 | 651 | 436,120 | 75,223 |  | 436,120 | 75,223 |
| 02/27/98 | 149.0 | 0.0432 | 3,447 | 457 | 439,567 | 75,546 |  | 439,567 | 75,546 |
| 02/28/98 | 162.0 | 0.0423 | 3,827 | 931 | 443,394 | 75,901 |  | 443,394 | 75,901 |
| 03/01/98 | 97.0 | 0.0421 | 2,302 | 820 | 445,697 | 76,116 |  | 445,697 | 76,116 |
| 03/02/98 | 123.0 | 0.0432 | 2,850 | 666 | 448,546 | 76,386 |  | 448,546 | 76,386 |
| 03/03/98 | 74.0 | 0.0425 | 1,741 | 653 | 450,287 | 76,550 |  | 450,287 | 76,550 |
| 03/04/98 | 81.2 | 0.0425 | 1,909 | 458 | 452,196 | 76,729 |  | 452,196 | 76,729 |
| 03/05/98 | 49.0 | 0.0382 | 1,283 | 483 | 453,479 | 76,835 |  | 453,479 | 76,835 |
| 03/06/98 | 52.0 | 0.0359 | 1,447 | 1,471 | 454,926 | 76,958 |  | 454,926 | 76,958 |
| 03/07/98 | 142.0 | 0.0412 | 3,451 | 1,219 | 458,377 | 77,276 |  | 458,377 | 77,283 |
| 03/08/98 | 124.0 | 0.0328 | 3,786 | 4,742 | 462,162 | 77,669 | Parr | 3,786 | 5,986 |
| 03/09/98 | 402.0 | 0.0330 | 12,198 | 4,922 | 474,361 | 78,628 |  | 15,984 | 7,805 |
| 03/10/98 | 394.0 | 0.0326 | 12,086 | 2,987 | 486,446 | 79,475 |  | 28,070 | 8,568 |
| 03/11/98 | 242.0 | 0.0331 | 7,317 | 2,466 | 493,764 | 80,010 |  | 35,387 | 9,127 |
| 03/12/98 | 352.0 | 0.0326 | 10,804 | 4,496 | 504,568 | 80,854 |  | 46,191 | 10,515 |
| 03/13/98 | 68.0 | 0.0304 | 2,234 | 5,284 | 506,802 | 81,160 |  | 48,425 | 11,845 |
| 03/14/98 | 77.0 | 0.0298 | 2,583 | 301 | 509,384 | 81,310 |  | 51,008 | 11,940 |
| 03/15/98 | 78.0 | 0.0333 | 2,342 | 574 | 511,727 | 81,475 |  | 53,350 | 12,047 |
| 03/16/98 | 108.0 | 0.0320 | 3,374 | 2,664 | 515,100 | 81,740 |  | 56,724 | 12,469 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 03/17/98 | 238.0 | 0.0308 | 7,721 | 3,615 | 522,821 | 82,296 |  | 64,445 | 13,278 |
| 03/18/98 | 20.0 | 0.0323 | 619 | 3,803 | 523,441 | 82,425 |  | 65,064 | 13,838 |
| 03/19/98 | 29.0 | 0.0304 | 953 | 404 | 524,394 | 82,484 |  | 66,017 | 13,884 |
| 03/20/98 | 43.9 | 0.0268 | 1,634 | 507 | 526,028 | 82,564 |  | 67,651 | 13,955 |
| 03/21/98 | 55.0 | 0.0325 | 1,692 | 781 | 527,719 | 82,683 |  | 69,343 | 14,052 |
| 03/22/98 | 7.3 | 0.0434 | 169 | 654 | 527,888 | 82,702 |  | 69,511 | 14,076 |
| 03/23/98 | 58.5 | 0.0249 | 2,347 | 1,154 | 530,235 | 82,804 |  | 71,859 | 14,212 |
| 03/24/98 | 54.0 | 0.0249 | 2,172 | 295 | 532,407 | 82,891 |  | 74,030 | 14,299 |
| 03/25/98 | 48.0 | 0.0234 | 2,053 | 11,134 | 534,460 | 83,703 |  | 76,083 | 18,186 |
| 03/26/98 | 504.0 | 0.0267 | 18,887 | 8,684 | 553,346 | 85,038 |  | 94,970 | 20,724 |
| 03/27/98 | 279.2 | 0.0243 | 11,500 | 8,668 | 564,847 | 85,921 |  | 106,470 | 22,845 |
| 03/28/98 | 85.0 | 0.0226 | 3,755 | 4,738 | 568,602 | 86,174 |  | 110,225 | 23,466 |
| 03/29/98 | 102.3 | 0.0240 | 4,272 | 893 | 572,874 | 86,343 |  | 114,497 | 23,645 |
| 03/30/98 | 123.0 | 0.0231 | 5,315 | 1,299 | 578,189 | 86,542 |  | 119,812 | 23,887 |
| 03/31/98 | 68.4 | 0.0227 | 3,014 | 1,386 | 581,203 | 86,656 |  | 122,826 | 24,050 |
| 04/01/98 | 71.0 | 0.0233 | 3,048 | 361 | 584,251 | 86,769 |  | 125,874 | 24,181 |
| 04/02/98 | 62.0 | 0.0209 | 2,968 | 1,127 | 587,218 | 86,855 |  | 128,842 | 24,333 |
| 04/03/98 | 105.0 | 0.0233 | 4,505 | 3,681 | 591,723 | 87,103 |  | 133,346 | 24,806 |
| 04/04/98 | 227.0 | 0.0230 | 9,887 | 4,422 | 601,610 | 87,576 |  | 143,234 | 25,631 |
| 04/05/98 | 302.0 | 0.0233 | 12,976 | 2,683 | 614,586 | 88,121 |  | 156,209 | 26,376 |
| 04/06/98 | 194.7 | 0.0263 | 7,389 | 2,143 | 621,975 | 88,528 |  | 163,598 | 26,829 |
| 04/07/98 | 254.0 | 0.0235 | 10,796 | 2,699 | 632,771 | 89,018 |  | 174,394 | 27,522 |
| 04/08/98 | 312.0 | 0.0231 | 13,496 | 4,148 | 646,267 | 89,666 |  | 187,890 | 28,558 |
| 04/09/98 | 133.3 | 0.0238 | 5,595 | 5,098 | 651,862 | 90,061 |  | 193,485 | 29,324 |
| 04/10/98 | 80.4 | 0.0236 | 3,404 | 1,347 | 655,266 | 90,223 |  | 196,890 | 29,551 |
| 04/11/98 | 79.0 | 0.0218 | 3,621 | 445 | 658,887 | 90,360 |  | 200,510 | 29,770 |
| 04/12/98 | 71.0 | 0.0209 | 3,391 | 1,459 | 662,278 | 90,489 |  | 203,901 | 30,013 |
| 04/13/98 | 24.0 | 0.0200 | 1,200 | 1,341 | 663,478 | 90,536 |  | 205,101 | 30,118 |
| 04/14/98 | 25.0 | 0.0167 | 1,500 | 541 | 664,977 | 90,560 |  | 206,601 | 30,222 |
| 04/15/98 | 39.0 | 0.0164 | 2,374 | 568 | 667,352 | 90,596 |  | 208,975 | 30,388 |
| 04/16/98 | 27.0 | 0.0183 | 1,479 | 653 | 668,831 | 90,633 |  | 210,454 | 30,493 |
| 04/17/98 | 16.0 | 0.0182 | 879 | 572 | 669,710 | 90,656 |  | 211,333 | 30,558 |
| 04/18/98 | 36.5 | 0.0201 | 1,821 | 1,471 | 671,531 | 90,727 |  | 213,155 | 30,712 |
| 04/19/98 | 74.0 | 0.0180 | 4,119 | 1,552 | 675,651 | 90,835 |  | 217,274 | 31,031 |
| 04/20/98 | 23.0 | 0.0198 | 1,160 | 1,511 | 676,811 | 90,885 |  | 218,434 | 31,146 |
| 04/21/98 | 21.0 | 0.0182 | 1,154 | 222 | 677,965 | 90,914 |  | 219,588 | 31,227 |
| 04/22/98 | 27.0 | 0.0158 | 1,711 | 628 | 679,676 | 90,939 |  | 221,299 | 31,360 |
| 04/23/98 | 39.0 | 0.0178 | 2,190 | 726 | 681,866 | 90,994 |  | 223,489 | 31,521 |
| 04/24/98 | 51.0 | 0.0168 | 3,026 | 664 | 684,892 | 91,055 | Smolt | 3,026 | 469 |
| 04/25/98 | 56.1 | 0.0169 | 3,312 | 618 | 688,204 | 91,125 |  | 6,338 | 991 |
| 04/26/98 | 42.0 | 0.0175 | 2,395 | 535 | 690,598 | 91,185 |  | 8,733 | 1,350 |
| 04/27/98 | 44.0 | 0.0177 | 2,487 | 1,087 | 693,086 | 91,254 |  | 11,220 | 1,944 |
| 04/28/98 | 75.0 | 0.0161 | 4,671 | 1,203 | 697,756 | 91,345 |  | 15,891 | 2,699 |
| 04/29/98 | 67.0 | 0.0175 | 3,830 | 557 | 701,586 | 91,446 |  | 19,721 | 3,134 |
| 04/30/98 | 72.0 | 0.0162 | 4,443 | 1,291 | 706,029 | 91,546 |  | 24,164 | 3,867 |
| 05/01/98 | 101.0 | 0.0168 | 6,006 | 1,558 | 712,035 | 91,705 |  | 30,169 | 4,790 |
| 05/02/98 | 57.0 | 0.0168 | 3,388 | 1,798 | 715,423 | 91,809 |  | 33,557 | 5,486 |
| 05/03/98 | 45.0 | 0.0166 | 2,718 | 669 | 718,140 | 91,880 |  | 36,275 | 5,840 |
| 05/04/98 | 39.0 | 0.0145 | 2,685 | 2,402 | 720,825 | 91,953 |  | 38,960 | 6,649 |
| 05/05/98 | 102.0 | 0.0158 | 6,453 | 2,200 | 727,278 | 92,127 |  | 45,413 | 7,712 |
| 05/06/98 | 65.0 | 0.0158 | 4,111 | 2,798 | 731,389 | 92,269 |  | 49,524 | 8,667 |
| 05/07/98 | 15.0 | 0.0162 | 925 | 2,081 | 732,314 | 92,318 |  | 50,449 | 9,018 |
| 05/08/98 | 0.0 | 0.0159 | 0 | 1,671 | 732,314 | 92,333 |  | 50,449 | 9,171 |
| 05/09/98 | 52.2 | 0.0156 | 3,348 | 3,061 | 735,662 | 92,465 |  | 53,797 | 10,034 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 05/10/98 | 95.0 | 0.0153 | 6,222 | 1,768 | 741,884 | 92,630 |  | 60,019 | 10,883 |
| 05/11/98 | 88.0 | 0.0150 | 5,880 | 956 | 747,764 | 92,776 |  | 65,899 | 11,629 |
| 05/12/98 | 94.0 | 0.0154 | 6,104 | 1,951 | 753,869 | 92,964 |  | 72,003 | 12,520 |
| 05/13/98 | 45.0 | 0.0142 | 3,175 | 3,113 | 757,043 | 93,090 |  | 75,178 | 13,320 |
| 05/14/98 | 133.0 | 0.0146 | 9,083 | 4,259 | 766,127 | 93,428 |  | 84,261 | 15,085 |
| 05/15/98 | 158.0 | 0.0153 | 10,355 | 1,857 | 776,481 | 93,771 |  | 94,616 | 16,451 |
| 05/16/98 | 132.0 | 0.0153 | 8,604 | 1,962 | 785,085 | 94,085 |  | 103,220 | 17,647 |
| 05/17/98 | 113.0 | 0.0161 | 7,023 | 1,673 | 792,108 | 94,373 |  | 110,243 | 18,589 |
| 05/18/98 | 89.0 | 0.0150 | 5,943 | 5,041 | 798,051 | 94,724 |  | 116,186 | 20,037 |
| 05/19/98 | 229.4 | 0.0156 | 14,670 | 5,736 | 812,721 | 95,483 |  | 130,856 | 22,594 |
| 05/20/98 | 80.0 | 0.0151 | 5,305 | 6,667 | 818,026 | 95,935 |  | 136,161 | 24,231 |
| 05/21/98 | 37.0 | 0.0151 | 2,458 | 1,462 | 820,484 | 96,051 |  | 138,619 | 24,594 |
| 05/22/98 | 59.0 | 0.0150 | 3,929 | 2,796 | 824,413 | 96,261 |  | 142,547 | 25,261 |
| 05/23/98 | 117.3 | 0.0152 | 7,722 | 2,600 | 832,135 | 96,638 |  | 150,269 | 26,378 |
| 05/24/98 | 53.0 | 0.0148 | 3,573 | 2,811 | 835,707 | 96,841 |  | 153,842 | 27,003 |
| 05/25/98 | 40.0 | 0.0155 | 2,579 | 1,066 | 838,287 | 96,970 |  | 156,421 | 27,356 |
| 05/26/98 | 71.0 | 0.0152 | 4,677 | 2,268 | 842,964 | 97,218 |  | 161,099 | 28,061 |
| 05/27/98 | 5.0 | 0.0156 | 320 | 2,093 | 843,284 | 97,257 |  | 161,419 | 28,181 |
| 05/28/98 | 41.0 | 0.0145 | 2,824 | 1,707 | 846,108 | 97,404 |  | 164,242 | 28,622 |
| 05/29/98 | 51.0 | 0.0137 | 3,716 | 787 | 849,824 | 97,576 |  | 167,958 | 29,175 |
| 05/30/98 | 39.0 | 0.0138 | 2,817 | 895 | 852,641 | 97,712 |  | 170,775 | 29,600 |
| 05/31/98 | 29.8 | 0.0130 | 2,294 | 1,354 | 854,935 | 97,825 |  | 173,069 | 29,988 |
| 06/01/98 | 6.0 | 0.0122 | 492 | 1,932 | 855,426 | 97,866 |  | 173,561 | 30,132 |
| 06/02/98 | 54.0 | 0.0144 | 3,746 | 1,753 | 859,173 | 98,068 |  | 177,307 | 30,711 |
| 06/03/98 | 29.0 | 0.0144 | 2,009 | 1,876 | 861,181 | 98,188 |  | 179,316 | 31,053 |
| 06/04/98 | 0.0 | 0.0138 | 0 | 2,738 | 861,181 | 98,226 |  | 179,316 | 31,173 |
| 06/05/98 | 76.0 | 0.0132 | 5,761 | 3,331 | 866,942 | 98,560 |  | 185,077 | 32,226 |
| 06/06/98 | 5.1 | 0.0112 | 454 | 3,295 | 867,396 | 98,636 |  | 185,531 | 32,475 |
| 06/07/98 | 17.9 | 0.0118 | 1,515 | 822 | 868,911 | 98,710 |  | 187,046 | 32,743 |
| 06/08/98 | 0.0 | 0.0124 | 0 | 2,703 | 868,911 | 98,747 |  | 187,046 | 32,855 |
| 06/09/98 | 66.0 | 0.0130 | 5,064 | 2,997 | 873,976 | 99,046 |  | 192,110 | 33,771 |
| 06/10/98 | 1.0 | 0.0124 | 81 | 2,712 | 874,056 | 99,087 |  | 192,191 | 33,893 |
| 06/11/98 | 15.0 | 0.0130 | 1,157 | 773 | 875,214 | 99,150 |  | 193,348 | 34,084 |
| 06/12/98 | 19.9 | 0.0131 | 1,521 | 465 | 876,735 | 99,230 |  | 194,870 | 34,324 |
| 06/13/98 | 25.0 | 0.0128 | 1,955 | 688 | 878,690 | 99,334 |  | 196,824 | 34,642 |
| 06/14/98 | 10.0 | 0.0126 | 795 | 994 | 879,485 | 99,380 |  | 197,619 | 34,785 |
| 06/15/98 | 0.0 | 0.0124 | 0 | 398 | 879,485 | 99,381 |  | 197,619 | 34,788 |
| 06/16/98 | 6.0 | 0.0123 | 490 | 274 | 879,974 | 99,407 |  | 198,109 | 34,870 |
| 06/17/98 | 1.0 | 0.0102 | 98 | 254 | 880,073 | 99,412 |  | 198,207 | 34,890 |
| 06/18/98 | 2.0 | 0.0110 | 182 | 97 | 880,255 | 99,421 |  | 198,389 | 34,924 |
| 06/19/98 | 0.0 | 0.0113 | 0 | 95 | 880,255 | 99,421 |  | 198,389 | 34,924 |
| 06/20/98 | 1.7 | 0.0115 | 152 | 93 | 880,406 | 99,428 |  | 198,541 | 34,950 |
| 06/21/98 | 1.8 | 0.0118 | 154 | 48 | 880,560 | 99,436 |  | 198,695 | 34,977 |
| 06/22/98 | 1.0 | 0.0121 | 83 | 46 | 880,643 | 99,441 |  | 198,777 | 34,991 |
| 06/23/98 | 2.0 | 0.0135 | 148 | 77 | 880,791 | 99,449 |  | 198,925 | 35,013 |
| 06/24/98 | 3.0 | 0.0102 | 293 | 159 | 881,083 | 99,463 |  | 199,218 | 35,070 |
| 06/25/98 | 0.0 | 0.0099 | 0 | 147 | 881,083 | 99,463 |  | 199,218 | 35,070 |
| 06/26/98 | 1.6 | 0.0096 | 170 | 98 | 881,253 | 99,471 |  | 199,388 | 35,105 |
| 06/27/98 | 1.4 | 0.0093 | 150 | 42 | 881,403 | 99,477 |  | 199,538 | 35,136 |
| 06/28/98 | 1.2 | 0.0090 | 137 | 40 | 881,540 | 99,484 |  | 199,674 | 35,165 |
| 06/29/98 | 1.0 | 0.0087 | 115 | 65 | 881,655 | 99,489 |  | 199,789 | 35,190 |
| 06/30/98 | 2.0 | 0.0093 | 215 | 116 | 881,869 | 99,498 |  | 200,004 | 35,234 |
| 07/01/98 | 0.0 | 0.0099 | 0 | 113 | 881,869 | 99,498 |  | 200,004 | 35,234 |
| 07/02/98 | 0.0 | 0.0106 | 0 | 106 | 881,869 | 99,498 |  | 200,004 | 35,235 |

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| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 07/03/98 | 2.0 | 0.0113 | 176 | 94 | 882,045 | 99,507 |  | 200,180 | 35,266 |
| 07/04/98 | 0.7 | 0.0113 | 66 | 70 | 882,111 | 99,511 |  | 200,246 | 35,278 |
| 07/05/98 | 0.5 | 0.0113 | 46 | 21 | 882,158 | 99,513 |  | 200,292 | 35,286 |
| 07/06/98 | 0.3 | 0.0113 | 28 | 24 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/07/98 | 0.0 | 0.0113 | 0 | 16 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/08/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/09/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/10/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/11/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/12/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/13/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/14/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/15/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |
| 07/16/98 | 0.0 | 0.0113 | 0 | 0 | 882,186 | 99,515 |  | 200,320 | 35,291 |

Appendix A.6.d. 1999 Outmigration index estimates based on efficiency predictor that excluded turbidity > 10

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 01/18/99 | 13.0 | 0.0260 | 501 | 94 | 501 | 94 | Fry | 501 | 94 |
| 01/19/99 | 16.0 | 0.0260 | 616 | 2,160 | 1,117 | 2,163 |  | 1,117 | 2,163 |
| 01/20/99 | 112.0 | 0.0260 | 4,313 | 34,183 | 5,430 | 34,252 |  | 5,430 | 34,252 |
| 01/21/99 | 1606.0 | 0.0260 | 61,851 | 36,538 | 67,281 | 50,143 |  | 67,281 | 50,143 |
| 01/22/99 | 1849.0 | 0.0260 | 71,209 | 21,872 | 138,490 | 55,496 |  | 138,490 | 55,496 |
| 01/23/99 | 812.0 | 0.0260 | 31,272 | 32,355 | 169,762 | 64,849 |  | 169,762 | 64,849 |
| 01/24/99 | 185.0 | 0.0260 | 7,125 | 15,477 | 176,887 | 66,834 |  | 176,887 | 66,834 |
| 01/25/99 | 938.0 | 0.0260 | 36,125 | 15,514 | 213,012 | 69,452 |  | 213,012 | 69,452 |
| 01/26/99 | 766.0 | 0.0260 | 29,501 | 4,928 | 242,512 | 70,442 |  | 242,512 | 70,442 |
| 01/27/99 | 746.0 | 0.0260 | 28,730 | 4,372 | 271,242 | 71,469 |  | 271,242 | 71,469 |
| 01/28/99 | 909.0 | 0.0260 | 35,008 | 6,433 | 306,250 | 72,950 |  | 306,250 | 72,950 |
| 01/29/99 | 623.0 | 0.0260 | 23,993 | 7,091 | 330,243 | 74,199 |  | 330,243 | 74,199 |
| 01/30/99 | 591.0 | 0.0260 | 22,761 | 2,276 | 353,004 | 75,148 |  | 353,004 | 75,148 |
| 01/31/99 | 621.0 | 0.0260 | 23,916 | 6,961 | 376,920 | 76,480 |  | 376,920 | 76,480 |
| 02/01/99 | 310.0 | 0.0260 | 11,939 | 11,844 | 388,859 | 77,918 |  | 388,859 | 77,918 |
| 02/02/99 | 925.0 | 0.0260 | 35,624 | 12,410 | 424,483 | 80,479 |  | 424,483 | 80,479 |
| 02/03/99 | 533.0 | 0.0260 | 20,527 | 8,419 | 445,010 | 81,891 |  | 445,010 | 81,891 |
| 02/04/99 | 582.0 | 0.0260 | 22,414 | 2,418 | 467,424 | 83,025 |  | 467,424 | 83,025 |
| 02/05/99 | 586.0 | 0.0260 | 22,568 | 11,840 | 489,993 | 85,000 |  | 489,993 | 85,000 |
| 02/06/99 | 1110.0 | 0.0260 | 42,749 | 11,187 | 532,741 | 87,925 |  | 532,741 | 87,925 |
| 02/07/99 | 723.0 | 0.0260 | 27,844 | 13,684 | 560,586 | 90,485 |  | 560,586 | 90,485 |
| 02/08/99 | 411.0 | 0.0260 | 15,829 | 19,232 | 576,414 | 93,374 |  | 576,414 | 93,374 |
| 02/09/99 | 1390.0 | 0.0260 | 53,532 | 36,988 | 629,946 | 103,187 |  | 629,946 | 103,187 |
| 02/10/99 | 2322.0 | 0.0260 | 89,426 | 19,822 | 719,372 | 109,836 |  | 719,372 | 109,836 |
| 02/11/99 | 1903.0 | 0.0260 | 73,289 | 10,968 | 792,661 | 114,640 |  | 792,661 | 114,640 |
| 02/12/99 | 2232.0 | 0.0260 | 85,960 | 17,386 | 878,621 | 121,172 |  | 878,621 | 121,172 |
| 02/13/99 | 1436.0 | 0.0260 | 55,304 | 22,238 | 933,924 | 126,728 |  | 933,924 | 126,728 |
| 02/14/99 | 1143.0 | 0.0260 | 44,020 | 8,697 | 977,944 | 129,933 |  | 977,944 | 129,933 |
| 02/15/99 | 1522.0 | 0.0260 | 58,616 | 27,606 | 1,036,560 | 136,698 |  | 1,036,560 | 136,698 |


| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 02/16/99 | 156.0 | 0.0260 | 6,008 | 26,277 | 1,042,568 | 139,606 |  | 1,042,568 | 139,606 |
| 02/17/99 | 743.0 | 0.0260 | 28,615 | 11,866 | 1,071,182 | 142,031 |  | 1,071,182 | 142,031 |
| 02/18/99 | 567.0 | 0.0260 | 21,837 | 8,175 | 1,093,019 | 143,752 |  | 1,093,019 | 143,752 |
| 02/19/99 | 978.0 | 0.0260 | 37,665 | 9,086 | 1,130,684 | 146,613 |  | 1,130,684 | 146,613 |
| 02/20/99 | 897.0 | 0.0260 | 34,546 | 4,145 | 1,165,229 | 149,071 |  | 1,165,229 | 149,071 |
| 02/21/99 | 1027.0 | 0.0260 | 39,552 | 5,909 | 1,204,782 | 151,969 |  | 1,204,782 | 151,969 |
| 02/22/99 | 1134.0 | 0.0260 | 43,673 | 7,716 | 1,248,455 | 155,274 |  | 1,248,455 | 155,274 |
| 02/23/99 | 802.0 | 0.0260 | 30,887 | 8,154 | 1,279,342 | 157,725 |  | 1,279,342 | 157,725 |
| 02/24/99 | 780.0 | 0.0260 | 30,040 | 7,244 | 1,309,381 | 160,088 |  | 1,309,381 | 160,088 |
| 02/25/99 | 491.0 | 0.0260 | 18,910 | 7,876 | 1,328,291 | 161,679 |  | 1,328,291 | 161,679 |
| 02/26/99 | 396.0 | 0.0260 | 15,251 | 3,058 | 1,343,542 | 162,842 |  | 1,343,542 | 162,842 |
| 02/27/99 | 354.0 | 0.0260 | 13,633 | 1,837 | 1,357,175 | 163,871 |  | 1,357,175 | 163,871 |
| 02/28/99 | 329.0 | 0.0260 | 12,671 | 2,653 | 1,369,846 | 164,843 |  | 1,369,846 | 164,843 |
| 03/01/99 | 237.0 | 0.0260 | 9,127 | 2,083 | 1,378,973 | 165,543 |  | 1,378,973 | 165,543 |
| 03/02/99 | 314.0 | 0.0260 | 12,093 | 3,461 | 1,391,066 | 166,492 |  | 1,391,066 | 166,492 |
| 03/03/99 | 144.0 | 0.0260 | 5,546 | 4,290 | 1,396,612 | 166,968 |  | 1,396,612 | 166,968 |
| 03/04/99 | 105.2 | 0.0260 | 4,053 | 2,984 | 1,400,665 | 167,302 |  | 1,400,665 | 167,302 |
| 03/05/99 | 254.0 | 0.0260 | 9,782 | 4,272 | 1,410,447 | 168,099 |  | 1,410,447 | 168,099 |
| 03/06/99 | 317.0 | 0.0260 | 12,208 | 4,889 | 1,422,655 | 169,097 |  | 1,422,655 | 169,097 |
| 03/07/99 | 78.0 | 0.0260 | 3,004 | 5,191 | 1,425,659 | 169,406 |  | 1,425,659 | 169,406 |
| 03/08/99 | 88.0 | 0.0260 | 3,389 | 436 | 1,429,048 | 169,666 |  | 1,429,048 | 169,666 |
| 03/09/99 | 93.0 | 0.0260 | 3,582 | 542 | 1,432,630 | 169,941 |  | 1,432,630 | 169,941 |
| 03/10/99 | 109.0 | 0.0260 | 4,198 | 1,462 | 1,436,828 | 170,268 |  | 1,436,828 | 170,268 |
| 03/11/99 | 39.0 | 0.0260 | 1,502 | 1,556 | 1,438,330 | 170,390 |  | 1,438,330 | 170,390 |
| 03/12/99 | 39.0 | 0.0260 | 1,502 | 406 | 1,439,832 | 170,506 |  | 1,439,832 | 170,506 |
| 03/13/99 | 56.2 | 0.0260 | 2,163 | 432 | 1,441,995 | 170,672 |  | 1,441,995 | 170,672 |
| 03/14/99 | 39.0 | 0.0260 | 1,502 | 417 | 1,443,497 | 170,788 |  | 1,443,497 | 170,788 |
| 03/15/99 | 38.0 | 0.0260 | 1,464 | 141 | 1,444,960 | 170,901 |  | 1,444,960 | 170,901 |
| 03/16/99 | 38.4 | 0.0260 | 1,478 | 141 | 1,446,438 | 171,014 | Parr | 1,478 | 141 |
| 03/17/99 | 38.4 | 0.0260 | 1,480 | 456 | 1,447,918 | 171,128 |  | 2,958 | 517 |
| 03/18/99 | 58.0 | 0.0260 | 2,234 | 505 | 1,450,152 | 171,301 |  | 5,191 | 802 |
| 03/19/99 | 36.4 | 0.0260 | 1,401 | 472 | 1,451,552 | 171,409 |  | 6,592 | 999 |
| 03/20/99 | 39.0 | 0.0260 | 1,502 | 263 | 1,453,054 | 171,525 |  | 8,094 | 1,116 |
| 03/21/99 | 28.0 | 0.0260 | 1,078 | 363 | 1,454,132 | 171,608 |  | 9,172 | 1,240 |
| 03/22/99 | 21.0 | 0.0260 | 809 | 211 | 1,454,941 | 171,671 |  | 9,981 | 1,310 |
| 03/23/99 | 18.0 | 0.0260 | 693 | 201 | 1,455,634 | 171,724 |  | 10,674 | 1,372 |
| 03/24/99 | 27.7 | 0.0260 | 1,066 | 288 | 1,456,700 | 171,806 |  | 11,739 | 1,474 |
| 03/25/99 | 14.0 | 0.0260 | 539 | 483 | 1,457,239 | 171,848 |  | 12,279 | 1,587 |
| 03/26/99 | 39.0 | 0.0260 | 1,502 | 913 | 1,458,741 | 171,967 |  | 13,781 | 1,920 |
| 03/27/99 | 61.0 | 0.0260 | 2,349 | 502 | 1,461,090 | 172,148 |  | 16,130 | 2,128 |
| 03/28/99 | 57.0 | 0.0260 | 2,195 | 259 | 1,463,285 | 172,318 |  | 18,325 | 2,289 |
| 03/29/99 | 53.0 | 0.0260 | 2,041 | 802 | 1,465,326 | 172,477 |  | 20,366 | 2,562 |
| 03/30/99 | 20.0 | 0.0260 | 770 | 808 | 1,466,097 | 172,538 |  | 21,136 | 2,739 |
| 03/31/99 | 14.0 | 0.0260 | 539 | 1,026 | 1,466,636 | 172,583 |  | 21,676 | 2,960 |
| 04/01/99 | 63.0 | 0.0260 | 2,426 | 1,289 | 1,469,062 | 172,775 |  | 24,102 | 3,373 |
| 04/02/99 | 77.0 | 0.0260 | 2,966 | 641 | 1,472,027 | 173,005 |  | 27,067 | 3,618 |
| 04/03/99 | 47.0 | 0.0260 | 1,810 | 1,201 | 1,473,838 | 173,149 |  | 28,877 | 3,927 |
| 04/04/99 | 109.0 | 0.0260 | 4,198 | 1,702 | 1,478,035 | 173,482 |  | 33,075 | 4,530 |
| 04/05/99 | 26.0 | 0.0260 | 1,001 | 1,749 | 1,479,037 | 173,568 |  | 34,077 | 4,917 |
| 04/06/99 | 35.0 | 0.0260 | 1,348 | 272 | 1,480,385 | 173,673 |  | 35,424 | 5,009 |
| 04/07/99 | 23.0 | 0.0260 | 886 | 468 | 1,481,270 | 173,742 |  | 36,310 | 5,087 |
| 04/08/99 | 11.0 | 0.0260 | 424 | 384 | 1,481,694 | 173,775 |  | 36,734 | 5,129 |
| 04/09/99 | 30.8 | 0.0260 | 1,185 | 396 | 1,482,879 | 173,867 |  | 37,919 | 5,220 |
| 04/10/99 | 21.0 | 0.0260 | 809 | 465 | 1,483,688 | 173,931 |  | 38,728 | 5,294 |


| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 04/11/99 | 7.0 | 0.0260 | 270 | 500 | 1,483,958 | 173,952 |  | 38,998 | 5,335 |
| 04/12/99 | 33.0 | 0.0260 | 1,271 | 515 | 1,485,229 | 174,051 |  | 40,268 | 5,443 |
| 04/13/99 | 18.0 | 0.0260 | 693 | 405 | 1,485,922 | 174,106 |  | 40,962 | 5,505 |
| 04/14/99 | 38.0 | 0.0260 | 1,464 | 570 | 1,487,385 | 174,220 |  | 42,425 | 5,632 |
| 04/15/99 | 46.0 | 0.0260 | 1,772 | 836 | 1,489,157 | 174,359 |  | 44,197 | 5,812 |
| 04/16/99 | 78.3 | 0.0260 | 3,017 | 693 | 1,492,174 | 174,595 |  | 47,214 | 6,057 |
| 04/17/99 | 57.0 | 0.0260 | 2,195 | 576 | 1,494,369 | 174,766 |  | 49,409 | 6,237 |
| 04/18/99 | 52.0 | 0.0260 | 2,003 | 236 | 1,496,372 | 174,922 |  | 51,412 | 6,384 |
| 04/19/99 | 59.0 | 0.0260 | 2,272 | 464 | 1,498,644 | 175,099 |  | 53,684 | 6,564 |
| 04/20/99 | 38.0 | 0.0260 | 1,464 | 639 | 1,500,108 | 175,214 |  | 55,147 | 6,703 |
| 04/21/99 | 27.0 | 0.0260 | 1,040 | 253 | 1,501,147 | 175,295 |  | 56,187 | 6,785 |
| 04/22/99 | 28.0 | 0.0260 | 1,078 | 170 | 1,502,226 | 175,379 |  | 57,266 | 6,867 |
| 04/23/99 | 33.6 | 0.0260 | 1,293 | 435 | 1,503,519 | 175,480 |  | 58,559 | 6,978 |
| 04/24/99 | 49.0 | 0.0260 | 1,887 | 347 | 1,505,406 | 175,627 |  | 60,446 | 7,129 |
| 04/25/99 | 40.0 | 0.0260 | 1,541 | 227 | 1,506,946 | 175,747 |  | 61,986 | 7,250 |
| 04/26/99 | 45.0 | 0.0260 | 1,733 | 239 | 1,508,680 | 175,882 |  | 63,719 | 7,388 |
| 04/27/99 | 36.0 | 0.0260 | 1,386 | 370 | 1,510,066 | 175,991 |  | 65,106 | 7,505 |
| 04/28/99 | 54.0 | 0.0260 | 2,080 | 428 | 1,512,146 | 176,153 |  | 67,185 | 7,679 |
| 04/29/99 | 52.2 | 0.0260 | 2,009 | 445 | 1,514,154 | 176,310 |  | 69,194 | 7,849 |
| 04/30/99 | 35.0 | 0.0260 | 1,348 | 491 | 1,515,502 | 176,416 |  | 70,542 | 7,972 |
| 05/01/99 | 59.0 | 0.0260 | 2,272 | 1,118 | 1,517,774 | 176,597 |  | 72,814 | 8,229 |
| 05/02/99 | 92.0 | 0.0260 | 3,543 | 1,165 | 1,521,317 | 176,877 |  | 76,357 | 8,588 |
| 05/03/99 | 34.0 | 0.0260 | 1,309 | 1,543 | 1,522,627 | 176,986 |  | 77,667 | 8,829 |
| 05/04/99 | 15.0 | 0.0260 | 578 | 1,013 | 1,523,204 | 177,034 |  | 78,244 | 8,933 |
| 05/05/99 | 67.1 | 0.0260 | 2,586 | 2,009 | 1,525,790 | 177,248 |  | 80,830 | 9,355 |
| 05/06/99 | 119.0 | 0.0260 | 4,583 | 1,374 | 1,530,373 | 177,611 |  | 85,413 | 9,804 |
| 05/07/99 | 55.0 | 0.0260 | 2,118 | 1,344 | 1,532,491 | 177,782 |  | 87,531 | 10,061 |
| 05/08/99 | 64.0 | 0.0260 | 2,465 | 347 | 1,534,956 | 177,975 |  | 89,996 | 10,260 |
| 05/09/99 | 68.0 | 0.0260 | 2,619 | 357 | 1,537,575 | 178,181 |  | 92,615 | 10,470 |
| 05/10/99 | 55.0 | 0.0260 | 2,118 | 911 | 1,539,693 | 178,349 | Smolt | 2,118 | 891 |
| 05/11/99 | 23.0 | 0.0260 | 886 | 802 | 1,540,579 | 178,420 |  | 3,004 | 1,212 |
| 05/12/99 | 62.0 | 0.0260 | 2,388 | 788 | 1,542,967 | 178,609 |  | 5,392 | 1,490 |
| 05/13/99 | 47.0 | 0.0260 | 1,810 | 512 | 1,544,777 | 178,752 |  | 7,202 | 1,631 |
| 05/14/99 | 72.0 | 0.0260 | 2,773 | 697 | 1,547,550 | 178,971 |  | 9,975 | 1,873 |
| 05/15/99 | 40.0 | 0.0260 | 1,541 | 645 | 1,549,090 | 179,093 |  | 11,515 | 2,050 |
| 05/16/99 | 50.0 | 0.0260 | 1,926 | 280 | 1,551,016 | 179,244 |  | 13,441 | 2,165 |
| 05/17/99 | 49.0 | 0.0260 | 1,887 | 1,023 | 1,552,903 | 179,396 |  | 15,328 | 2,489 |
| 05/18/99 | 95.0 | 0.0260 | 3,659 | 1,211 | 1,556,562 | 179,687 |  | 18,987 | 2,946 |
| 05/19/99 | 38.0 | 0.0260 | 1,464 | 1,248 | 1,558,025 | 179,807 |  | 20,450 | 3,277 |
| 05/20/99 | 40.0 | 0.0260 | 1,541 | 1,932 | 1,559,566 | 179,938 |  | 21,991 | 3,879 |
| 05/21/99 | 126.0 | 0.0260 | 4,853 | 1,915 | 1,564,418 | 180,330 |  | 26,843 | 4,544 |
| 05/22/99 | 121.7 | 0.0260 | 4,687 | 1,173 | 1,569,106 | 180,703 |  | 31,531 | 4,931 |
| 05/23/99 | 75.0 | 0.0260 | 2,888 | 1,193 | 1,571,994 | 180,935 |  | 34,419 | 5,233 |
| 05/24/99 | 65.0 | 0.0260 | 2,503 | 1,298 | 1,574,497 | 181,137 |  | 36,922 | 5,535 |
| 05/25/99 | 13.0 | 0.0260 | 501 | 2,013 | 1,574,998 | 181,187 |  | 37,423 | 5,918 |
| 05/26/99 | 118.0 | 0.0260 | 4,545 | 2,066 | 1,579,542 | 181,558 |  | 41,967 | 6,510 |
| 05/27/99 | 73.0 | 0.0260 | 2,811 | 1,765 | 1,582,354 | 181,788 |  | 44,779 | 6,902 |
| 05/28/99 | 27.0 | 0.0260 | 1,040 | 893 | 1,583,394 | 181,873 |  | 45,819 | 7,020 |
| 05/29/99 | 54.7 | 0.0260 | 2,107 | 624 | 1,585,501 | 182,040 |  | 47,926 | 7,172 |
| 05/30/99 | 52.6 | 0.0260 | 2,024 | 220 | 1,587,525 | 182,201 |  | 49,950 | 7,297 |
| 05/31/99 | 49.3 | 0.0260 | 1,897 | 693 | 1,589,422 | 182,352 |  | 51,847 | 7,446 |
| 06/01/99 | 81.0 | 0.0260 | 3,120 | 1,082 | 1,592,541 | 182,602 |  | 54,966 | 7,717 |
| 06/02/99 | 27.0 | 0.0260 | 1,040 | 1,046 | 1,593,581 | 182,688 |  | 56,006 | 7,854 |
| 06/03/99 | 59.0 | 0.0260 | 2,272 | 703 | 1,595,853 | 182,869 |  | 58,278 | 8,031 |

S.P. Cramer \& Associates, Inc. 1999 Caswell Report September 2000

| Date | Count | Efficiency | Outmigration |  |  |  | Life-Stage Cohort | Cohort Cumulative Outmigration |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Daily |  | Cumulative |  |  |  |  |
|  |  |  | Estimate | S.E. | Estimate | S.E. |  | Estimate | S.E. |
| 06/04/99 | 55.0 | 0.0260 | 2,118 | 783 | 1,597,972 | 183,038 |  | 60,397 | 8,207 |
| 06/05/99 | 23.0 | 0.0260 | 886 | 843 | 1,598,857 | 183,111 |  | 61,282 | 8,308 |
| 06/06/99 | 13.2 | 0.0260 | 507 | 346 | 1,599,364 | 183,151 |  | 61,789 | 8,350 |
| 06/07/99 | 31.0 | 0.0260 | 1,194 | 401 | 1,600,558 | 183,246 |  | 62,983 | 8,439 |
| 06/08/99 | 30.0 | 0.0260 | 1,155 | 136 | 1,601,713 | 183,338 |  | 64,138 | 8,518 |
| 06/09/99 | 34.0 | 0.0260 | 1,309 | 343 | 1,603,023 | 183,442 |  | 65,448 | 8,614 |
| 06/10/99 | 18.0 | 0.0260 | 693 | 321 | 1,603,716 | 183,498 |  | 66,141 | 8,668 |
| 06/11/99 | 29.0 | 0.0260 | 1,117 | 236 | 1,604,833 | 183,586 |  | 67,258 | 8,748 |
| 06/12/99 | 23.7 | 0.0260 | 911 | 203 | 1,605,744 | 183,659 |  | 68,169 | 8,814 |
| 06/13/99 | 19.5 | 0.0260 | 750 | 349 | 1,606,493 | 183,719 |  | 68,919 | 8,873 |
| 06/14/99 | 6.6 | 0.0260 | 253 | 286 | 1,606,747 | 183,739 |  | 69,172 | 8,895 |
| 06/15/99 | 6.6 | 0.0260 | 253 | 122 | 1,607,000 | 183,759 |  | 69,425 | 8,914 |
| 06/16/99 | 12.0 | 0.0260 | 462 | 170 | 1,607,462 | 183,796 |  | 69,887 | 8,948 |
| 06/17/99 | 15.0 | 0.0260 | 578 | 164 | 1,608,040 | 183,842 |  | 70,465 | 8,991 |
| 06/18/99 | 7.0 | 0.0260 | 270 | 179 | 1,608,310 | 183,863 |  | 70,735 | 9,012 |
| 06/19/99 | 7.0 | 0.0260 | 268 | 34 | 1,608,578 | 183,885 |  | 71,003 | 9,031 |
| 06/20/99 | 6.0 | 0.0260 | 229 | 103 | 1,608,807 | 183,903 |  | 71,232 | 9,048 |
| 06/21/99 | 2.0 | 0.0260 | 77 | 212 | 1,608,884 | 183,909 |  | 71,309 | 9,056 |
| 06/22/99 | 12.9 | 0.0260 | 498 | 217 | 1,609,382 | 183,949 |  | 71,807 | 9,094 |
| 06/23/99 | 6.0 | 0.0260 | 231 | 155 | 1,609,613 | 183,967 |  | 72,038 | 9,112 |
| 06/24/99 | 6.0 | 0.0260 | 231 | 31 | 1,609,844 | 183,986 |  | 72,269 | 9,129 |
| 06/25/99 | 5.0 | 0.0260 | 193 | 43 | 1,610,037 | 184,001 |  | 72,462 | 9,142 |
| 06/26/99 | 4.0 | 0.0260 | 154 | 31 | 1,610,191 | 184,013 |  | 72,616 | 9,154 |
| 06/27/99 | 3.6 | 0.0260 | 140 | 55 | 1,610,330 | 184,024 |  | 72,755 | 9,164 |
| 06/28/99 | 1.4 | 0.0260 | 55 | 54 | 1,610,386 | 184,029 |  | 72,811 | 9,168 |
| 06/29/99 | 1.0 | 0.0260 | 39 | 62 | 1,610,424 | 184,032 |  | 72,849 | 9,171 |
| 06/30/99 | 4.0 | 0.0260 | 154 | 83 | 1,610,578 | 184,044 |  | 73,003 | 9,182 |

Appendix B. Doug Neeley's response to statistical reviews by Skalski and McDonald and Howlin.

## IntSTATS

# International Statistical Training and Technical Services 

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S. P. Cramer, Doug Demko, Andrea Phillips

From: Doug Neeley
Subject: Response to reviews of Appendix A. in Outmigrant Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site 1999

I am responding to the reviews of John R. Skalski (March 21, 2000) and Lyman L. McDonald and Shay Howlin (April 28, 2000) to the model used to estimate trapping efficiency and smolt outmigration at Caswell State Park. Their reviews seem to based on S. P. Cramer and Associations report on the 1997 outmigration ${ }^{1}$ and on a draft report of the 1999 outmigration. Below, I respond to the reviews separately as they apply to my contribution (i.e., Appendix A). However, the reviewers may wish to read all of my comments because there are some related responses even though their reviews centered on different issues.

## Response to Skalski's review.

While I was not responsible for the release schedule, it is important that the reader of the review be aware of the limitations imposed on the field team. As I understand it, the California

[^1]Department of Fish and Game (CDFG) imposed the restriction that no more than one release of river-run fish could be made a week, and throughout a large portion of the outmigration, river-run fish were all that were available. More releases should have been made, but I believe that during periods of low recovery, the field team felt that there were too few fish available to make a release. They probably did not realize that the logistic regression effectively weights the efficiencies by the release size and that I could have easily accommodated small releases without their results having a disproportionate effect on the parameter estimates. I should note that a major benefit of Skalski's review was that CDFG has reversed its policy and that more frequent releases can now be made. The decision was probably too late for efficiency estimates during the "fry" portion of the 2000 outmigration.

Now to specific comments regarding the efficiency predictions and count expansions.

## Page 3, Paragraph 3 of Skalski's Review

Skalski criticizes the use of only night-time releases in 1999 and the abundance of hatchery-fish fish (as opposed to river-run fish) in 1998.

Day-time releases: Regarding the lack of 1999 day-time releases. This lack was based on my recommendation which stemmed from the analysis done on the 1998 releases and presented in the report of 1998 outmigration $^{2}$ (which I assume that Skalski did not have for his review). I quote from my appendix. "... The decision was made to drop day releases because the day release efficiency would not have been representative of day passage."

There were three paired day-time and night-time releases made and discussed in the 1998 report. Out of the total of 9172 fish released in the day time, only 17 were recovered and 16 of those 17 were captured at night. For comparable night-time releases, out of a total of 5153 night-time releases, there was a total of 83 fish recovered, all at night. The day-release and night-release efficiency estimates, ignoring time of recovery, respectively were 0.0019 and 0.0162 . Assuming that the night-time efficiency would be independent of whether a fish was released during the day or night, the much lower efficiency associated with the day-time releases and based predominantly on nighttime recoveries implies that many day-time released fish were passing during the day but were not being recovered. While it would be possible to partition the day-time release recoveries into day and night recoveries and use maximum-likelihood techniques to estimate day-time efficiency based on both day-time and night-time release estimates, the estimated day-time efficiency would be near 0 and would be based on only one day-time recovered fish. The result would be that the small day-time catches would be expanded by a huge number based on an estimate derived from one recovered fish.

[^2]The holding of fish may trigger a mass movement of fish when released which may not be representative of river-run fish movement. We hope that radio-telemetry studies can address the movement of fish during the day and night. Some consideration is being given to making paired day-time and night-time releases this year. But until issues are resolved regarding day-time and night-time movement, I would not use 1 unadjusted day-time recovery to estimate day-time efficiency when most of the day-time-released fish are recovered at night. The lack of separate day and night expansions is one reason that I state the outmigration estimates should only be regarded as an index, and I always tried to use the term "index" throughout my appendix. This issue needs a clearer presentation in future reports.

Hatchery releases: The frequency of hatchery and river-run fish is driven by the relative availability of those two sources of fish. With the changes in CDFG policy mentioned above, there should now be greater numbers of river-run-fish releases in 2000. There was a total of four paired releases of river-run ("natural") and hatchery fish in 1998 for the purpose of comparing trapping efficiency estimates. A logistic analysis of variance resulted in no statistically significant differences between the efficiency estimates ( $\mathrm{P}=0.5063$ with a pooled river-run efficiency estimate of 0.0337 from 5100 released fish and a pooled hatchery efficiency estimate of 0.0197 from 11249 released fish.

The report on the 1998 outmigration did present the paired comparisons between hatchery and river-run fish and between day-time and the night-time releases.

Skalski suggests that I use a $\mathrm{r}^{2}$ analog

$$
\mathrm{r}^{2}=(\text { Model Deviance }) /(\text { Total Deviance })
$$

analogous to the sums-of-squares-based estimate from least squares analysis. The least squares estimate is the equivalent to the square of the correlation coefficient between the response variable values and the predictor estimates (thus the term r2). Such an equality does not exist for non-linear fits. I have no problem in presenting a "r2" measure, but I believe that an estimate based on the weighted estimate of the correlation coefficient would be more appropriate (the weight being the release number).

## Page 4, Paragraph 2 of Skalski's Review

Skalski suggests that I use Anscombe's residuals instead of Pearson's residuals, and I will consider doing so in any future residual analysis. The reason that I used Pearson standardized residuals is that I felt that they would be more familiar to the fishery biologists and that the release sizes were sufficiently large for the Central Limit to hold.

Page 5, Paragraph 2 of Skalski's Review

It is unlikely that we will be able to adopt stratified sampling procedures with either 1) releases made each day or 2 ) release days randomly sampled within strata. Even with modifications in CDFG's policies, fish availability will dictate the number of releases we can make, but we will be striving for at least two releases per week. A model approach will still probably be necessary but it will be over a more representative distribution of the predictor variables, and I hope that post-data-collection stratification will be possible (e.g., according to cohort: fry, parr, smolt). Cross stratification of recovery and release days would probably not be necessary for us because, with rare exceptions, all fish are recovered within the night of the night release (or the night following the day release). However, the technique that Skalsi recommends may be beneficial for paired night and day releases, treating the day and night recoveries as strata. I appreciate the information on the availability of the SPAS program.

## Variance estimate (Pages 6-10 Skalski's Review)

On page 10, Skalski states that my estimate of the variance is "largely" correct but that individual components need to be adjusted. He doesn't state at this point what the individual components should be, I list below what I think the components probably are:

Variance of efficiency estimate: I believe my estimate of the variance of the efficiency (e)

$$
s^{2}\{e\}=s^{2}\left\{1 /\left[\exp \left[-b(0)-b(1)^{*} x(1)-\ldots\right]\right\}\right.
$$

is correct. But Skalski is correct in stating that I should be using

$$
s^{2}\left(e^{*}\right)=s^{2}(1)+s^{2}\left\{1 /\left[\exp \left[-b(0)-b(1)^{*} x(1)-\ldots\right]\right\}\right.
$$

[my formulation, not Skalski's ${ }^{3}$ ] for the variance when performing the expansion.
The $s^{2}(1)$ term, which I didn't use but should have, can be estimated by

$$
s^{2}(1)=\frac{\text { Residualdeviance }}{\text { Residualdegreesf Freedom }} * \frac{e(1 e)}{N}
$$

wherein N is the number of outmigrating fish on the date of the expansion. N would then have to be the day's estimated expanded catch.

The over-dispersion correction [(Residual Deviance)/(Degrees of Freedom)] used above is necessary because the efficiency variance among releases made on the same day significantly differed from the binomial.

[^3]S.P. Cramer \& Associates, Inc. 1999 Caswell Report

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Variance of unexpanded catch estimate: Skalski suggests that I use an estimated variance based on the binomial. I believe the appropriate "binomially-based" variance estimate would be

$$
\mathrm{s}^{2}\left(\text { catc } \boldsymbol{h}=\frac{\text { Residualldeviance }}{\text { ResidualDegreeof Freedom }} * \text { Nere}^{*}(1-\mathrm{e})\right.
$$

I, instead, used the variance among three days' catch (the day's catch, the previous day's catch, and the next day's catch) as an estimate of $s^{2}$ (catch). Skalski correctly states that my estimate will be biased upwards (too big); however, I will still probably use my conservative estimate. If I use the above over-dispersion-adjusted binomially-based estimate of the catch variance, it will be based strictly on my efficiency predictor. Given the limited number of releases on which the efficiency estimates were based, I feel that it is better to base catch estimates on the catch data which are collected each day. As will be seen in my response to the review by McDonald and Howlin, my estimate of the variance will sometimes be too liberal.

I know of no other "components" that Skalski identified to adjust the variance for.

## Error in Skalski's variance estimator

I do want to flag an error in Skalski's development. On page 6, ignoring his covariance terms, he uses the following form for the variance of a product.

$$
\sigma^{2}(x y)=E^{2}(x)^{*} \sigma^{2}(y)+\sigma^{2}(x)^{*} E^{2}(y)+\sigma^{2}(x)^{*} \sigma^{2}(y)
$$

(My notation not his.) While this is correct when the expected values and variances are known, the unbiased estimate is of the form

$$
s^{2}(x y)=x^{2} x^{2}(y)+s^{2}(x)^{*} y^{2}-s^{2}(x)^{*} s^{2}(y)
$$

Note the subtraction of the last term instead of the addition. This is because

$$
E\left(x^{2}\right)=E^{2}(x)+\sigma^{2}(x)
$$

(refer to Kendal and Stuart's The Advanced Theory of Statistics, Volume 1 under the variance of a product).

## Response to McDonald and Howlin's review.

## Missing Count Estimates (Page 2 of McDonald and Howlin's review)

I appreciate their considering my missing-value estimates (which were also applied to some counts when there were trap stoppages of both screw traps) as reasonable. The missing value estimates were of great concern to me since they were what to McDonald and Howlin referred to as "ad hoc" and were not statistically based. It should be noted that when the variance of counts was estimated, the procedure did not distinguish whether the counts were missing or not. If they were missing value estimates, then the counts from the evaluated day and from the previous day and the following day were not independent and would be positively correlated. Under these conditions, the estimated variance of the count could be too small rather than too large (as discussed earlier when responding to Skalski's comments).

Bootstrapping would be difficult for us because the dates on which releases were made did not represent a "random sampling" of the outmigration dates. Underlying bootstrapping procedures is the assumption that the sample is a random sample from the underlying population.

## Daily Efficiency (Page 2 of McDonald and Howlin's review)

I mentioned in my comments to Skalski's review that the report on the 1998 outmigration (which the reviewers apparently didn't receive) compared various release strategies; i.e., day-time versus night-time releases and hatchery- versus river-run-fish releases referred to in my comments to Skalski's review. There were also comparisons of the effect of different release positions. There was a release set that involved the "standard" location and positions located upstream and downstream of the standard. The efficiency estimates were compared using a logistic analysis of variance. The estimates did not differ substantially or significantly ( $\mathrm{P}=0.89$ for standard versus upstream and $\mathrm{P}=0.57$ for standard versus downstream; the efficiency estimates being 0.015 for standard, 0.016 for upstream, and 0.020 for downstream based on 2123, 826, and 1003 standard, upstream, and downstream released fish). In my opinion, an undo number of comparative releases within-day were being made at the expense of daily releases (there were only 8 release days in 1996 and 1997 combined and 7 release days in 1998). I felt resources could be better allocated by having more release days at the expense of multiple daily releases, and abandoning multiple daily releases was my idea. Unfortunately, only 8 release days were included in 1999.

I should note that, if multiple release sites/times are used, then they should all be used on every release day or the inclusion/exclusion of release sites should be based on a random process; otherwise, bias will result from associating different release-site/time effects with random "error" or residual variation when they are in fact biasing the coefficient estimates.

## Logistic Regression (Page 2-3 of McDonald and Howlin's review)

More discussion of the logistic in future reports is clearly warranted.
The "sample size" when making comparisons in size of fish is the number of fish measured. When applied to logistic regression, sample size should refer to the number of releases. For the 1996-1997 releases it would have included 18 releases with multiple releases within days (8 release days), in 1998 it would have included 16 releases with multiple releases within days ( 7 release days), and in 1999 it would have included 9 releases with multiple releases within days ( 8 release days). Because of homogeneity in the residual mean deviance over years ${ }^{4}$, the samples were combined to boost the degrees of freedom ( $18+16+9=43$ releases with 7 parameters estimated ${ }^{5}$ for 43-7 = 36 residual degrees of freedom).

McDonald and Howlin are correct in assuming that the fish size predictor was the size measured from a sample of recovered fish. I was inconsistent in previous years, inadvertently using size of released fish in some cases and size of recovered fish in others. In 1999 I went back through all previous years' data and corrected this inconsistency, and those 1999 revised estimates are based on recovered fish size.

The "adjustment to the binomial" was discussed. It was the multiplication of the variances of the coefficients by the mean deviance ${ }^{6}$ (or the multiplying of the standard errors by the square root of the mean deviance). This was the adjustment for overdispersion mentioned in Skalski's review.

Regarding the 10 ntu turbidity "threshold". All integer threshold values were evaluated and 10 gave by far the smallest mean deviance. I am embarrassed to say that I found no reference to this evaluation in any of the reports (I think I must have left part of my discussion out of the report on the 1996 outmigration and will have to include this discussion in the 2000 outmigration report). For the estimation of the regression coefficients, if the turbidity was less than the threshold value, zero was substituted for turbidity, otherwise the actual turbidity value was used. Recovery-day (morning following release) turbidity was used as the predictor variable and gave a coefficient that substantially and significantly differed from 0 ( $\mathrm{P}<0.0001$ ) when the threshold value of 10 was reached or exceeded. The turbidity values were equated to 0 for efficiency prediction whenever the value was less than the threshold.

## Model Selection (Page 3 of McDonald and Howlin's review)

[^4]I understand the desire for models to be driven by biological considerations as well as by statistical significance. The question is how can this be assessed. For example, should efficiency increase as flow increases or decrease as flow increases? Arguments can be put forward for either case. If flow increases, the fish may have a more difficult time avoiding the trap in which case the efficiency-to-flow coefficient should be positive. However, as flow increases, the percent volume of water entering the trap decreases; therefore the proportion of fish trapped should drop and the efficiency-to-flow coefficient should be negative. I used a very liberal P for inclusion. If $\mathrm{P}<0.2$, then the coefficient would be included. It turned out that all of the included coefficients had significance levels of $\mathrm{P}<0.06$.

## Estimated Outmigration and its Variance (Page 3-5 of McDonald and Howlin's review)

The estimate of the catch-to-efficiency ratio, Sci/Sei, given by the reviewers is only unbiased if there is truly a random sampling of days for release within reasonably homogeneous strata. This is not the case. Further, the counts are made every day ( $\mathrm{n}=\mathrm{N}$ ) but the daily efficiency is predicted from estimates derived from "sampled" days on which releases are made (true n is a small fraction of N ). Using the efficiency estimates in the reviewer's ratio estimate would be appropriate if each day's efficiency were independently estimated as the counts are (with the exception of missing value estimates). They are not; they are all predicted using the same coefficient estimates.

I believe that the estimator that I use is the appropriate one for our situation, and so we are stuck with the complicated variance estimate.

I believe that Monte Carlo simulation techniques would be take time to develop, and the resources devoted to this effort (even though they would be financially rewarding to me) would be better directed to more field releases and monitoring.

APPENDIX C. Statistical reviews by John Skalski and Lyman McDonald and Shay Howlin.

Appendix D. Stillwater analysis of data.

# Stillwater Ecosystem, Watershed \& Riverine Sciences <br> 2532 Durant Avenue Suite 201 <br> Berkeley, CA 94704 <br> Phone (510) 848-8098 Fax (510) 848-8398 

## MEMORANDUM

DATE: July 3, 2000
TO: Scott Spaulding and Craig Fleming USFWS-AFRP
FROM: Jennifer Vick, Stillwater Sciences
SUBJECT: Revised Preliminary Analysis of Data From 1999 Stanislaus River Rotary Screw Trap Experiments

This report has been revised to include additional recaputure information provided by S.P. Cramer and Associates following completion of the original report, dated March 17, 2000.

## Background

Flows in the lower Stanislaus River are controlled by the New Melones Project. Under the authority of the Central Valley Improvement Act (CVPIA), the Bureau of Reclamation and the U.S. Fish and Wildlife Service are developing operating criteria for the New Melones Project to maximize production of fish resources in the lower Stanislaus River. With this in mind, the 1999 Annual Work Plan for the Stanislaus River Basin Water Needs, CVPIA Section 3406(c)(2), identified general objectives and actions to provide information useful to the overall planning efforts for the operation of the New Melones Project. A key objective in this plan was to evaluate elements of biological water needs and flow effects, including relationship of flow volume and patterns to biological processes.

In 1999, the U.S. Fish and Wildlife Service conducted releases of marked chinook salmon smolts to assess smolt survival in the Stanislaus River. Stillwater Sciences worked with the U.S. Fish and Wildlife Service to design the releases so that a multinomial mark-recapture model could be used to estimate survival in specific reaches of the river. Field implementation of the studies was conducted by S.P. Cramer and Associates. Marked salmon were captured in rotary screw traps at two locations - Oakdale (RM 40) and Caswell (RM 8). Salmon were released at five locations - Knights Ferry (RM 56.7), immediately upstream of the Oakdale trap (RM 40), immediately upstream of the Oakdale Recreational Area (RM 40), immediately downstream of the Oakdale Recreational Area (RM 38), and immediately upstream of the Caswell trap (RM 8)
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(Figure 1). All marked fish captured in the traps were re-marked and re-released. Release and recovery data are shown in Table 1.

Stillwater Sciences developed a simple mark-recapture model to estimate survival in specific river reaches and in all reaches combined based on the releases described above. Typically, survival has been estimated by expanding the number of recaptured fish based on estimated trap efficiency. In this approach, trap efficiency is estimated by releasing marked fish immediately upstream of the trap, and efficiency is defined as the ratio of fish recaptured to fish released. This approach is vulnerable to problems of estimating trap efficiency, which may result in significant over- or under-estimation of survival. This was demonstrated on the Tuolumne River, where releases of marked fish immediately upstream of a rotary screw trap consistently and significantly underestimated trap efficiency (TID/MID 1998, 1999). The multinomial model does not rely on traditional estimates of trap efficiency but rather analyzes the data as an interlocking set of paired releases.

This study consisted of five smolt release groups totaling approximately 50,000 coded-wire tagged (CWT) fish (Figure 1). One group of approximately 25,000 CWT fish was released at Knights Ferry (RM 54.7); the second group of approximately 800 CWT fish (split into two subgroups) was released upstream of the Oakdale rotary screw trap (RM 40); the third group of approximately 10,000 CWT fish (split into two sub-groups) was released just upstream of the Oakdale Recreation ponds; a fourth group of approximately 10,000 CWT fish (split into two subgroups) was released below the Oakdale Recreation ponds at about RM 39; and the final group of 5,000 CWT fish (also split into two separate groups) was released around RM 8 upstream of the Caswell rotary screw trapping (RST) site, which served as the efficiency release for the Caswell traps. The existing RST sites, one trap at RM 40 near Oakdale and two traps near Caswell State Park (RM 8), served as the primary recovery locations for the marked fish. All release groups bore unique marks, and any fish captured in the Oakdale rotary screw trap was given a new mark and re-released. The numbers of fish released and marks used are shown in Table 1. All fish released for this study were coded-wire tagged and adipose fin-clipped. Two CWT lots were used (one for the Knights Ferry release and one for the downstream release groups). The fish groups released at the lower four sites (constituting one tag lot) also had a secondary dye inoculation mark so that their release location could be identified without sacrificing the fish, which is necessary for recovering the CWT.

The fish were released over a three day period during flows of 1,230-1,370 cfs (Figure 1, Table 1). The first day included the release of 25,000 fish at Knights Ferry, 400 fish upstream of the Oakdale rotary screw trap, and one group each of 5,000 fish above Oakdale Recreation Area and below Oakdale Recreation Area, respectively. The second day included the release of the remaining 400 fish upstream of the Oakdale rotary screw trap and the remaining 5,000 fish groups at the two Oakdale sites downstream of the trap and 2,500 fish at the Caswell site. On the final day, the remaining group of 2,500 fish was released at the Caswell site. This release strategy was intended to allow the fish to disperse in as natural a pattern as possible and to
maximize the likelihood that the fish would move through the same segments of river under the same environmental conditions.

Table 1. Fish release groups used for 1999 smolt survey evaluations and recaptures at the Oakdale and Caswell Traps.


## Analysis Methods

The tasks, as defined in the scope of work are as follows:

Task 1: Estimate survival (with confidence intervals) in the upper and middle reaches (Knights Ferry-RM 40 and RM 40-RM 38, respectively) using a multinomial model. This model treats the Knights Ferry, RM 40, and RM 38 releases as an interlocking set of three paired-release
experiments with recoveries at the Caswell trap. The model does not rely on estimated trap efficiency at Caswell.

Task 2a: Estimate survival (with confidence intervals) in all three reaches (Knights Ferry-RM 40, RM 40-RM 38, and RM 38-Caswell) and river-wide using the traditional approach, which expands recovery based on estimated trap efficiency. This task relies on recaptures at Caswell and estimated daily trap efficiency (based on the efficiency relationships developed by S.P. Cramer and Associates). For reaches Knights Ferry-RM 40 and RM 40-RM 38, compare the results of the more traditional estimates to the results of the multinomial modeling completed in Task 1.

Task 2b: Estimate survival (with confidence intervals) between Knights Ferry-RM 40 using Oakdale recovery data and the Oakdale efficiency experiment (conducted during the survival releases). Compare this with the reach Knights Ferry-RM 40 survival estimate from Tasks 1 and $2 a$.

The methods used to complete these analyses are described in Appendix A.

## Results

Results of the survival analyses are shown in Table 2 and Figures 2, 3 and 4.

Table 2. Estimated smolt survival in the Stanislaus River, 1999.

| Reach | Estimated Survival and 95 \% confidence intervals (\%) |  |  |
| :---: | :---: | :---: | :---: |
|  | Multinomial Model <br> (Task 1) | Traditional based on Caswell recoveries (Task 2a) | Traditional based on Oakdale recoveries (Task 2b) |
| Knights Ferry-RM 40 | 80 (51-100) | 77 (44-100) | 77 (40-100) |
| RM 40-RM 38 | 100 (57-100) | 100 (55-100) |  |
| RM 38-Caswell | $8.2(6.3-13)^{1}$ | $7.8(4.2-12)^{1}$ |  |
| River-wide | 6.6 (4.5-8.5) ${ }^{1}$ | 6.7 (4.4-9.6) ${ }^{1}$ |  |

${ }^{1}$ These estimates rely on traditional estimates of trap efficiency at Caswell. The estimate used is $2.1 \%$.

The benefits of the multinomial approach are limited because only two recovery locations (i.e., trap locations) were available in the design. As such, the multinomial model (Task 1) and the traditional method using Caswell recoveries (Task 2a) use exactly the same data and the same assumptions about survival and recovery of each group individually. (Superficially, the
traditional method for estimating survival in the Knights Ferry-RM 40 and the RM 40-RM 38 reaches uses an efficiency estimate at the Caswell trap, but this term cancels out algebraically, contributing nothing to the final estimator). The only differences between the two approaches are that the multinomial model is constrained by the requirement that all survival parameters in the model must be $\# 100 \%$, and that the multinomial model is able to form slightly smaller confidence intervals because it treats all three releases as a single experiment, rather than three separate experiments. Neither method provides any check on the validity of the efficiency assumption at the Caswell trap.

The multinomial model and the traditional method using Oakdale recoveries (Task 2b) use different data sets and different assumptions. In particular, the former makes no assumption about efficiency at either trap, whereas the latter depends on an efficiency-release for the Oakdale trap. In this reach, the general agreement of the estimates indicates that, at least for this particular experiment, the assumptions of the trap efficiency releases were met. It is not known whether this would be the case under other flow conditions or for other releases.

## Appendix A. Description of Models Used for Survival Analysis

## TASK 1. MULTINOMIAL MODEL

Assumptions:

- All smolts released at Knights Ferry have the same probability $\mathrm{n}_{1}$ of surviving to RM 40.
- All smolts released at RM 40, and all smolts from Knights Ferry reaching RM 40, have the same probability $\mathrm{n}_{2}$ of surviving to RM 38 .
- All smolts released at RM 38, and all smolts from Knights Ferry or RM 40 reaching RM 38, have the same probability $\sigma$ of appearing in the Caswell traps.

Let $\mathrm{n}^{\prime}\left(\mathrm{n}_{1}, \mathrm{n}_{2}, \sigma\right)$.
Under these assumptions, the probability of recovering $\boldsymbol{m}^{'}\left\{m_{1}, m_{2}, m_{3}\right\}$ smolts from the Knights Ferry, RM 40, and RM 38 releases, respectively, out of releases of $\boldsymbol{n}^{\prime}\left\{n_{1}, n_{2}, n_{3}\right\}$ smolts at these locations, is

$$
\begin{aligned}
p(\boldsymbol{m} \mid \mathrm{n}, \boldsymbol{n}) \cdot & \binom{n_{1}}{m_{1}}\left(\mathrm{n}_{1} \mathrm{n}_{2} \sigma\right)^{m_{1}}\left(1 \& n_{1} \mathrm{n}_{2} \sigma\right)^{n_{1} \delta m_{1}} \\
& \times\binom{ n_{2}}{m_{2}}\left(\mathrm{n}_{2} \sigma\right)^{m_{2}}\left(1 \& n_{2} \sigma\right)^{n_{2} \delta m_{2}} \\
& \times\binom{ n_{3}}{m_{3}} \sigma^{m_{3}}(1 \delta \sigma)^{n_{3} \delta m_{3}}
\end{aligned}
$$

The likelihood, $L(\mathbf{n} \mid \boldsymbol{m}, \boldsymbol{n})$, is any function proportional to this, considered as a function of $\boldsymbol{n}$.
Temporarily ignoring the requirement that $\mathrm{n} 0[0,1]^{3}$, the maximum value of $L$ is easily found to occur at

$$
\begin{equation*}
\hat{\mathrm{n}}_{1}, \frac{m_{1}}{n_{1}} / \frac{m_{2}}{n_{2}}, \quad \hat{\mathrm{n}}_{2}, \frac{m_{2}}{n_{2}} / \frac{m_{3}}{n_{3}}, \quad \hat{\sigma} \quad \frac{m_{3}}{n_{3}} . \tag{2a}
\end{equation*}
$$

This will be the maximum likelihood estimate when it is in the parameter space.
If the point (2a) does not lie in the parameter space, the maximum likelihood is attained somewhere on the boundary, and the estimator should be modified accordingly. The following cases can arise:

- If $\frac{m_{1}}{n_{1}}>\frac{m_{2}}{n_{2}}$ and $\frac{m_{1} \% m_{2}}{n_{1} \% n_{2}} \# \frac{m_{3}}{n_{3}}$, the estimator is

$$
\begin{equation*}
\hat{\mathrm{n}}_{1}^{\prime} 1, \quad \hat{\mathrm{n}}_{2}^{\prime} \frac{m_{1} \% m_{2}}{n_{1} \% m_{2}} / \frac{m_{3}}{n_{3}}, \quad \hat{\sigma} ' \frac{m_{3}}{n_{3}} \tag{2b}
\end{equation*}
$$

- If $\frac{m_{2}}{n_{2}}>\frac{m_{3}}{n_{3}}$ and $\frac{m_{1}}{n_{1}} \# \frac{m_{2} / m_{3}}{n_{2} \% n_{3}}$, the estimator is

$$
\begin{equation*}
\hat{\mathrm{n}}_{1}, \frac{m_{1}}{n_{1}} / \frac{m_{2} \% m_{3}}{n_{2} \% n_{3}}, \hat{\mathrm{n}}_{2} ' 1, \quad \hat{\sigma} ' \frac{m_{2} / m_{3}}{n_{2} \% n_{3}} \tag{2c}
\end{equation*}
$$

- Finally, if $\frac{m_{1}}{n_{1}}>\frac{m_{2}}{n_{2}}$ and $\frac{m_{1} \% m_{2}}{n_{1} \% n_{2}}>\frac{m_{3}}{n_{3}}$, or if $\frac{m_{2}}{n_{2}}>\frac{m_{3}}{n_{3}}$ and $\frac{m_{1}}{n_{1}}>\frac{m_{2} m_{3}}{n_{2} \sigma_{3}}$, the estimator is

$$
\begin{equation*}
\hat{\mathrm{n}}_{1}^{\prime} 1, \hat{\mathrm{n}}_{2}^{\prime} 1, \hat{\sigma}^{\prime} \frac{m_{1} \% m_{2} \% m_{3}}{n_{1} \% m_{2} \% m_{3}} \tag{2d}
\end{equation*}
$$

For the Stanislaus River data, $\boldsymbol{n}^{\prime}(25536,9378,9988), \boldsymbol{m}^{\prime}(35,17,16)$, the estimator (2c) applies, and the fitted model is

$$
\hat{\mathrm{n}}_{1}^{\prime} 0.80, \quad \hat{\mathrm{n}}_{2}^{\prime} 1.00, \quad \hat{\sigma}^{\prime} 0.0017
$$

## Classical Confidence Regions

By definition, confidence intervals for model parameters arise from the distribution of the parameters re-estimated from samples drawn from the fitted model. These distributions can be derived analytically in some cases, but when the model is non-standard, or the estimators are complicated (as here), we may as well just calculate them via simulation.

Using parametric bootstrapping ( $B=10,000$ ) with the 1999 Stanislaus River data, and applying the routine sm. density from the smoothing library of Bowman and Azzalini (1997), ten smoothed density curves were generated for each of the three parameters. These curves are shown in Figure 2, along with the consensus curve obtained by averaging.

The $95 \%$ confidence intervals associated with these marginal densities were
$0.51 \# \mathrm{n}_{1} \# 1.00, \quad 0.57 \# \mathrm{n}_{2} \# 1.00, \quad 0.0013 \# \sigma \# 0.0026$.

## Problems With Confidence Regions

When the form of the estimator can vary from sample to sample, as in (2a-d) above, the distribution of re-estimated parameters, on which the classical confidence intervals are based, can look very strange. Indeed, this was the case in Figure 2.

The problem here goes beyond aesthetics, however. Because the classical intervals are based on samples from the fitted model, "accidental" features of the basic estimate carry over to these intervals. This is particularly troublesome when, as here, the general behavior of the model is very sensitive to the parameter values. For example, if a basic estimate of survival or capture probability is exactly zero, all the re-estimated values will be also, so that the confidence
intervals will have width zero. Although technically correct, such a result is not easy to explain to non-statistical readers, nor particularly useful once explained.

This has been a problem for us in the past and has prompted us to explore other ways of quantifying parameter uncertainty. The only methods which seem applicable here are those which rely on the shape of the likelihood, regarded as a function of the possible states of nature n when the data $\boldsymbol{m}$ are held fixed.

## Marginal Likelihood

The likelihood is a joint function of all three parameters. It is hard to visualize a four-dimensional object such as the graph of this likelihood, or even three-dimensional objects such as its contour surfaces. Ordinarily, we want to consider parameters one or two at a time.

A very simple way to reduce the dimensionality is to consider cross-sections of the likelihood hypersurface along planes (or hyperplanes) perpendicular to the parameter space and passing through the maximum-likelihood estimate. Such cross-sections are shown in Figure 3.

The right way to do things, however, is to integrate out some parameters, and obtain marginal likelihoods on those remaining.

As it turns out, none of the desired integrals can be written in terms of standard functions (or at least in terms of built-in functions of S-Plus). With an eye toward generalization to a greater number of reaches (and consequently higher-dimensional integrals), and the possible introduction of Bayesian methods at some point, we chose to use a form of Monte-Carlo integration. Our algorithm is equivalent to sampling from the joint distribution proportional to the likelihood. The marginal distributions of the components of these samples are then proportional to the marginal likelihoods.

To sample from this joint distribution, consider the change of variables

$$
\theta_{1}^{\prime} \mathrm{n}_{1} \mathrm{n}_{2} \sigma, \quad \theta_{2}^{\prime} \mathrm{n}_{2} \sigma, \quad \theta_{3}^{\prime} \quad \sigma .
$$

Sampling from the distribution

$$
P_{\mathrm{n}} \% L(\mathrm{n}) d \mathrm{n},
$$

supported on the unit n -cube, is equivalent to sampling from the distribution

$$
P_{\boldsymbol{\theta}}^{S} \% L(\mathrm{n}(\boldsymbol{\theta}))\left|\frac{\mathrm{M}}{\mathrm{M} \boldsymbol{\theta}}\right| d \boldsymbol{\theta}
$$

supported on the simplex $S^{\prime}\left\{\boldsymbol{\theta} \mid 0 \# \theta_{1} \# \theta_{2} \# \theta_{3} \# 1\right\}$.
Interpret $P_{\theta}^{S}$ as the conditional distribution $P_{\theta \mid S}$, where $P_{\theta}$ is proportional to the extension of $L|M h / M \theta| d \boldsymbol{\theta}$ to the entire unit $\boldsymbol{\theta}$-cube. Then

$$
P_{\theta} \% \theta_{1}^{m_{1}}\left(1 \& \theta_{1}\right)^{n_{1} \delta m_{1}} d \theta_{1}, \theta_{2}^{m_{2} \& 1}\left(1 \& \theta_{2}\right)^{n_{2} \delta m_{2}} d \theta_{2}, \theta_{3}^{m_{3} \& 1}\left(1 \& \theta_{3}\right)^{n_{3} \delta m_{3}} d \theta_{3}
$$

which is just the product of three independent beta distributions (but notice the subtle effect of the Jacobian $|\mathbf{M} / \mathbf{M} \boldsymbol{\theta}| ' \theta_{2}^{\delta_{1}} \theta_{3}^{\delta 1}$ on the parameters of these distributions).

We sample from $P_{\theta \mid S}$ by simply drawing random samples from $P_{\theta}$ and rejecting those which are not in $S$.

This worked well for the 1999 Stanislaus River data. The marginal likelihood curves shown in Figure 4 were drawn by the method described in Section 2, using a total of 100,000 samples.

## Bayesian HPD Regions

The likelihood function is proportional to the Bayesian posterior distribution for the prior consisting of the product of independent uniform distributions on $\mathrm{n}_{1}, \mathrm{n}_{2}$, and $\sigma$. The marginal posterior distributions are simply the normalizations of the marginal likelihoods. This interpretation allows us to use the Bayesian concept of HPD (highest posterior density) regions in place of classical confidence regions.

For the 1999 Stanislaus River data, the marginal posterior distributions are just the normalizations of the marginal likelihoods, presented in Figure 4.

The 95\% HPD intervals associated with these were:
$0.54 \# \mathrm{n}_{1} \# 0.99, \quad 0.58 \# \mathrm{n}_{2} \# 0.99, \quad 0.0014 \# \sigma \# 0.0028$

## Survival in the Lowermost Reach

It is impossible to separate survival in the RM 38-Caswell reach from capture efficiency at the Caswell traps without additional data. If the capture efficiency at Caswell, $p$, were known, survival in this lowermost reach could be estimated by simply dividing the estimate for $\sigma$ by $p$. The confidence and HPD intervals would scale in the same way.

The 1999 Stanislaus River Rotary Screw Trap Program included experiments designed to estimate this efficiency. In these experiments, a total of 4,987 marked smolts were released a short distance upstream of the Caswell traps, of which 103 were subsequently recovered. This yields the efficiency $p^{\prime} 0.0207$; treating this as if it were an exact value yields an estimate of 0.082 for survival in the RM 38-to-Caswell reach, with $95 \%$ confidence and HPD intervals (0.063-0.13) and (0.068-0.14) respectively.

Of course, this value of $p$ is only an estimate, whose uncertainty should be taken into account. This would yield broader intervals for the survival, and shift the survival estimate itself slightly to the right. There are several ways this could be done; the tidiest would be to modify the basic model to have three release locations and four reaches, the last representing the segment between the efficiency release location and the trap, and setting survival in this reach to 1 . Alternatively,
one could simply treat the recovery of efficiency fish as a separate binomial or Poisson experiment to get estimates the mean and variance of $p$, use the delta method to approximate the mean and variance of $\sigma / p$, and inflate the intervals calculated above accordingly.

We do neither of these here, however, because our experience with similar experiments on the Tuolumne River has led us to suspect that conventional trap efficiency experiments like these, in which smolts are released closely enough to the trap that mortality between release and recovery can be safely neglected, may be badly biased as estimators of the efficiency appropriate to groups released much further upstream. We believe that the effect of such bias on the accuracy of the survival estimate are potentially more important than the effect of sampling error on the precision of the estimate.

## TASKS 2 AND 3. TRADITIONAL APPROACH

## Survival from Knights Ferry to RM 40, Using Data from Oakdale Trap

Survival in the Knights Ferry-RM 40 reach can be estimated using recovery of the Knights Ferry release group at the Oakdale Trap (at RM 40), together with data from the Oakdale Trap efficiency experiments.

Usually, this survival estimate is described as a two-stage process: First, capture efficiency at the trap is estimated as $\hat{p}^{\prime} m_{e} / n_{e}$, where $n_{e}$ is the number released in the efficiency experiment and $m_{e}$ is the number of these recovered at the trap. Second, survival from the upstream site is estimated as $\hat{\mathrm{n}}_{1}{ }^{\prime} m_{s} /\left(\hat{p}\left(n_{s}\right)\right.$, where $n_{s}$ is the number released in the survival experiment and $m_{s}$ is the number of these recovered at the trap.

This is equivalent to treating the survival and efficiency releases together as a paired-release experiment (note that this would not be the case if the capture efficiency were estimated separately, e.g., by using the logistic model described in (Demko and Cramer 1998) to predict efficiency from environmental variables). Confidence intervals were constructed on this basis by simulation.

For the Stanislaus River data $n_{e}{ }^{\prime} 761, m_{e}^{\prime} 6, n_{s}{ }^{\prime} 25,536, m_{s}{ }^{\prime} 156$, the survival estimated in this way is

$$
\hat{\mathrm{n}}_{1}{ }^{\prime} 0.77
$$

with $95 \%$ confidence interval

$$
0.40 \text { \# n }{ }_{1} \# 1.00
$$

## Survival from Knights Ferry to RM 40, RM 40 to RM 38, and RM 38 to Caswell, and River-Wide Survival, Using Data from Caswell Trap

The same method described above can be used with recoveries of the Knights Ferry, RM 40, and RM 38 release group at the Caswell Trap, together with data from the Caswell Trap efficiency experiments. The efficiency data at Caswell were $n_{e}{ }^{\prime} 4,987, m_{e}{ }^{\prime} 103$.

Estimates of survivals from Knights Ferry to RM $40\left(\mathrm{n}_{1}\right)$ and from RM 40 to RM $38\left(\mathrm{n}_{2}\right)$ can be found as
$\hat{n}_{1}=$ Survival from Knights Ferry to Caswell / Survival from RM 40 to Caswell, $\mathrm{n}_{2}=$ Survival from RM 40 to Caswell / Survival from RM 38 to Caswell.

These are mathematically equivalent to treating the Knights Ferry and RM 40 releases, and the RM 40 to RM 38 releases, as paired release experiments. However, the point estimates are slightly different, because the constraint $\mathrm{n}_{2} \# 1.00$ does not affect other parameters, and the confidence intervals are slightly broader, since these experiments are treated independently here:

$$
\begin{array}{lll}
\hat{n}_{1}^{\prime} \quad 0.77, \quad \hat{\mathrm{n}}_{2}^{\prime} & 1.00 \\
0.44 \# \mathrm{n}_{1} \# 1.00, & 0.55 \# \mathrm{n}_{2} \# 1.00
\end{array}
$$

Similarly, the confidence interval reported above for survival from RM 38 to Caswell is slightly broader than that found for Task 1, although the estimate itself is identical.

## REFERENCES

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Mossdale trawl


Figure 1. Smolt survival study design for the Stanislaus River.

Appendix E. Results of mark-recapture tests to estimate survival of juvenile chinook migrating through the Stanislaus River to Caswell, 1996-1999.
Release locations labeled Oakdale refer to efficiency releases in which individuals were recaptured at the Caswell traps.

| Release <br> Location | Date of <br> Release | Fish Stock | Adjusted \# Released | Total \# Recap. | \% <br> Recap. | Predicted Efficiency | Expanded Catch | Survival Index | Avg. Flow at OBB | Mean Release | Mean Recap. |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| KF | 11-Apr-98 | Hatchery | - | 1 | - | 0.01565 | 64 | VOID | 2,066 | - | 78.0 |
| KF | 02-May-98 | Hatchery | 2763 | 6 | 0.217\% | 0.01399 | 429 | 15.5\% | 1,972 | 83.2 | 88.7 |
| Oakdale | 02-Mar-98 | Natural | 929 | 2 | 0.215\% | 0.03909 | 51 | 5.5\% | 3,508 | 35.4 | 36.0 |
| Oakdale | 18-Mar-98 | Natural | 479 | 1 | 0.209\% | 0.15736 | 6 | 1.3\% | 1,768 | 62.2 | 71.0 |
| Oakdale | 02-May-98 | Natural | 392 | 1 | 0.255\% | 0.02794 | 36 | 9.1\% | 1,972 | 81.1 | 85.0 |
| Oakdale | 12-Feb-96 | Natural | 969 | 3 | 0.310\% | 0.03477 | 86 | 8.9\% | 681 | 34 | 35 |
| Oakdale | 22-Mar-96 | Hatchery | 617 | 1 | 0.162\% | 0.02726 | 37 | 5.9\% | 3,413 | 43.9 | 100 |
| Oakdale | 06-Apr-96 | Hatchery | 500 | 2 | 0.400\% | 0.02355 | 85 | 17.0\% | 1,791 | 70.6 | 76.5 |
| KF | 22-Apr-96 | Natural | 930 | 3 | 0.323\% | 0.0249 | 120 | 13.0\% | 1,673 | 86.1 | 88.3 |
| Oakdale | 04-May-96 | Natural | 547 | 1 | 0.183\% | 0.0256 | 39 | 7.1\% | 1,674 | 75.5 | 80 |
| Oakdale | 26-May-96 | Hatchery | 304 | 1 | 0.329\% | 0.04369 | 23 | 7.5\% | 921 | 72.2 | 80 |
| Oakdale | 19-Feb-99 | Natural | 326 | 1 | 0.307\% | 0.02597 | 39 | 11.8\% | 4,129 | 34.2 | 37 |
| Oakdale | 22-Feb-99 | Natural | 316 | 3 | 0.949\% | 0.02597 | 116 | 36.6\% | 4,158 | 35.8 | 37.7 |
| Oakdale | 01-Mar-99 | Natural | 193 | 1 | 0.518\% | 0.02597 | 39 | 20.0\% | 3,535 | 35.2 | 45 |
| Oakdale | 05-Mar-99 | Natural | 519 | 1 | 0.193\% | 0.02597 | 39 | 7.4\% | 2,641 | 35.8 | 38 |
| Oakdale | 30-Mar-99 | Natural | 391 | 1 | 0.256\% | 0.02597 | 39 | 9.8\% | 1,146 | 49.6 | 75 |
| KF | 01-Jun-99 | Hatchery | 25536 | 35 | 0.137\% | 0.02597 | 1348 | 5.3\% | 1,229 |  | 89.2 |
| RM38 | 01-Jun-99 | Hatchery | 4981 | 8 | 0.161\% | 0.02597 | 308 | 6.2\% | 1,229 | 85.3 | 85.4 |
| RM38 | 02-Jun-99 | Hatchery | 5007 | 8 | 0.160\% | 0.02597 | 308 | 6.2\% | 1,365 | 84.8 | 86 |
| RM40 | 01-Jun-99 | Hatchery | 4975 | 10 | 0.201\% | 0.02597 | 385 | 7.7\% | 1,229 | 84.4 | 87.3 |
| RM40 | 02-Jun-99 | Hatchery | 4403 | 7 | 0.159\% | 0.02597 | 270 | 6.1\% | 1,365 | 83.2 | 85.7 |


[^0]:    8 Mean Deviance = Deviance/(Degrees of Freedom), analogous to the Mean Square = (Sum of Squares)(Degrees of Freedom) from least squares' analysis of variance.

[^1]:    1 D. B. Demko and S. P. Cramer. June 1998. Outmigrant Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site 1997.

[^2]:    2 D.B. Demko, C. Gemperle, S.P. Cramer, and A. Phillips. June 1999. Outmigrant Trapping of Juvenile Salmonids in the Lower Stanislaus River Caswell State Park Site 1998.

[^3]:    3 Skalski's formulation was applicable to least squares linear regression

[^4]:    4 It should be noted that the among-day model-based residual deviances was not significantly greater than the within-day deviance; therefore the among-day and within-day deviances were also pooled.

    5 Four parameters in 1996-1997; two parameters in 1998 and 1 parameter in 1999 (the 1996-based turbidity coefficient estimate was used for all years)

    6 Mean Deviance $=($ Residual Deviance $) /($ Residual Degrees of Freedom $)$

